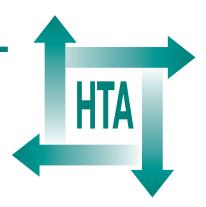
Clinical effectiveness and costeffectiveness of pioglitazone and rosiglitazone in the treatment of type 2 diabetes: a systematic review and economic evaluation

C Czoski-Murray, E Warren, J Chilcott, C Beverley, MA Psyllaki and J Cowan

April 2004

Health Technology Assessment NHS R&D HTA Programme







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Clinical effectiveness and cost-effectiveness of pioglitazone and rosiglitazone in the treatment of type 2 diabetes: a systematic review and economic evaluation

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Objectives: To evaluate the use of pioglitazone and rosiglitazone, in terms of both clinical and cost-effectiveness in the treatment of type 2 diabetes. **Data sources:** Electronic databases and the reference lists of relevant articles, in addition 14 health services research-related resources were consulted via the Internet.

Review methods: A systematic review of the literature, involving a range of databases, was performed to identify all papers relating to the glitazones. The methodological quality of the included randomised controlled trials (RCTs) was assessed using the Jadad method. A generic proforma for the critical appraisal of modelling studies in health economics was used in systematically reviewing the economic assessment studies identified. This was supplemented by a detailed review of the diseasespecific factors within the studies. Where possible, key outcomes were compared. Readers should note that information from the sponsor's submission was submitted in confidence to the National Institute for Clinical Excellence (NICE). Such information was made available to the NICE Appraisals Committee, but has been removed from this version of the report. Results: Of the 1272 studies identified, nine studies met the inclusion criteria. The clinical evidence available showed that glitazones reduce glycosylated haemoglobin by approximately 1%, and are more effective at higher doses than at lower doses. Glitazone treatment is associated with weight gain. No published data were available on the long-term effects of glitazone use. No prospective RCTs were found comparing pioglitazone to rosiglitazone, but the available evidence indicated that the two treatments had similar effects. There are no published economic studies on either pioglitazone or rosiglitazone. Economic evaluations for both glitazones were provided by the manufacturers. Sensitivity analyses undertaken by the assessment team

suggest that the cost per quality-adjusted life-year (QALY) of rosiglitazone is most sensitive to dosage and treatment effect, that is, the effect of rosiglitazone on β -cell function and insulin sensitivity. In the two scenarios where rosiglitazone is compared with metformin and sulfonylurea combination therapy, the cost-effectiveness of rosiglitazone switches from around £10,000 per QALY to being dominated by the comparator strategy. Since the baseline estimate of cost-effectiveness is not robust to changes in the treatment effect and is reliant on the many assumptions included within the metabolic and long-term economic models, caution should be used in interpreting the baseline result. **Conclusions:** The clinical evidence available showed that glitazones can reduce glycosylated haemoglobin; however there were no peer-reviewed data available on the long-term effects of their use or any prospective RCTs found comparing pioglitazone with rosiglitazone. No published economic studies on either pioglitazone or rosiglitazone were found, although sensitivity analyses undertaken by the assessment team suggest that the cost per QALY of rosiglitazone is most sensitive to dosage and treatment effect. It is suggested that research already undertaken in this area should be published, preferably in peer-reviewed journals. Direct head-to-head comparisons of the glitazones in combination with metformin or sulfonylurea would be helpful. The current licence arrangements do not allow for routine use of the glitazones in triple oral combination therapy or in combination with insulin. Evidence is emerging of use of the glitazones within such combinations; therefore, prospective RCTs would be useful. These studies could examine short-term transition strategies and longer term management. The impact of the glitazones in delaying transfer to insulin and the impact on long-term outcomes should also be considered for investigation.



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List of abbreviations

ACE	angiotensin-converting enzyme	FPI	fasting plasma insulin
AMI	acute myocardial infarction	GDM	gestational diabetes mellitus
B1	blindness in one eye	GPMDP	General Practice Morbidity Database Project
B2	blindness in both eyes	GPR	gross proteinuria
BDR	background diabetic	GSK	GlaxoSmithKline
BMI	retinopathy		
BNF	body mass index	HbA _{1c}	glycated (glycosylated) haemoglobin
	British National Formulary	HDL	high-density lipoprotein
BP	blood pressure	НОМА	homeostasis model assessment
CCTR	Cochrane Controlled Trials Register	Ι	insulin
CDSR	Cochrane Database of Systematic Reviews	ICER	incremental cost-effectiveness ratio
CHD	coronary heart disease	IDDM	insulin-dependent diabetes
CI	confidence interval		mellitus
CIGMA	continuous infusion of glucose	IFG	impaired fasting glucose
	with model assessment	IGT	impaired glucose tolerance
СРК	creatinine phosphokinase	LDL	low-density lipoprotein
CVD	cardiovascular disease	LEA	lower extremity amputation
D&E	diet and exercise	LS	least squares
DARE	Database of Abstracts of	LVH	left ventricular hypertrophy
	Reviews of Effectiveness	LYG	life-year gained
DARTS	Diabetes Audit and Research in Tayside Scotland	М	metformin
DBP	diastolic blood pressure	MI	myocardial infarction
DCCT	Diabetes Control and	МО	macular oedema
Deel	Complications Trial	NA	not applicable
DFU	diabetic foot ulceration	NHANES	National Health and Nutrition
DKA	diabetic ketoacidosis		Examination Survey
ECG	electrocardiogram	NHS EED	NHS Economic Evaluation Database
EQ-5D	EuroQol 5 Dimensions	NICE	National Institute for Clinical
ESRD	end-stage renal disease		Excellence
FDA	Food and Drug Administration	NIDDM	non-insulin-dependent
FPG	fasting plasma glucose		diabetes mellitus

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List of abbreviations continued

OGTT	oral glucose tolerance test	SBP	systolic blood pressure
OHE HEED	Office of Health Economics	SCI	Science Citation Index
	Health Economic Evaluation Database	SIGN	Scottish Intercollegiate Guidelines Network
OPCS	Office of Population, Censuses and Surveys	SmPC	Summary of Product Characteristics
Р	pioglitazone	SSCI	Social Sciences Citation Index
PDR	proliferative diabetic retinopathy	TZD	thiazolidinedione
PPAR-γ	peroxisome proliferator- activated receptor-gamma	UKDIPG	UK Drug Information Pharmacists Group
PSSRU	Personal Social Services Research Unit	UKPDS	United Kingdom Prospective Diabetes Study
PVD	peripheral vascular disease	VA-HIT	Veterans Affairs High-Density Lipoprotein Cholesterol
QALY	quality-adjusted life-year		Intervention Trial
R	rosiglitazone	VAS	visual analogue scale
RCT	randomised controlled trial	WESDR	Wisconsin Epidemiologic Study
REP	Rochester Epidemiology		of Diabetic Retinopathy
	Project, Minnesota	WHO	World Health Organization
S	sulfonylurea	WMD	weighted mean difference

All abbreviations that have been used in this report are listed here unless the abbreviation is well known (e.g. NHS), or it has been used only once, or it is a non-standard abbreviation used only in figures/tables/appendices in which case the abbreviation is defined in the figure legend or at the end of the table.



Objectives

The aim of this review was to evaluate the use of pioglitazone and rosiglitazone, in terms of both clinical and cost-effectiveness in the treatment of type 2 diabetes.

Methods

Electronic databases and the reference lists of relevant articles, in addition 14 health services research-related resources were consulted via the Internet. A systematic review of the literature, involving a range of databases, was performed to identify all papers relating to the glitazones.

The methodological quality of the included randomised controlled trials (RCTs) was assessed using the Jadad method. A generic proforma for the critical appraisal of modelling studies in health economics was used in systematically reviewing the economic assessment studies identified. This was supplemented by a detailed review of the diseasespecific factors within the studies. Where possible, key outcomes were compared.

Readers should note that information from the sponsor's submission was submitted in confidence to the National Institute for Clinical Excellence (NICE). Such information was made available to the NICE Appraisals Committee, but has been removed from this version of the report.

Results and conclusions

The total number of studies identified from these searches was 1272.

Number and quality of studies

Nine studies met the inclusion criteria.

Clinical effectiveness

The clinical evidence available showed that glitazones reduce glycosylated haemoglobin by approximately 1%, and are more effective at higher doses than at lower doses. Glitazone treatment is associated with weight gain. No published data were available on the long-term effects of glitazone use. No prospective randomised controlled trials (RCTs) were found comparing pioglitazone to rosiglitazone, but the available evidence indicated that the two treatments had similar effects.

Health economics

There are no published economic studies on either pioglitazone or rosiglitazone. Economic evaluations for both glitazones have been provided by the sponsors. In spite of the emphasis in the NICE 'Guidance for Manufacturers and Sponsors' that sponsors provide transparent economic models with a full range of sensitivity analyses performed, neither GlaxoSmithKline nor Takeda fulfilled this requirement. Even though this review is an update of the original glitazone reviews, all the economic evidence presented in the Takeda submission and the majority of the new evidence presented in the GSK submission is still marked 'Commercial in Confidence'.

Sensitivity analyses undertaken by the assessment team suggest that the cost per quality-adjusted life-year (QALY) of rosiglitazone is most sensitive to dosage and treatment effect, that is, the effect of rosiglitazone on β -cell function and insulin sensitivity. In the two scenarios where rosiglitazone is compared with metformin and sulfonylurea combination therapy, the cost-effectiveness of rosiglitazone switches from around £10,000 per QALY to being dominated by the comparator strategy. Since the baseline estimate of costeffectiveness is not robust to changes in the treatment effect and is reliant on the many assumptions included within the metabolic and long-term economic models, caution should be used in interpreting the baseline result.

Recommendations for further research

It is recommended that research already undertaken should be published, preferably in peer-reviewed journals. Direct head-to-head comparisons of the glitazones in combination with metformin or sulfonylurea would be helpful. The current licence arrangements do not allow for

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routine use of the glitazones in triple oral combination therapy or in combination with insulin. Evidence is emerging of use of the glitazones within such combinations; therefore, prospective RCTs would be useful. These studies could examine short-term transition strategies and longer term management. The impact of the glitazones in delaying transfer to insulin and the impact on long-term outcomes should also be considered for investigation.

Chapter I Aim of the review

The aim of this review is to evaluate the incremental clinical effectiveness and costeffectiveness of pioglitazone and rosiglitazone for the treatment of type 2 diabetes.

The specific objectives are:

- to evaluate the relative clinical effectiveness in terms of glycaemic control and the incidence of hypoglycaemic events
- to estimate the relative clinical effectiveness in terms of prevention of the longer term complications of diabetes mellitus
- to estimate the relative effect on overall mortality and quality-of-life adjusted mortality
- to estimate the incremental cost-effectiveness of the glitazones
- to estimate the possible cost impact on the NHS in England and Wales.

Chapter 2 Background

Description of underlying health problem

Definition of diabetes mellitus

Diabetes mellitus is a group of chronic disorders characterised by elevated blood glucose levels (hyperglycaemia). This is a consequence of inadequate control of glucose in the blood by the pancreatic hormone insulin and/or abnormal resistance to insulin. A specialised definition by the World Health Organization (WHO) is given below.

The WHO defines diabetes mellitus as "a metabolic disorder of multiple aetiology characterised by chronic hyperglycaemia with disturbances of carbohydrate, fat and protein metabolism resulting from defects in insulin secretion, insulin action, or both".¹

Glucose is the principal energy source for cellular metabolism and efficient metabolism depends on an optimum blood glucose concentration. Insulin is secreted by β -cells in the islets of Langerhans of the pancreas. Normally, the concentration of insulin in the blood increases in response to an elevation in blood glucose levels that occur naturally after eating. The action of insulin on a number of cells, including muscle and fat cells, results in absorption of glucose out of the blood, thus maintaining blood glucose levels within the normal range. Hyperglycaemia results from a total or partial lack of insulin available or ineffectual for this function. The potential consequences of hyperglycaemia are damage to many of the body's systems; in particular, the blood vessels and nerves. Loss of glycaemic control is associated with long-term complications and people with diabetes are at increased risk of cardiovascular, peripheral vascular and cerebrovascular disease.²

There are two main aetiological types of diabetes.

 Type 1 diabetes mellitus (previously termed insulin-dependent diabetes mellitus, IDDM) is a condition in which the pancreas makes little or no insulin because the islet β-cells have been destroyed through an autoimmune mechanism. The insulin-dependent tissues are less able to take up glucose and therefore there is a buildup of glucose in the body. Type 2 diabetes mellitus (previously termed non-insulin-dependent diabetes mellitus, NIDDM) is caused by two factors: the reduction in insulin production and the presence of insulin resistance in skeletal muscle and liver. Type 2 diabetes is a progressive disease in which insulin production declines as the disease progresses, resulting in increasing failure of glucose uptake into cells. In early stages of type 2 diabetes, the most significant pathology is lack of insulin secretion. Type 2 diabetes appears when insulin secretion from the pancreas falls below a level that can maintain normoglycaemia. As the disease progresses, insulin resistance remains relatively stable and insulin production declines progressively.

In addition to type 1 and 2 diabetes, the WHO classification system includes a number of other aetiological types:

- other specific types
- genetic defects of islet β -cell function
- genetic defects in insulin action
- diseases of the exocrine pancreas
- endocrinopathies
- drug or chemical-induced diabetes
- uncommon forms of immune-mediated diabetes
- other genetic syndromes associated with diabetes
- gestational diabetes mellitus (GDM) (diagnosed during pregnancy).

Individuals with diabetes mellitus may be further subdivided according to treatment:

- patients not requiring insulin
- patients who use insulin to control blood glucose levels
- patients who require insulin for survival.

The criteria for the diagnosis of diabetes in nonpregnant adults are (all values given for venous plasma):

- symptoms of diabetes and a casual plasma glucose ≥ 11.1 mmol/l (200 mg/dl). Casual is defined as any time of day without regard to time since last meal
- fasting plasma glucose (FPG) ≥ 7.0 mmol/l (126 mg/dl). During the test a sample of blood

is obtained following a period of not eating or drinking (except water) for at least 8 hours

2 hour plasma glucose ≥ 11.1 mmol/l (200 mg/dl) during an oral glucose tolerance test (OGTT). During the test a fasting blood sugar is obtained initially. The person is then asked to drink a sugary beverage (75 g anhydrous glucose dissolved in water). Blood glucose levels are then obtained every 30 minutes for the next 2 hours. A blood glucose level below 140 mg/dl at 2 hours is considered normal. A blood glucose level of greater than 200 mg/dl at 2 hours is indicative of diabetes. A blood glucose level of 140–200 mg/dl at 2 hours is indicative of the test.

Three ways to diagnose diabetes are available and each must be confirmed on a subsequent day. FPG is the preferred test because of its lower cost and ease of use. Hyperglycaemia not sufficient to meet the diagnostic criteria for diabetes is categorised as either impaired fasting glucose (IFG) or impaired glucose tolerance (IGT). IFG is diagnosed in the presence of FPG \geq 6.1 but < 7.0 mmol/l, whereas IGT implies a plasma glucose value 2 hours after a glucose load (during an OGTT) of \geq 7.8 but < 11.1 mmol/l. Both categories, IFG and IGT, are risk factors for future diabetes and cardiovascular disease (CVD).

Symptoms and complications

The main symptoms of diabetes are:

- unexplained weight loss (although appetite often increases)
- polyphagia (frequently hungry)
- polyuria (frequently urinating)
- polydipsia (frequently thirsty)
- blurred vision
- severe fatigue
- poor wound healing (cuts, scrapes, etc.)
- dry or itchy skin
- recurrent infections such as vaginal yeast infections, groin rash, or external ear infections (swimmer's ear).

The main complications of diabetes are as follows.

• Atherosclerosis can lead to widespread CVD including coronary artery disease, cerebrovascular disease and peripheral vascular disease (PVD). Diabetes is associated with a two-to three-fold increase in the risk of coronary heart disease (CHD) and stroke. CVD is also frequently related to hypertension. In type 2 diabetes up to 30% of patients in the United Kingdom Prospective Diabetes Study (UKPDS)

had hypertension at diagnosis. Nearly all patients will be hypertensive after 5–10 years, increasing the risk of CVD.

- Diabetic kidney disease is caused mainly by high blood glucose levels. Owing to damage in the small blood vessels in the kidneys, protein is released into the urine. Diabetic kidney disease is often associated with high blood pressure, which may not develop until after the kidneys have been affected. Renal failure is a common effect of diabetes. In the UK 20% of patients entering renal replacement therapy programmes have diabetes as a cause of their renal failure.
- **Diabetic retinopathy** is an eye disease generally associated with long-standing diabetes. It is a major cause of poor vision in the UK and, if left untreated, diabetic retinopathy can lead to blindness. Prolonged periods of high blood sugar levels cause damage to the small blood vessels in the retina at the back of the eye. These blood vessels initially become leaky and may become blocked off. The leakiness causes haemorrhages and exudates (leakage of fats) from the vessels onto the retina. Leakage may also cause swelling (oedema of the retina). Blocked vessels can starve the retina of oxygen, leading to the growth of new abnormal vessels from the retina. Diabetic retinopathy is the most common cause of blindness in people of working age.
- Diabetic neuropathy, leading to foot ulceration and infection can be either acute or chronic. Neuropathy can affect the nervous system, as a painful or reduced muscle function (motor control), sense of touch, or function of the inner organs and blood vessels (the autonomic system). There is also a risk of foot ulceration and amputation. Diabetic neuropathy is caused by a prolonged high blood glucose level. Once the blood glucose level rises above a certain point, nerves throughout the body gradually begin to be damaged. About 15% of people with diabetes develop foot ulcers, and 5–15% of people with diabetic foot ulcers require amputations.
- **Peripheral vascular disease** may lead to ischaemia and ultimately gangrene and amputation distally, for example, amputation of the feet or legs.
- **Susceptibility to infections**; for example, urinary tract infections.
- **Hypoglycaemia** occurs when there is a fall in blood glucose levels, and may lead to loss of consciousness. Although potentially serious, it is easily treated by oral or intravenous glucose. Patients undergoing therapies that increase the

levels of insulin in the blood have increased susceptibility to hypoglycaemia. Each individual's therapy must balance the goal of normoglycaemia against the risk of hypoglycaemia.

- Diabetic ketoacidosis (DKA) develops when there is absolute or relative insulin deficiency, secondary either to omitting insulin or under treatment during an infection or other major stress. Because of the insulin deficiency hyperglycaemia and ketoacidosis develop. Severe high blood glucose and ketoacidosis are serious and potentially life-threatening medical problems which can occur in diabetes. Ketoacidosis is rare in people with type 2 diabetes.
- Hyperosmolar non-ketotic coma: this form of acute metabolic disturbance is more common than ketoacidosis in people with type 2 diabetes. A hyperosmolar non-ketotic coma occurs when there seems to be sufficient insulin to prevent the breakdown of fat and thus prevent ketoacidosis, but where blood glucose levels become extremely high. The person becomes very dehydrated and hyperosmolar.

The risk of chronic complications can be reduced by good blood glucose and blood pressure control, and also aggressive management of cardiovascular risk factors. In addition, there is a need for regular screening for early complications of diabetes, as early identification may prevent and can certainly slow the progression of complications. Thus, for example, early identification of retinopathy can lead to laser treatment and prevention of loss of sight.

Diabetic complications are a major cause of morbidity.³ Estimates of diabetes-related mortality based on death-certificate data are seriously misleading, because diabetes will have been a contributory factor in many deaths attributed to other underlying cause. Age- and gender-specific mortality rates are higher for people with diabetes than for non-diabetic individuals.⁴

Epidemiology

Diabetes mellitus affects 2.4% of the adult population,³ of whom 200,000 have type 1 diabetes, and more than a million have type 2 diabetes.⁵ Without taking into account improved detection, the prevalence of both type 1 and type 2 diabetes will increase over the next two decades. Type 2 diabetes is more common in the elderly population, more prevalent in men than women and varies depending on ethnicity. It has been estimated that the prevalence of type 2 diabetes in the UK will more than double between 1997 and 2010.⁶ Diabetes is much more common in people of Asian Indian and Afro-Caribbean origin than people of Caucasian origin. In a Newcastle study,⁷ 17.9% of South Asians aged 25–74 years were found to have the disorder, with a further 18.7% having impaired glucose tolerance, which implies a 30–50% higher risk of the development of diabetes in 5–10 years. Weight is another important risk factor for type 2 diabetes. It is estimated that 75% of people who develop type 2 diabetes are, or have been, obese.

Current treatment options and service provision

Diabetes is a chronic illness that requires continuing medical care and patient self-management education to prevent acute complications and to reduce the risk of long-term complications. Diabetes care is complex and requires that many issues, beyond glycaemic control, be addressed.

The goal of insulin treatment is to control the amount of insulin in the bloodstream so that glucose levels are normal or near normal (normoglycaemia). The treatment of diabetes is based on individual needs.

The treatment protocol may include:

- appropriate diet to manage blood glucose level
- exercise to lower and help the body to use blood glucose
- regular blood testing for blood glucose levels.

The goal of nutrition intervention is to assist and facilitate individual lifestyle and behavioural changes that will lead to improved metabolic control. This addresses not only glycaemic control but also other aspects such as dyslipidaemia and hypertension.

Specific treatment is determined based on:

- the patient's age, overall health and medical history
- the extent of the disease
- the patient's tolerance for specific medications, procedures or therapies
- expectations for the course of the disease
- the patient's opinion or preference.

The objective of any insulin delivery regimen is to simulate the body's normal secretion of insulin in

response to dietary intake, exercise levels and the underlying metabolic state, keeping blood glucose levels as close to normal as possible.

With the help of the healthcare team, people with diabetes will maintain control of their blood glucose, blood pressure and other risk factors that may help to prevent the development of complications. This will maximise their quality of life and reduce the risk of developing long-term complications.

Medication

There is a variety of medications, along with insulin formulations, that help people with diabetes achieve adequate blood glucose control. These drugs are described below, including their actions and the role they play in helping people with diabetes to attain a healthy blood glucose range.

Type 2 therapies

Type 2 diabetes can be managed by diet and exercise alone, at least in its early stages. The purpose of a diet is to reduce energy input in order to promote weight loss, and hence insulin sensitivity. However, type 2 diabetes is a progressive disease. Nearly all patients require oral glucose-lowering drugs after some time, and most patients eventually need insulin in order to maintain satisfactory blood glucose levels.

Insulin cannot be taken orally, and must therefore be given by injection, usually through the subcutaneous route. The aim of insulin treatment is to achieve the best possible control of blood glucose levels while avoiding hypoglycaemia. There are various types of insulin and possible schedules. The three main types of insulin are classified by the speed of action. Short-acting insulin has a relatively rapid onset of action, for example for use immediately following a meal, and includes soluble insulin, insulin lispro and insulin aspart. Intermediate-acting insulin is used when longer periods of action are required and includes isophane insulin and insulin zinc suspension. The action of the third type of insulin is slow in onset and lasts for long periods, for example crystalline insulin zinc suspension.

There are four main groups of oral glucoselowering drugs currently listed in the British National Formulary (BNF).⁸

• Sulfonylurea stimulates insulin production in the pancreas and increases insulin sensitivity at the cellular level. Weight gain is the most common side-effect; other side-effects include skin rash, jaundice, sensitivity to sunlight and hypoglycaemia.

- Metformin increases insulin sensitivity at the cellular level with no effect on the pancreas, eliminating the danger of hypoglycaemia. Side-effects include gastrointestinal disorders, usually nausea, vomiting and diarrhoea, in up to 30% of patients.
- Alpha-glucosidase inhibitors work in the small intestine to slow carbohydrate and glucose absorption. Side-effects include nausea, diarrhoea and flatulence.
- Thiazolidinediones (TZD) are oral glucoselowering drugs specifically designed for type 2 diabetes. They reduce insulin resistance through the activation of peroxisome proliferatoractivated receptor-gamma (PPAR-γ).

Management guidelines

Several reviews and clinical practice guidelines for the treatment of type 2 diabetes have been developed recently.^{3,9–13} These all recommend a 'step-up' policy of treatment, starting with diet and lifestyle advice alone, and adding various oral glucose-lowering agents and eventually insulin if targets are not achieved. Type 2 diabetes is progressive. Hence, although patients may be adequately managed initially on diet alone, within 3 years of onset 50% of patients require multiple therapy, and after 9 years this figure has increased to 75%.¹⁴

The guidelines recommend that individual treatment targets should be set, based on the need to achieve good control of blood glucose and cardiovascular risk factors, while avoiding the risk of hypoglycaemia and maintaining an acceptable quality of life. The WHO blood glucose cut-offs are designed for diagnosis, and should not be used as therapeutic targets. Commonly used figures for high risk values are shown in *Table 1*.

The commencement of an oral glucose-lowering drug is advocated if blood glucose levels remain high after an adequate trial of lifestyle education. Initiation of an oral agent is suggested (by the now outdated European guidelines) when HbA_{1c} > 6.5% (FPG > 6.0 mmol/l), or occasionally (if other risk factors are low) when HbA_{1c} > 7.5% (FPG > 7.0 mmol/l).³ Attempts to modify lifestyle factors should continue alongside medical treatment.

The choice of initial oral glucose-lowering drug depends on the patient's weight (metformin is advocated for obese patients) and on their expected susceptibility to the various side-effects. Dose titration is recommended, starting with a low dose

TABLE I Vascular risk assessment	guidelines
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	Low risk	At risk	High risk
Blood glucose			
HbA _{Ic} (%)	≤ 6.5	6.5–7.0	> 7.0
Venous fasting plasma glucose (mmol/l)	≤ 6.0	6.0–7.0	> 7.0
Self-monitored fasting blood glucose (mmol/l)	≤ 5.5	5.5-6.0	> 6.0
Blood lipids			
Total serum cholesterol (mmol/l)	< 4.8	4.8-5.0	> 5.0
Serum LDL cholesterol (mmol/l)	< 3.0	3.0	> 3.0
Serum HDL cholesterol (mmol/l)	> 1.2	1.2–1.0	< 1.0
Serum triglycerides (mmol/l)	< 1.7	1.7–2.2	> 2.2
Blood pressure			
Blood pressure (mmHg)	< 140/85		> 140/80

and gradually increasing towards the ceiling dose if targets are not met. Dosages should be reviewed and reduced if adverse effects are observed or if blood glucose is well within the target range.

The guidelines differ with respect to the recommended sequence and timing of the next step, after failure with a single oral glucose-lowering agent. Some recommend a trial of another single oral agent, before moving to combination therapy.⁹ Other guidelines recommend adding another oral agent to current medication.^{3,11} The European guidelines³ suggest that triple therapy with three differently acting agents may be tried if targets cannot be achieved on the maximum tolerated doses of two drugs.

If blood glucose levels remain high after an adequate trial of oral glucose-lowering drugs then insulin therapy is recommended (unless the patient has a poor life expectancy and is asymptomatic). The European guidelines suggest that, for most patients, insulin should be added to oral medication if $HbA_{1c} > 7.5\%$ after "maximum attention" to diet and oral medication.³

The guidelines also make a range of other recommendations relating to:

- antihypertensive therapy
- the location and organisation of services (primary/secondary/shared care)
- the professional skills that should be included in the diabetes team (GP and practice nurse, consultant physician, diabetes specialist nurse, dietitian, chiropodist and other specialists as necessary)
- the need for structured patient education and self-care programmes
- the need for self-monitoring and regular

professional checks to ensure that blood glucose levels are maintained as close to optimal levels as is possible

 the need for a range of screening tests to monitor other risk factors, side-effects and complications (e.g. blood pressure monitoring, an annual test for urinary protein and microalbuminuria, regular eye and foot checks).

The burden of disease

Estimates of the financial cost of diabetes vary enormously, depending on whether they include all costs or only healthcare costs and on whether they include costs of disease associated with or caused by diabetes.^{15–19}

The estimated total cost to the NHS, including inpatient, prescription and GP consultation costs, of treating diabetes mellitus (all types) has been estimated at £243 million for the UK in 1995/96.²⁰ This represents in real terms (i.e. inflation adjusted) an increase of around 25% since 1989. Prescriptions make up the largest component of this cost estimate, closely followed by inpatient care. However, this figure only includes the direct cost of treating disease specifically attributed to diabetes. It does not include the cost of treatments where diabetes was a contributory factor.

Another estimate, based on a survey of a district in South Wales,²¹ found that the additional hospital costs for people with diabetes were £1800 per person. This represents 9% of UK hospital costs, around £1900 million per year.

Description of intervention

The thiazolidinediones (glitazones) are a recently developed class of oral glucose-lowering drugs.^{22,23}

They work through the activation of PPAR- γ receptors, so reducing insulin resistance.²⁴ Glitazones are not intended for type 1 diabetes.

There are currently two thiazolidinedione drugs licensed in the UK:

- rosiglitazone: Avandia (SmithKlineBeecham)
- pioglitazone: Actos (Takeda & Eli Lilly).

The main adverse effect of these drugs is weight gain. There have been some instances of hepatocellular dysfunction, and therefore it is recommended in the Summaries of Product Characteristics (SmPC)^{25,26} that patients being treated with these drugs undergo liver function tests every 2 months during the first year of treatment, and at regular intervals thereafter.

Outcome measures

Principal goals of treatment

The principal aim of treatment in diabetes is the reduction of mortality and morbidity resulting from increased blood glucose levels, while maintaining a good quality of life. HbA_{1c} should ideally be $\leq 7\%$, but adjusted to accommodate

rates of hypoglycaemia acceptable to people living with diabetes. Insulin secretion in non-diabetics is characterised by continuous basal secretion with peaks immediately after meals and steady release throughout the night. Insulin requirements are at a low during the early morning.

Glycaemic control

The Diabetes Control and Complications Trial (DCCT) and the UKPDS demonstrated that HbA_{1c} must be reduced to <7% to minimise the development of microvascular complications.

Cardiovascular risk factors

CVD is a major complication and the leading cause of premature death among people with diabetes. The UKPDS recommends reporting this increased risk, as for every 1.0% haemoglobin higher HbA_{1c} there is an increase in risk of:

- 21% for any diabetes-related end-point [95% confidence interval (CI) 17 to 24, *p* < 0.0001]
- 21% for any diabetes-related deaths [95% confidence interval (CI) 15 to 27, *p* < 0.0001]
- 14% for myocardial infarction (MI) [95% confidence interval (CI) 8 to 21, p < 0.0001]
- 37% for microvascular complications [95% confidence interval (CI) 33 to 41, p < 0.0001].^{27,28}

Chapter 3 Clinical effectiveness

Search strategy

Sources searched

Twelve electronic bibliographic databases were searched, covering biomedical, health-related, science, social science and grey literature. A list of databases is provided in Appendix 1.

In addition, the reference lists of relevant articles were checked and 14 health services research-related resources were consulted via the Internet. These included health technology assessment organisations, guideline-producing bodies, generic research and trials registers and specialist diabetes sites. A list of these additional sources is given in Appendix 2. Finally, citation searches of key papers were undertaken using the Science Citation Index (SCI) citation facility and the reference lists of included studies were checked for additional studies.

Search terms

A combination of free-text and thesaurus terms was used. 'Intervention' terms included: glitazone(s), thiazole(s), thiazolidinedione, PPAR gamma agonist(s), pioglitazone, actos, 111025-46-8 (CAS registry number), ad 4833 and u 72107 (product codes), rosiglitazone, avandia, 122320-73-4 (CAS registry number) and BRL 49653 (product code). Copies of the search strategies used in the major databases are included in Appendix 3.

Search restrictions

No date, language, study or publication type restrictions were applied.

Inclusion and exclusion criteria

The above search strategy identified 1272 references. These were screened on the basis of their titles and abstracts, for relevance to the study question. There were a considerable number of animal studies, preclinical studies, general comment, general review papers or specifically monotherapy or placebo studies. All relevant reviews were looked at for additional references to primary research. Full copies of papers that appeared relevant were requested.

Studies were then assessed against the following criteria, and all studies that met the criteria were included:

- **intervention**: pioglitazone or rosiglitazone in combination with other hypoglycaemic drugs
- comparator: other antidiabetic drugs.
- subjects: patients with type 2 diabetes mellitus
- **outcome measures**: to include at least one of:
- glycaemic control (blood glucose or HbA_{1c})
- cardiovascular risk factors, lipids, weight
- adverse events
- **study methodology**: to include at least one of:
 - randomised controlled trial (RCT)
 - systematic review
 - economic evaluations.
- **length of study**: at least 12 weeks on study medication.

Quality assessment strategy

A standard checklist, the Jadad,²⁹ was used to assess the methodological quality of the included RCT. Two reviewers considered the papers independently.

Clinical effectiveness of rosiglitazone in the treatment of type 2 diabetes

Background

This report is an update of the review undertaken by Lord and colleagues.³⁰ *Tables 2–5* summarise the results of the meta-analyses undertaken for the original review, examining the effectiveness of the two licensed doses of rosiglitazone (4 and 8 mg) as a combination therapy (with metformin or sulfonylurea). In addition, the evidence provided by the additional trials is appended to the appropriate table.

Overall, a meta-analysis of trials showed that the addition of rosiglitazone to metformin and sulfonylurea led to significantly greater reductions in blood glucose over 6 months, with a higher response rate than placebo therapy. Rosiglitazone combination therapy was found to increase β -cell function and, at high doses, to decrease significantly insulin resistance compared to placebo combinations.

The addition of rosiglitazone led to significantly higher increases in HDL and LDL cholesterol over 6 months compared with metformin alone. This



	Original report	Study 044
	WMD (95% CI fixed)	Information from the sponsor's submission was submitted in confidence
HbAlc	-0.75 (-1.17 to -0.33)	to NICE. This information was made available to the NICE Appraisals
FPG	–2.13 (–2.95 to –1.31)	Committee but has been removed from this version of the report.
HDL cholesterol	0.12 (0.04 to 0.20)	
LDL cholesterol	0.33 (0.08 to 0.58)	
BP systolic	-2.10 (-5.96 to 1.76)	
BP diastolic	-0.90 (-3.40 to 1.60)	
Weight	1.90 (1.17 to 2.63)	
		nitted in confidence to NICE. This information was made available to the NIC
WMD (95% CI fixe differences in HbA ₁	c change between trial and pla	his version of the report. (95% confidence interval fixed); HbA _{1c} , values given as percentage icebo arms; FPG, fasting plasma glucose; values given as mmol/l difference given as mmol/l difference in cholesterol change between arms; BP, blood

TABLE 2 Effectiveness of rosiglitazone 4 mg with metformin

 TABLE 3 Effectiveness of rosiglitazone 4 mg with sulfonylurea

	Original report	Study 128 ($n = 53$)	Study 132
	WMD (95% CI Fixed)	Information from the sponsor's	Information from the sponsor's
HbAlc	-1.08 (-1.30 to -0.86)	submission was submitted in	submission was submitted in
FPG	-2.58 (-3.03 to -2.14)	confidence to NICE. This	confidence to NICE. This
HDL cholesterol	0.05 (0.01 to 0.08)	information was made available to	information was made available to
LDL cholesterol	0.27 (0.17 to 0.36)	the NICE Appraisals Committee	the NICE Appraisals Committee
BP systolic	Ì NA Í	but has been removed from this	but has been removed from this
BP diastolic	NA	version of the report.	version of the report.
Weight	2.35 (1.86 to 2.83)		

Note: confidence intervals were not available for some results.

TABLE 4	Effectiveness	of rosiglitazone	8 mg with metformin
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	Original report	Study 044
	WMD (95%CI Fixed)	Information from the sponsor's submission was submitted in confidence to
HbAlc	–0.88 (–1.18 to 0.57)	NICE. This information was made available to the NICE Appraisals
FPG	-2.75 (-3.38 to -2.13)	Committee but has been removed from this version of the report.
HDL cholesterol	0.13 (0.07 to 0.20)	
LDL cholesterol	0.29 (0.11 to 0.47)	
BP systolic	-0.07 (-2.89 to 2.75)	
, BP diastolic	–1.59 (–3.32 to 0.13)	
Weight	3.33 (2.68 to 3.98)	

Note: confidence intervals were not available for some results.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

was also true when compared with sulfonylurea therapy in most cases; however, changes in HDL levels were not significantly different for the 4 mg group. At 6 months, there was no significant difference in HDL cholesterol levels. The 4 mg dose had an insignificant effect on LDL cholesterol levels, but the 8 mg dose produced significantly higher levels of LDL cholesterol at 6 months. Some data suggest that the initial increase in LDL cholesterol seen with rosiglitazone

TABLE 3 Effectiveness of rosignuazone of the with surjoinvioled	TABLE 5	Effectiveness of rosiglitazone 8 mg with sulfonylurea
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	Study 132
HbA _{Ic} FPG HDL cholesterol LDL cholesterol BP systolic BP diastolic Weight	Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Note: no results were available from the previous metaanalysis for this combination.

stabilises, whereas HDL cholesterol continues to increase over 18 months. This reduction appears to be statistically significant for the rosiglitazone/metformin group but not for the rosiglitazone/sulfonylurea group.

The mean diastolic blood pressure (DBP) was lower in the rosiglitazone group in the previous meta-analysis. Although increases in body weight were significantly greater for the rosiglitazone/metformin combination therapy groups than for the metformin control groups, there were no significant absolute differences between the groups in body weight at 6 months. The rosiglitazone/sulfonylurea combination groups showed significantly greater weight increases over 6 months than the sulfonylurea control groups.

There was no significant difference between the rosiglitazone/metformin combination groups and the metformin group in terms of the proportion of patients who experienced at least one adverse event, withdrew from the studies because of an adverse event or withdrew for any reason. There was also no significant difference in the incidence of adverse events for the sulfonylurea combination arms compared with the control arms, but significantly lower proportions of patients withdrew in the 8 mg rosiglitazone/sulfonylurea combination groups compared with the controls. However, it is not licensed at the 8 mg dose for this combination in the UK.

The addition of rosiglitazone to therapy was associated with a significant reduction in the risk of hyperglycaemia in the case of metformin; however, there was no significant effect in the case of sulfonylurea. Significant increases in the incidence of hypercholesterolaemia and hyperlipaemia were observed with rosiglitazone/sulfonylurea therapy compared with sulfonylurea alone. For rosiglitazone/metformin combination therapy, a significant increase in hyperlipaemia (but not hypercholesterolaemia) was observed. Rosiglitazone/metformin therapy compared with metformin alone was associated with a reduction in the incidence of hypertension and diarrhoea. The incidence of anaemia and oedema was higher for the rosiglitazone combination therapies than for the controls. No other significant differences were noted between rosiglitazone and placebo arms in the area of adverse events.

New studies

This section critically reviews the new evidence on the effectiveness of rosiglitazone identified through the search strategy described above or submitted by the sponsor [GlaxoSmithKline (GSK)], and compares it with the estimates in the previous report³⁰ (see *Table 6*).

Information from the sponsor's submission was submitted in confidence to the National Institute for Clinical Excellence (NICE). This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Three additional study reports were identified in the searches of the published literature. One of these reported the results of the Mexican arm of study 044 (which was submitted to the Institute by the sponsor in confidence),³¹ which is reviewed below. Therefore, this paper is not considered separately. A second report³² had a Jadad score of 0 and therefore is not considered further. The third study, by Raskin and colleagues,³³ examines the use of rosiglitazone in subjects treated with insulin. This paper was not considered in the previous report and is reviewed below. This combination is not licensed in the UK.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Raskin and colleagues: A randomised trial of rosiglitazone therapy in patients with inadequately controlled insulin-treated type 2 diabetes³³

This three-armed study randomised 319 patients between insulin plus placebo, insulin plus rosiglitazone 4 mg and insulin plus rosiglitazone 8 mg. The primary outcome measure was the change in HbA_{1c} at 26 weeks compared to baseline. (The mean baseline measurement \pm SD was 8.9 \pm 1.1 to 9.1 \pm 1.3.) The mean difference in the HbA_{1c} change from baseline was -0.7% for the low dose rosiglitazone group and -1.2% for the high dose rosiglitazone group. Compared to

TABLE 6 Methodological quality of new trials submitted by the sponsor

Jadad criteria	Study 044	Study 085	Study 108	Study 128	Study 132	Study 134	Study 136	Raskin et <i>al.</i> , 2001 ²⁸
AI	Information fro	m the sponsor's subn	nission was submitte	ed in confidence to N	NCE. This information	n was made availabl	e to the	Y
A2	NICE Appraisal	s Committee but has	been removed from	this version of the i	report.			Α
BI								А
CI								Y
C2								Y
C3								Y
C4								Y
DI								Ν
D2								Y
D3 dropouts								
Jadad								4
Sample size achieved?								
inadequate (B), or masked from the in intention-to-treat a	unclear (C)? CI, W nvestigators? C4, W nalysis performed?	/as the trial describe /as the treatment all ? D3, What were the	d as 'double-blind' location masked fro e dropout rates in e	C2, Was the treat om the outcome as each group of the t	ment allocation mas sessments? D1, Was	ked from the partic the number of wit nain outcomes? (Dr	cipants? C3, Was the hdrawals in each gr	ealment adequate (A), e treatment allocation oup stated? D2, Was an en for each trial in the three

baseline, the mean insulin dose for the low-dose rosiglitazone reduced by 5.6 units, compared with 0.6 units for the placebo group. The mean insulin dose in the high-dose rosiglitazone group reduced by 12 units. There was no statistically significant difference in HDL cholesterol or LDL HDL cholesterol ratios in the patients treated with rosiglitazone.

Summary of new evidence on rosiglitazone

The additional studies confirm the efficacy of rosiglitazone as described in the original review by Lord and colleagues.³⁰

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Clinical effectiveness of pioglitazone in the treatment of type 2 diabetes

Background

The clinical effectiveness of pioglitazone was reviewed by Chilcott and co-workers for NICE in March 2001.⁴

Summary of evidence reviewed by Chilcott and colleagues

Eleven studies were included in the previous review. Of these, seven referred to pioglitazone as a monotherapy in trials of placebo versus various doses of pioglitazone, and hence are not relevant to this analysis. One study was kept in confidence and is therefore not reviewed here. The remaining studies included study PNFP-010, comparing sulfonylurea with placebo to sulfonylurea with 15 mg or 30 mg pioglitazone daily, study PNFP-014, comparing insulin with placebo to insulin with pioglitazone 15 mg and 30 mg (contrary to the indication in the SmPC, which contraindicates combination therapy with insulin), and study PNFP-27, comparing metformin with placebo to metformin with pioglitazone 30 mg (*Table 7*).

Information was available on the characteristics of the study populations in studies PNFP-010, PNFP-014 and PNFP-027. No information was available on co-morbidities of patients recruited to PNFP-010 or PNFP-027. There were no statistically significant differences in baseline characteristics between the study groups in either study. In the former study, 30/187, 29/184 and 23/189 patients were withdrawn from the placebo, 15 mg and 30 mg arms, respectively. Of these, lack of efficacy was the reason in 13, 12 and 4 patients,

Study	Subjects	Treatment groups (no. randomised)	Study procedure	Outcomes reported
PNFP-010	Type 2 diabetics, aged $30-75$ years, treated with sulfonylurea alone or with acarbose or metformin. HbA _{1c} > 8% at screening and randomisation, fasting C peptide > 1 ng/ml		2 week screening period, then 4 weeks on sulfonylurea + placebo, then 16 weeks on allocated treatment. Patients were maintained on previous dose of sulfonylurea	HbA _{1c} , FBG, insulin C peptide, triglycerides, HDL and LDL cholesterol, body weight
PNFP-014	Type 2 diabetics, treated with insulin > 30 units/day for at least 30 days. HbA _{1c} > 8% at screening and randomisation, fasting C peptide > 0.7 ng/ml	I + placebo (187) I + P 15 mg/day (191) I + P 30 mg/day (161)	2 week screening period, then 4 weeks on insulin + placebo, then 16 weeks on allocated treatment. "No attempt made to change insulin regimen"	HbA _{1c} , FBG, insulin C peptide, triglycerides, HDL and LDL cholesterol, body weight
PNFP-027	Type 2 diabetics, treated with metformin for > 30 days. HbA _{1c} $> 8\%$ at screening and randomisation, fasting C peptide > 1 ng/ml	M + placebo (153) M + P 30 mg/day (161)	2 week screening period, then 4 weeks on metformin + placebo, then 16 weeks on allocated treatment	HbA _{1c} , FBG, insulin C peptide, triglycerides, HDL and LDL cholesterol, body weight

TABLE 7 Studies identified in the previous review

respectively [Food and Drug Administration (FDA) website].

A meta-analysis of results was not appropriate for the pioglitazone studies because each trial evaluated pioglitazone in combination with a different drug. All studies took place over a period of 16 weeks.

When used in combination with metformin, sulfonylurea or insulin, pioglitazone, at doses of 15 mg or 30 mg daily, led to a significant fall in blood glucose and HbA_{1c}. The effect was greater at the higher than the lower dose. The extent of the fall in HbA_{1c} was between 0.64 and 1.26 percentage points (*Table 8*). Because the full fall in blood sugar took 8 (studies PNFP-010, PNFP-014) to 12 (study PNFP-027) weeks, the HbA_{1c} changes at 16 weeks may not fully reflect the fall in blood glucose and so may underestimate the overall effect. This effect is maintained for at least 40 weeks. Secondary analysis suggests that the effect was greater in women than in men, and was also greater in those with a higher than those in a lower initial HbA_{1c} level.

There was a significant fall in both fasting C peptide and insulin levels on pioglitazone combination treatment in studies PNFP-010 and PNFP-027. In study PNFP-014 (pioglitazone combined with insulin treatment), the fall in fasting C peptide levels was significant for the 15 mg pioglitazone group, but not for the 30 mg group.

The level of triglyceride in the 30 mg pioglitazone combination group was significantly reduced compared with the placebo combination group in each of the three trials, with a fall in the order of 0.3 mmol/l. HDL levels increased in pioglitazone combination treatment groups compared with placebo combination groups (*Table 9*). There was no change in total cholesterol or LDL cholesterol levels (*Table 10*).

Unfortunately, the table with details of the LDL levels in study PNFP-014 appears to have been misprinted in the statistical review on the FDA website. The results are therefore not reproduced here.

The changes seen at 16 weeks (a fall in triglycerides and an increase in HDL) are maintained at 40 weeks in both the metformin and sulfonylurea combination studies.³⁴ There is no further change in lipid levels at 40 weeks over that reported at 16 weeks. The changes seen at 16 weeks (a fall in triglycerides and an increase in HDL) are maintained at 40 weeks in both the metformin and sulfonylurea combination studies.³⁴ There is no further change in lipid levels at 40 weeks and an increase in HDL) are maintained at 40 weeks in both the metformin and sulfonylurea combination studies.³⁴ There is no further change in lipid levels at 40 weeks over that reported at 16 weeks.

While there is no direct evidence available on the effect of pioglitazone on diabetic complications, including cardiovascular mortality, there is evidence from the UKPDS study^{27,28,35} that improved glycaemic control reduces the incidence of microvascular complications, so it would not be unreasonable to expect that this would hold true if the improved glycaemic control is achieved through using pioglitazone.

There is evidence from the clinical trials that pioglitazone does have an impact on recognised

TABLE 8	Effect of	Dioglitazone on	HbA.	(%) in	combination therapy
		progrituzone on	TIDATC	(/0)	combination therapy

PNFP-010	S + placebo	S + P I5 mg/day	S + P 30 mg/day
Baseline	9.86	10.01	9.93
Mean change	0.06	-0.82	-1.22
LS mean difference between			
placebo and P arms		–0.88 (Cl –1.17 to –0.58)	-1.28 (Cl -1.57 to -0.99)
PNFP-014	I + placebo	I + P I5 mg/day	I + P 30 mg/day
Baseline	9.75	9.75	9.84
Mean change	-0.26	-0.99	-1.26
LS mean difference between			
placebo and P arms		–0.73 (Cl –1.00 to –0.47)	-1.00 (Cl -1.27 to -0.74)
PNFP-027	M + placebo		M + P 30 mg/day
Baseline	9.77		9.92
Mean change	0.19		-0.64
LS mean difference between			–0.83 (Cl –1.15 to –0.51)
placebo and P arms			

cardiovascular risk factors. When pioglitazone is used in combination therapy there is a consistent fall in triglycerides when doses of 30 mg or more are used, and also a statistically significant increase in HDL levels. These changes are achieved within 8 weeks of treatment, and could be expected to lead to a reduction in cardiovascular risk, other things being equal.

Any consequent reduction in cardiovascular risk will, however, be countered by the increased risk associated with the significant and progressive weight gain observed on treatment. In the combination studies, body weight increased significantly in the pioglitazone groups compared with the placebo groups. The treatment differences from placebo were related to the dose of pioglitazone administered and were greater when pioglitazone was combined with insulin or sulfonylurea than when it was combined with metformin (*Table 11*).

A safety concern associated with pioglitazone was the possibility that it might be associated with hepatitis in the same way that troglitazone was. In the studies reviewed by the FDA, the reported incidence of elevation of alanine transaminase greater than three times the upper limit of normal was no different between pioglitazone-treated patients and placebo. At 0.26% it was lower than the reported rate in troglitazone-treated patients in controlled trials (1.90%), and therefore is in line with the reported rate in other diabetic agents. However, the relatively small number of patients with long-term exposure to pioglitazone means that a long-term tendency to produce hepatitis cannot be ruled out. Such a tendency would be plausible if the hepatotoxicity of pioglitazone were equivalent to that of troglitazone on a weight-for-weight basis (bearing in mind that the therapeutic dose of troglitazone is much greater than pioglitazone, so that the net exposure was greater for that drug).

Other possible adverse events relate to oedema, haemoglobin, creatinine phosphokinase (CPK) elevation and hypoglycaemia. Oedema is more commonly reported as an adverse event in patients treated with pioglitazone than with

TABLE 9 Effect of pioglitazone on HDL cholesterol (mmol/l) in combination therapy⁴

PNFP-010	S + placebo	S + P I5 mg/day	S + P 30 mg/day
Baseline	1.11	1.07	1.08
Mean change	-0.03	0.04	0.10
LS mean difference between placebo and P arms		0.06 (CI 0.02 to 0.11)	0.13 (CI 0.08 to 0.17)
PNFP-014	I + placebo	I + P I5 mg/day	I + P 30mg/day
Baseline	1.10	1.12	1.11
Mean change	-0.02	0.06	0.07
LS mean difference between		0.07 (Cl 0.02 to 0.13)	0.09 (CI 0.03 to 0.14)
placebo and P arms			
PNFP-027	M + placebo		M + P 30 mg/day
Baseline	1.09		1.11
Mean change	0.00		0.08
LS mean difference between			0.08 (Cl 0.03 to 0.13)
placebo and P arms			

TABLE 10 Effect of pioglitazone on LDL cholesterol (mmol/l) in combination therapy⁴

PNFP-010 Baseline Mean change LS mean difference	S + placebo 3.22 0.15	S + P I5 mg/day 3.22 0.08 -0.08 (CI -0.23 to 0.08)	S + P 30 mg/day 3.28 0.13 -0.02 (CI -0.18 to 0.14)
PNFP-027 Baseline Mean change LS mean difference between placebo and P arms	M + placebo 3.06 0.07		M + P 30 mg/day 3.09 0.18 0.11 (CI –0.03 to 0.24)

PNFP-010	S + placebo	S + P I5 mg/day	S + P 30 mg/day
Mean change	-0.83	+2.18	+3.06
PNFP-014	I + placebo	I + P I5mg/day	I + P 30 mg/day
Mean change	_0.↓↓	+2.53	+3.92
PNFP-027 Baseline Mean change LS mean difference between placebo and P arms	M + placebo 93.96 –1.06		M + P 30 mg/day 93.24 +1.41 2.48 (Cl 1.72 to 3.23)

TABLE 11 Effect of pioglitazone on bodyweight (kg) in combination therapy

placebo in both the monotherapy and combination therapy trials.⁴ The overall figures quoted on the FDA website for oedema are 6.6% for pioglitazone-treated patients and 2.3% for placebo-treated patients. There is a consistent, but not clinically significant, fall in haemoglobin in patients treated with pioglitazone, in the order of 0.38 g/dl. New electrocardiogram (ECG) findings were equally distributed between pioglitazone- and placebo-treated patients. (One patient, in study PNFP-026, was noted to have developed left ventricular hypertrophy (LVH) and left bundle branch block on the ECG, which resolved when the drug was withdrawn. Five other patients were noted to have cardiomegaly on X-ray.) Seven male patients in the studies were reported to have CPK values greater than ten times the normal upper limit. Four normalised on the drug, two off the drug, and one had falling, but not yet normal, levels on follow-up. As noted earlier, weight gain is a worrying side-effect of pioglitazone treatment. It was reported in many of the studies reviewed in this report.

New evidence

For this review Takeda provided a full trial report for one study not included in the original review.²⁶ This study examined pioglitazone as a monotherapy and is not therefore considered in this review. In addition, Takeda forwarded a list of 50 public domain abstracts and papers related to the efficacy of pioglitazone.

Three clinical effectiveness studies were identified in the information sent by Takeda, or through the search strategies described above.^{26,36,37} Comparison of these published papers with the information provided in the previous report established that two of these papers^{36,37} were reviewed in the original review, as studies PNFP-010 and PNPF-027. One additional effectiveness study for pioglitazone was identified. This study was an unlicensed higher dose and has not been included. The summary of the evidence on clinical effectiveness described by Chilcott and colleagues still stands:

"In combination therapy, pioglitazone appeared to be effective in reducing blood glucose in patients with poorly controlled Type 2 diabetes. When used in combination with metformin, sulfonylurea or insulin, pioglitazone led to a significant fall in blood glucose and glycalated haemoglobin (HbA_{1c}) at high and low doses, with greater effect at the higher than at the lower dose. Pioglitazone treatment is associated with significant weight gain in the short term, which appears to be greater than that seen with other thiazolidinediones. It also appears to be greater than that seen with sulfonylurea or insulin treatment in the UKPDS, which in turn was greater than that seen on metformin treatment. This weight gain continues, albeit at a lesser rate, for more than a year."⁴

Additional new evidence

After submission of the assessment report, Takeda provided NICE with the results of four new trials of pioglitazone in type 2 diabetes. Owing to the late arrival of these trials, the trial results were not included in this review. However, these trials were provided in time for them to be considered by the NICE Appraisal Committee.

A comparison of the clinical effectiveness of rosiglitazone and pioglitazone

One retrospective medical records study was identified which attempted to compare the clinical effectiveness of the two therapies.³⁸

Boyle and colleagues randomly selected 3175 patient records of adults with type 2 diabetes who received either pioglitazone or rosiglitazone between 1 August 1999 and 31 August 2000. After review with well-defined inclusion criteria, 1115 records were included in a comparison of changes in triglycerides, total cholesterol, HDL cholesterol, LDL cholesterol and HbA_{1c} between the two treatments. Triglyceride reduction was significantly greater in the pioglitazone group than in the rosiglitazone group. Rosiglitazone was associated with an increase in total cholesterol, whereas total cholesterol was significantly reduced in the pioglitazone group. HDL cholesterol was similar between the two groups at follow-up but LDL cholesterol was significantly reduced in the pioglitazone group and significantly increased in the rosiglitazone group. The reduction in HbA_{1c} was almost identical between the two groups.

In general, these findings were broadly supported by the results of the studies reviewed here, although no formal analysis was performed.

Although these results point to potential advantages in the use of pioglitazone compared with rosiglitazone, it must be remembered that this evidence is taken from a retrospective review of records. In addition, there was variation in the combinations of therapies and some doses were not licensed. Prospective RCTs would be necessary to establish the relative advantages and disadvantages of the two therapies.

Conclusions

The clinical evidence available showed that glitazones reduce HbA_{1c} by approximately 1%, and are more effective at higher doses than at lower doses. There is significant weight gain associated with glitazone treatment. No peer-reviewed data were available on the long-term effects of giltazone use. No prospective RCTs were found comparing pioglitazone with rosiglitazone, but the available evidence showed that the two treatments had similar effects.

Chapter 4 Economic analysis

Overview of economic assessment

The aim of this chapter is to assess the costeffectiveness of both pioglitazone and rosiglitazone in the treatment of type 2 diabetes. The economic analysis includes a systematic review of the costeffectiveness literature relating to pioglitazone and rosiglitazone and a review of the economic analyses submitted to NICE by the sponsoring bodies, Takeda and GlaxoSmithKline.^{25,26} In addition, modelling literature concerning the treatment of type 2 diabetes mellitus is reviewed in order to determine the appropriateness of the Takeda and GSK economic models. Finally, the two economic models are compared and the differences between them discussed.

Pioglitazone and rosiglitazone are antihyperglycaemic drugs. They are both indicated only in oral combination treatment of type 2 diabetes mellitus in patients with insufficient glycaemic control despite maximal tolerated doses of oral monotherapy with either metformin or a sulfonylurea:

- in combination with metformin only in obese patients
- in combination with a sulfonylurea only in patients who show intolerance to metformin or for whom metformin is contraindicated.

In the original submission regarding glitazones, NICE recommended to the NHS that patients whose blood glucose levels were not controlled with either metformin or a sulfonylurea alone should be offered metformin and a sulfonylurea in combination, unless there were reasons of contraindications or intolerability. Patients should be offered a glitazone in combination with metformin or sulfonylurea (as an alternative to injected insulin) if they are unable to take metformin and sulfonylurea as a combination therapy, or if their blood glucose level remains high despite adequate trial of this combination treatment. The combination of glitazone and metformin is preferred to the combination of glitazone and sulfonylurea, particularly for obese patients. Glitazone plus sulfonylurea may be offered to patients who are unable to take metformin.

After the original submissions and the NICE guidance for the glitazones, issues have been raised by the drug companies concerning the treatment pathways followed. In particular, it was suggested that adding a glitazone after failure of metformin monotherapy is a better treatment strategy than adding a sulfonylurea. Economic evaluation for this comparison of scenarios has been provided by the drug companies in the current submission. Moreover, it was suggested that the NICE proposal of adding sulfonylurea after metformin monotherapy and then switching to glitazone if the combination therapy fails is not followed in practice by clinicians. The reason is that an immediate substitution of sulfonylurea with glitazone in combination therapy with metformin results in significant loss of glycaemic control. As a result of that, clinicians tend to proceed to triple therapy with metformin, sulfonylurea and glitazone for some time before progressing to a combination of metformin with glitazone. Triple therapy is neither licensed nor recommended by NICE. However, it is applied in practice and it is an issue that has been raised since the original submission. Economic evaluation has not been provided by either of the two pharmaceutical companies for this issue.

Methods

A systematic literature search was undertaken for economic assessments of pioglitazone and rosiglitazone. In addition to the searches conducted for clinical effectiveness, searches were conducted in the NHS Economic Evaluation Database (NHS EED) and the Office of Health Economics Health Economic Evaluation Database (OHE HEED) specifically to identify costeffectiveness literature (Appendix 3). This was supplemented by searches in MEDLINE for economic and quality of life literature relating to the costs of insulin therapy (refer to Appendix 4 for the methodological search filters used).

A broader topic search was undertaken for economic or model-based assessments within diabetes. This search was used to identify assessments that attempt to estimate the long-term impact of glucose-lowering treatments in type 2

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diabetes mellitus and that do not limit their scope to individual complications of diabetes. The purpose of this topic review was to generate classification criteria for the evaluation of submitted economic evidence. A generic proforma for the critical appraisal of modelling studies in health economics is used in systematically reviewing the studies identified. This is supplemented by a detailed review of the disease-specific factors within all modelling studies identified. Where possible, key outcomes are compared. The key outcomes reported within these studies are:

- mean lifetime risk of complication
- cost per life-year gained (LYG)
- years free from first significant complication
- estimated incidence of complication
- total lifetime costs for diabetes
- duration of stay in given health state.

Results of topic review for issues in health economic modelling of diabetes

The topic search for economic or model-based studies identified 81 studies relating to the treatment of complications associated with type 2 diabetes. Details of the studies are available from the authors. Five of the studies identified focused on glucose-controlling interventions and addressed multiple complications of diabetes;³⁹⁻⁴³ these studies are summarised in Table 12 and in the original appraisal of pioglitazone.⁴ Of the five published studies (five publications, four studies), three studies focused on type 2 diabetes mellitus; the remaining two studies focused on type 1 diabetes.

The key clinical events within these models of diabetes^{39–43} are:

- nephropathy
- retinopathy
- acute myocardial infarction (AMI)
- stroke
- amputation
- hypoglycaemia •
- ketoacidosis
- lactic acidosis.

Other important events held within the models

- known epidemiology
- effects of interventions
- death from non-specific causes.

Pioglitazone

Results of systematic search for economic studies of pioglitazone

There are no published studies investigating the health economics of pioglitazone or, indeed, of any other TZD. The only available economic evidence concerning pioglitazone is that obtained as part of the confidential submission by the

Author	Study design	Economic outcomes	Intervention type	Intervention	Scope
Eastman et al. ³⁹	Modelling	Cost-effectiveness and cost-utility	Glucose control	Conventional vs intensive therapy	Туре 2
Palmer et <i>al</i> . ⁴⁰	Modelling	Cost-effectiveness	Glucose control and screening	ACE inhibitors, conventional and intensive insulin therapy	Туре I
Vijan et al. ⁴²	Modelling	None	Glucose control	Hypothetical	Type 2
DCCT ⁴¹	Modelling/cost of illness	Cost-effectiveness and cost-utility	Glucose control	Conventional vs intensive therapy	Type I
Bagust et al. ⁴³	Modelling/burden of illness	Cost	Glucose control	Conventional vs intensive therapy, diabetics vs non-diabetics	Туре 2

TABLE 12 Summary of classification of studies reviewed

sponsoring body, Takeda UK.²⁶ Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report. It should be noted that the economic model provided in the Takeda update submission is the same as was reviewed in the original appraisal of pioglitazone.⁴ Therefore, the review that is reported here is taken directly from the original economic review of pioglitazone.⁴

Critical appraisal of the economic submission for pioglitazone

A structured proforma was used in the critical appraisal of the economic submission for pioglitazone.⁴⁴ To determine how the unpublished submission compares with published models in diabetes the same proforma was used to summarise the five studies identified in *Table 12*.^{39–43} A detailed discussion of some of the key factors is given below.

Statement of the problem

The study²⁶ contains a clear statement of the problem, that is to estimate the impact of "the effects on HbA_{1c} , total cholesterol and HDL cholesterol in Type 2 patients" of treatment with pioglitazone combination therapy (added to either metformin or sulfonylureas) compared with other combination therapies or changing to insulin.

The population of interest is defined as those people with type 2 diabetes whose blood glucose levels are poorly controlled with oral monotherapy with either metformin or sulfonylureas.

The comparator therapies are well defined insofar as intensive glucose control is now accepted as conventional therapy.

The study focuses on the possible lifetime clinical and economic outcomes. The key economic results are reported in terms of the cost per life year gained. The perspective of the analysis is on direct medical costs, with a specific focus on the UK NHS costs.

Clinical benefits are discounted at 1.5%, and costs are discounted at 6%.

Cohort information

One of the key distinctions between the models is the focus on either type 1 or type 2 diabetes mellitus; given the different natural epidemiology of onset the cohort data will vary considerably between models. It is known that type 1 disease has an earlier onset than type 2 disease. As the study by Palmer and colleagues⁴⁰ and the DCCT⁴¹ considered the cost implications of type 1 diabetes, the age range of patients in the cohort is significantly lower than that of patients used in the models proposed by Eastman and colleagues^{39,45} and Vijan and colleagues.⁴² The cohorts used in the models are described in *Table 13*.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The patient populations used in the models proposed by Eastman and colleagues³⁹ used 10,000 patients as a baseline cohort. From this, 30.5%, 21.7%, 17.7% and 30% were within the age groups 25-44, 45-54, 55-64 and 65-74 years, respectively. The cohort used in the model by the DCCT⁴¹ also uses 10,000 patients. Equal proportions of males and females were included, and patients were weighted by ethnicity. The mean age of model entry at clinical diagnosis of type 2 diabetes in the USA was 51 years. The patient population included in the two cohorts of patients in the model proposed by the DCCT⁴¹ consists of a sample of patients with type 1 diabetes in the USA who were considered eligible for enrolment in the DCCT (dependent on demographic and clinical characteristics). These two cohorts are classified as follows.

TABLE 13	Cohort in	formation	used in	the	published	models
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Authors	Disease type	Cohort age range (years)	Source of cohort information	No. of patients in cohort
Eastman et al. ³⁹	Туре 2	25–74	WESDR	10,000
Palmer et al. ⁴⁰	Type I	19 (mean)	Not stated	Not stated
DCCT ⁴¹	Type I	13–39 (2 cohorts)	WESDR	10,000
Vijan et al. ⁴²	Type 2	45–75 (assumed)	REP, WESDR	Not stated
Bagust et al. ⁴³	Type 2	Not stated	WESDR, UKPDS	Not stated

REP, Rochester Epidemiology Project, Minnesota; WESDR, Wisconsin Epidemiologic Study of Diabetic Retinopathy.

- Patients in the conventional treatment arm (primary cohort) had no experience of retinopathy or MA, and a duration of disease of between 1 and 5 years.
- Patients in the intensive treatment arm had minimal to moderate non-proliferative retinopathy, excreted less than 200 mg of albumin in the urine per day and had a duration of diabetes of between 1 and 15 years. It was assumed that approximately 17% of the US population would be eligible for enrolment.

It is important to note that the individual characteristics assigned to the patients in the DCCT model will differ considerably from those of the Eastman^{39,45} model, given the difference in disease type.

The patient population entering the model by Vijan and colleagues⁴² is assumed to have an age range of 45–75 years; however, this is not explicitly stated in the literature. Patients in the cohort were assumed to have no clinically detectable microvascular complications at the time of diagnosis of diabetes. Patients who present with complications are already declared to be at high risk and, therefore, should be considered for intensive control. The study by Palmer and colleagues⁴⁰ focuses on male type 1 diabetes mellitus patients in Switzerland. The age of this representative cohort was 19 years, as this is known to be the median age of onset of type 1 diabetes in Swiss males; this presents a limitation for the Palmer model. Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Model structure and scope

The models by Eastman and colleagues^{39,45} and the DCCT⁴¹ use a Monte Carlo technique. These two models take the form of a microsimulation Markov model with a 1-year cycle, whereby patients enter the simulation as individuals rather than as a cohort. Upon beginning the simulation, patients are assigned individual characteristics, weighted to the incident cases of clinically diagnosed patients with type 2 diabetes in the US population, within the eligible age range (type 1 diabetes in the DCCT study). The eligible age ranges used in these models are shown in cohort information given in *Table 14*.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The models by Eastman^{39,45} and the DCCT⁴¹ reflect the typical model structure as described above, incorporating the same three major complications associated with diabetes: neuropathy, nephropathy and retinopathy (*Table 15*). The overall structures of the models proposed by the DCCT and Eastman are similar to those of a model used to predict retinopathy.⁴⁶

Authors	Type of simulation	Type of model	Decision analysis	Monte Carlo
Eastman et al. ^{39,45}	Micro	Markov	Yes	Yes
Palmer et al. ⁴⁰	Micro (assumed)	Markov	Yes	Not stated
Vijan et al. ⁴²	Micro (assumed)	Markov	Yes	Not stated
DCCT ⁴¹	Micro	Markov	Yes	Yes
Bagust et al. ⁴³	None	Markov	Yes	No

TABLE 14	Types of modelling	used by the stu	dies
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TABLE 15	Complications included in the models	

Complication	Eastman et al. ^{39,45}	Palmer et al. ⁴⁰	Vijan et <i>a</i> l. ⁴²	DCCT ⁴¹	Bagust et al. ⁴³
Retinopathy	*	*	*	*	*
Neuropathy	*	*		*	*
Nephropathy	*	*	*	*	*
Heart disease	*	*		*	*
Stroke		*			*
Hypoglycaemia		*			*
Ketoacidosis		*			
Lactic acidosis					

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The model proposed by Vijan and colleagues⁴² has by far the most limited scope. It calculates the risks of developing blindness and end-stage renal disease (ESRD) for patients at different ages of diabetes onset and different levels of glycaemic control. However, the model by Vijan excludes any complication-specific mortality and, therefore, considers only early-stage disease. Furthermore, although it is recognised that those patients at high risk of blindness and renal disease (as included in the model) have, in turn, a higher risk of developing neuropathy, Vijan and colleagues do not include amputation and neuropathy in the model, as a result of an apparent lack of evidence. The model by Vijan calculates the risks of developing blindness and ESRD for patients at different ages of diabetes onset and different levels of glycaemic control.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The data held within these two aspects of the models proposed by Eastman and colleagues and the DCCT are consistent with the known epidemiology of type 2 diabetes in the USA. Despite the DCCT⁴¹ model's difference in focus from that of Eastman and colleagues^{39,45} (from type 2 to type 1 diabetes), the underlying structure appears to be identical, as the complications represented as submodels are common to both type 1 and type 2 diabetes. The model proposed by Palmer and colleagues⁴⁰ has many similarities to those proposed by Eastman^{39,45} and the DCCT⁴¹ in terms of underlying model structure, but Palmer⁴⁰ and Takeda²⁶ simulate the disease with markedly wider scope.

The models presented by Eastman and colleagues^{39,45} and the DCCT⁴¹ include three complications, a heart disease submodel and a mortality submodel, which, together, are believed to reflect the natural history of vascular and neurological complications. Within the models proposed by Eastman, the DCCT, Vijan, Palmer and Takeda²⁶ there is no set sequence by which patients may experience the complications included in the model; rather, the submodels run in parallel. Another element of commonality is that all of the submodels for each study are assumed to be

mutually exclusive, and therefore no compound health states are included in the model.

It can be concluded here that, in terms of model scope, the model proposed by Takeda²⁶ is at least as comprehensive as the other models identified by the literature search. *Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.*

Structure of submodels

Neuropathy

The study by Eastman and colleagues^{39,45} has a major strength in the explicit statements of hazard rates and transition probabilities: these are not included by other authors. Apart from slight disparities in terms of clinical definitions of health states, the DCCT⁴¹ model is identical in structure to that of Eastman and colleagues. In the models proposed by Eastman and colleagues^{39,45} and the DCCT,⁴¹ patients may be in one of three disease states, through which they follow a consecutive progression. The amputation submodels presented by Palmer and co-workers⁴⁰ and Takeda²⁶ are largely identical in structure, each including five health states. These are, therefore, similar to the neuropathy structures proposed by Eastman^{39,45} and the DCCT,⁴¹ whereby the patient begins the simulation with no history of amputation. However, the submodels proposed by Palmer⁴⁰ and Takeda²⁶ also include non-specific mortality.

In the neuropathy submodel proposed by Eastman and colleagues,^{39,45} adjustments are made according to ethnicity. Patients enter the submodel with no neuropathy is present. At the time of clinical diagnosis, the prevalence of significant diabetic neuropathy was approximately 3.5% according to the National Health and Nutrition Examination Survey (NHANES) II in the Eastman^{39,45} model. The hazard rate allocated to this event predicted a cumulative incidence for symptomatic neuropathy of 13% 8 years after diagnosis, which is reflected in the results of the REP. The next health state in the submodel is that of first lower extremity amputation, which was also estimated by the REP. Hazard rates used in the progression to this state are conditional on the duration of diabetes onset. Similarly, hazard rates were calculated from the cumulative incidence of first lower extremity amputation (LEA) and, later, made conditional on the duration of diabetes represented in the model. Subsequent to experiencing a first LEA, patients are at a higher risk of a second LEA.

Palmer and colleagues⁴⁰ suggest that patients with type 1 diabetes are 14 times more at risk of nontraumatic LEA than a non-diabetic population. *Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.* The probability of amputation was assumed to decrease by 41% with intensive therapy within the Palmer⁴⁰ submodel. The annual incidence of second LEA is four times higher than that of the first. It is also known that patients have a higher risk of death once the first LEA has occurred.

The model proposed by Vijan and co-workers⁴² does not include a neuropathy submodel.

Nephropathy

The epidemiology of type 2 diabetes indicates that 25–50% of patients develop microalbuminuria.^{39,45} The nephropathy submodel contains four disease states within the DCCT⁴¹ and Eastman^{39,45} submodels. According to these models, patients progress from one state to the next without missing a step. Upon entering the model, patients begin in the disease state 'no nephropathy'. Using back-data from the WESDR, a baseline prevalence of microalbuminuria of 11.5% is assumed within the Eastman submodel. Adjustments are made again for hazard rates in ethnic minorities.

Patients progress from the initial health state to microalbuminuria; the respective hazard rate is universal for all durations of disease. This hazard rate is again dependent on ethnicity. The subsequent health state sees the patient progress to proteinuria. The hazard rate for this progression is universal for all durations of diabetes. The progression from proteinuria to ESRD is dependent on the duration of diabetes; the hazard rates for this progression are 0.0042, 0.0385 and 0.074 for durations of 1–11 years, 12–20 years and over 21 years, respectively. It should be noted that the clinical definitions for these two states differ among the various studies.

It is important to note that the intermediate disease states are referred to differently in the DCCT⁴¹ model and the model presented by Eastman and colleagues;^{39,45} hence, the differences between definitions may suggest differences in the internal structures of the submodels. The nephropathy submodel proposed by Vijan and colleagues⁴² is largely similar to the model proposed by the DCCT and Eastman, yet also includes a non-complication-specific mortality state. The structure of the submodel described by Palmer and colleagues⁴⁰ differs slightly from those models used by other authors in that they include ten health states. The four health states included in other submodels are also included here, and an additional six health states are also included. From ESRD, the final nephropathy health state in all submodels previously analysed, the model also includes the treatment of ESRD (e.g. haemodialysis) and includes a health state for ESRD-specific mortality. This represents a considerable amount of extra detail included in these models. This suggests a closer reflection of the complication in the models proposed by Palmer and co-workers.⁴⁰

Clearly, the transition probabilities may differ between each of the models proposed by various authors.

Retinopathy

As with the other submodels proposed by the DCCT,⁴¹ the retinopathy submodel is largely identical to that of Eastman and colleagues 39,45 in terms of structure, despite slightly different clinical definitions of health states. The epidemiology of the disease shows that most people with type 1 diabetes develop non-proliferative retinopathy and 62% develop proliferative retinopathy, so this information was used in the calculation of the transition probabilities in the model presented by the DCCT. The retinopathy submodel presented by Eastman and the DCCT includes five health states. This is the same for the submodels proposed by Vijan⁴² and Palmer⁴⁰ except that the macular oedema (MO) state is omitted and a noncomplication-specific mortality state is included. Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

There are, however, two different pathways through which patients may progress within the models by Eastman and the DCCT. The hazard rates derived by Eastman for the progression of one state to the next were again obtained from the WESDR. Patients begin in the disease state 'no retinopathy', with the exception of 20% of patients who, at the time of clinical diagnosis of diabetes, were assumed to have background retinopathy. The hazard rate of progression from 'no retinopathy' to 'background retinopathy' is dependent on the duration of the disease in the Eastman^{39,45} model.

From the 'background retinopathy' disease state, patients may progress either to the subsequent

disease state (proliferative retinopathy) or to significant MO. The hazard rates of progression from proliferative retinopathy to severe vision loss, and from MO to blindness, are conditional on whether or not the patient receives treatment for the disease state. The hazard rates for the progression to either of these states are also conditional on the duration of diabetes. The health state of MO is excluded from the model. Despite the authors mentioning this disease state in the literature, no explanation is provided explaining why this important factor is not included in the model. This is clearly a limitation of the model by Eastman and colleagues.^{39,45}

Eastman and co-workers make adjustments for ethnic minorities who are more at risk of background retinopathy, MO and proliferative retinopathy. As a result of insufficient data, the assumption is made that Asian Americans have the same risk as non-Hispanic white people. The final stage given either pathway is severe loss of vision, whereby vision is less than 20% of the better eye.

Vijan and colleagues⁴² used data derived from the DCCT⁴¹ to establish early rates of progression; these were used as base-case analyses. The incidence and progression of retinopathy were defined as in the DCCT.

Other complications associated with diabetes mellitus

The following complications are included only by Palmer and colleagues⁴⁰ and Takeda²⁶ (with the exception of the inclusion of CVD in the model proposed by Eastman^{39,45}). This is an advantage as it results in the model being considerably wider in scope, hence providing a truer representation of the complications encountered by diabetes patients.

Heart disease

It should first be noted that Eastman and colleagues^{39,45} do include a CVD submodel. Within this submodel the assumption is made that 50% of patients have CVD, as the disease accounts for 50% of the deaths in patients with diabetes-related ESRD.

Palmer and co-workers⁴⁰ state that, as with the probability of LEA, a patient's probability of developing AMI is dependent on previous heart conditions, and demographic and clinical factors. According to Palmer and colleagues, 6–10% of patients having first AMI die immediately,

dependent on age and gender. Patients with type 1 diabetes are at two to four times higher risk of developing AMI than the non-diabetic population.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Stroke

The stroke submodel proposed by Palmer and colleagues⁴⁰ suggests that from having no history of stroke, patients progress to first stroke before moving either to death or to recurrent stroke and then death. At any point in the model, the patient may die of a non-specific mortality. *Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.*

Palmer and colleagues⁴⁰ suggest that the known epidemiology of diabetes is that patients are at twice the risk of stroke in comparison with the non-diabetic population. The likelihood of experiencing a stroke is dependent on demographic and clinical factors. Approximately 16% of patients die in hospital.

Hypoglycaemia

It is known that hypoglycaemia is common and, ultimately, an important recurrent complication for diabetes patients, yet it is not included in the models proposed by Eastman,^{39,45} the DCCT⁴¹ or Vijan.⁴² Owing to the rapid rate at which the patient experiences the effects of hypoglycaemia, non-complication-specific death is not included in the submodel proposed by Palmer and colleagues.⁴⁰ The progression of hypoglycaemia is simple: patients enter the submodel without having experienced a hypoglycaemic event. Nonserious events are not included in the model. The patient may then progress to experience an event whereby he or she requires medical assistance. From this point, the patient either recovers and reverts back to the initial health state or progresses to the hypoglycaemia-specific death state. The patient cannot remain in the second health state described above, as ongoing hypoglycaemia is regarded as fatal. The Palmer model⁴⁰ assumes a case fatality probability of 0.0001.

Ketoacidosis

Ketoacidosis is a complication that is generally specific to type 1 diabetes and is included in the model by Palmer and colleagues.⁴⁰ Information from the sponsor's submission was submitted in confidence to

NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Lactic acidosis

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Weight gain

None of the models considered included the potential impact of weight gain on mortality. As discussed in the review of the clinical effectiveness of pioglitazone, pioglitazone has been shown to have a marked and progressive effect in increasing body weight. Although the effect of obesity on mortality independent of the effect of lipid concentrations is controversial, there remains the possibility that the increase in body weight due to pioglitazone usage may have an adverse impact on long-term mortality. It is a key shortcoming that these effects are not included in the models.

Mortality

The models by Eastman and colleagues^{39,45} and the DCCT⁴¹ include a separate submodel which simulates mortality of patients. Each year the mortality model defines whether the individual survives or not. In the model proposed by Palmer and co-workers,⁴⁰ mortality is not contained in a separate model, but is approached in the various submodels of complications. It is important to note that the model proposed by Vijan⁴² only includes early-stage disease and does not include a complication-specific mortality element. Eastman and colleagues^{39,45} use life tables to obtain the typical life expectancy of a non-diabetic patient; this is then multiplied by a factor of 2.75 to reflect the life expectancy of a patient with type 2 diabetes. The model proposed by the DCCT⁴¹ uses data from the US Department of Vital Statistics to obtain typical survival rates. It is not made clear how the model proposed by Palmer and colleagues⁴⁰ apportions rates of mortality.

Cost aspects

The costs included in each of the models are approached in different ways. The inevitable result of this is a severe difficulty in making close comparisons between the costs used in each of the models. The Vijan model⁴² addresses the risks and benefits associated with improved glycaemic control yet does not directly evaluate costs; the motive behind this is due to costs of decreasing HbA_{1c}-levels not being well defined for type 2 diabetes. The models proposed by the DCCT,⁴¹ Eastman^{39,45} and Palmer⁴⁰ include all direct medical costs (e.g. inpatient and outpatient care, laboratory tests and medical equipment), yet analyse these costs differently. Costs are in 1994 US dollars, with the exception of Palmer,⁴⁰ in which they are described in 1996 Swiss francs. Only Eastman and colleagues^{39,45} provide unit costs used. The Eastman models include costs of screening, treatment and disability. The sources of these data were the DCCT⁴¹ study, published literature and Medicare reimbursements.

The Palmer model⁴⁰ includes direct costs and takes a third-party payer perspective. *Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.* Palmer and colleagues⁴⁰ found that the cost driver in the model was the cost of renal failure, which is substantially reduced with the addition of screening for microalbuminuria and ACE therapy.

TABLE 16 Summary costs of implications and events used in the model proposed by Takeda

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The main costing areas included within the models are:

- screening costs
- treatment costs
- disability costs.

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Eastman and co-workers and the DCCT simply stated unit costs and costs per year.

A major issue in the comparison of relevant costs of complications contained in the literature is the issue of healthcare setting. Different settings have different implications for healthcare resource intensity and usage. This makes it difficult to define the costs of being in a particular health state. As largely the same sources have been used to derive data in the various models, one would expect the outcomes to be similar, yet evidently this is not the case.

Clinical outcomes

The results from the various studies are as follows.

- Vijan:⁴² the primary outcome measure used is lifetime risk. A reduction in HbA_{1c} levels from 9% to 7% for patients with onset before 50 years of age results in a 2.3% point decrease in lifetime risk of blindness due to retinopathy. A decrease from 11% to 9% in a patient with onset before 50 years of age results in a 5.3% decrease in blindness risk. The same relationship holds for the ESRD submodel. The conclusions drawn are that substantially greater effectiveness is achieved in moving from poor to moderate glycaemic control than from moderate to normal control.
- **Palmer**:⁴⁰ the primary outcome measures used in this study are mean total lifetime costs per patient, mean life expectancy and costeffectiveness (measured in terms of costs per LYG). Intensive therapy increased LYG but also increased total lifetime costs.
- **DCCT**:⁴¹ the primary outcome measure used is LYGs, but the study also tracked sight years, ESRD-free years, amputation-free years and quality-adjusted life-years (QALYs). QALY values were 0.69 for blindness, 0.61 for ESRD (Lawrence), 0.8 for LEA, 0 for death and 1 for all other health states. The incremental cost per LYG was found to be US\$28,661.

• **Eastman**:^{39,45} the primary outcome used is incremental cost per QALY. The incremental cost per QALY of intensive treatment over conventional is US\$16,002. This study uses the same utility outcomes as those used in the DCCT⁴¹ study, derived from largely the same sources. Maintaining an HbA_{1c} value of 7.2% is predicted to reduce the cumulative incidence of blindness, ESRD and LEA by 72%, 87% and 67%, respectively. Total estimated life expectancy is increased by 1.39 years.

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Although each of the models presents its findings in a different format, they all track the average increase in life expectancy from conventional to intensive therapy, with the exception of the model proposed by the DCCT.⁴¹ Despite this homogeneity, there are clearly differences between the findings due to the differences in the ages of diabetes onset and also the type of disease. It is likely that the increase in average life expectancy is higher for treatment of type 1 diabetes as the onset of the disease is earlier and hence the competing risks are less. Another major impact is in the definition of comparator therapies. The earlier studies focus on the comparison of intensive

Author	Outcome measure	Increase in life expectancy as a result of intensive therapy (years)	Target of intensive glycaemic therapy	Comments
Vijan et <i>al</i> . ⁴²	Average increase in life expectancy	1.3	2% decrease in HbA _{1c} level (actual start level not specified)	Age at onset: 45 years
DCCT ⁴¹	Mean years free from first significant complication	15.3	NA	None stated
Eastman et al. ^{39,45}	Average increase in life expectancy	3	A decrease in HbA _{1c} level of 2.8% (from 10% to 7.2%)	Assumes non-CVD mortality among diabetic population
Palmer et al. ⁴⁰	Average increase in life expectancy	7.4	Not stated	Risk of AMI and stroke reduced by 41%
				Conventional vs screening + intensive

TABLE 17 Average increases in life expectancy as described by the studies

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glucose control therapies with non-intensive therapies. Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The average life expectancies are shown in Table 17.

Utility scores

There has been considerable debate on the issue of utility scores for patients with diabetes. Of the four studies evaluated, only the DCCT study⁴¹ and Eastman and co-workers^{39,45} make an attempt to apportion quality of life scores to end-stage complications associated with type 2 diabetes. Both studies use identical scores for the end-stage diseases. This reflects the paucity of data in this area, as both studies use the same source, rather than a high level of certainty in the values used. The study by the DCCT⁴¹ makes the assumption that, as compound health states are not incorporated in the model, where patients reach the end stage in two or more of the complications, they assume the lower utility of the complications that they have experienced. For example, where a patient reaches blindness and LEA, the quality of life score used is 0.61, the score for blindness. This implies that the models are likely to underestimate the quality of life impact for an individual, as patients who are blind and have had an amputation would clearly prefer not to be blind; there should be a difference in quality of life between these two scenarios. A suggested (and more realistic) alternative would be to multiply the two utility scores, so, for example, in the compound health state described above the resulting utility score would be $0.80 \times 0.69 = 0.552$.

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Validation of the Takeda model²⁶

TABLE 18 Comparison of incidence estimates for the model against UKPDS data

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Assumptions made within the Takeda model²⁶

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Conclusions on the critical appraisal of the Takeda model²⁶

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Key economic results for pioglitazone

TABLE 19 Baseline cost per LYG for pioglitazone

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Impact of structural assumptions within the model

TABLE 20 Economic impact of assumptions in the Takeda model

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Conclusions on the health economics of treatment with pioglitazone in type 2 diabetes

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Rosiglitazone

Results of systematic search for economic studies of rosiglitazone

There are no published studies investigating the health economics of rosiglitazone or, indeed, of any other TZD. The only available economic evidence concerning rosiglitazone is that obtained as part of the confidential submission by the sponsoring body, GSK.²⁵ This confidential economic submission is based closely on a published economic model by Bagust and colleagues.⁴³

Critical appraisal of the economic submission for rosiglitazone

A structured proforma has been used in the critical appraisal of the economic submission for rosiglitazone.⁴⁴ A detailed discussion of some of the key factors is given below.

The GSK cost-effectiveness model

The model provided by GSK has two submodels, a metabolic model and an economic model. The metabolic model uses treatment-specific values for insulin sensitivity and β -cell function to predict FPG, which is then converted into HbA_{1c}. The economic model then uses the predicted value of HbA_{1c} to estimate the future incidence of long-term complications and mortality for different treatment strategies.

These two models are linked together. In particular, the metabolic model represents the natural interrelationship of clinical metabolic variables and their normal progression in type 2 diabetes. The basic variables that drive this model are insulin sensitivity and β -cell function. The metabolic model provides the progression of glucose levels (measured in HbA_{1c}) through time for each one of the available treatment pathways followed. In addition, it presents the progression of other secondary clinical variables, such as body mass index (BMI), blood pressure, triglycerides and cholesterol. The outputs of the metabolic model are then used as inputs to the economic model. The HbA_{1c} levels mainly contribute to the calculation of microvascular complications in the economic model. The values of systolic blood pressure (SBP) and of the ratio total cholesterol to HDL cholesterol are averaged per age group and are used in the economic model, to estimate the CVD death rates. The economic model is a Markov model for the long-term type 2 complications of newly diagnosed patients. Data from well-attested published models, as well as very recent clinical results, are used to estimate mortality rates, transition probabilities and the number of patients in each state of the major complications of type 2 diabetes. Moreover, the model includes a very detailed module for the calculation of healthcare costs. Its main outputs are cumulative QALYs, lifeyears and health costs by time from diagnosis, and lifetime healthcare costs.

Statement of the problem

The submission²⁵ contains a clear statement of the problem, that is to perform a cost–utility analysis of adding rosiglitazone to sulfonylurea or metformin compared with other combination therapies or changing to insulin.

GSK supports that there is strong clinical and economic evidence available for the effectiveness and toleration of rosiglitazone in the treatment of type 2 diabetes patients. Moreover, the company believes that there is ambiguity in the existing NICE guidance with respect to the positioning of rosiglitazone in the treatment pathway of type 2 diabetes.

Thus, GSK suggests that the Institute should carefully consider the following aspects when reviewing the current guidance for rosiglitazone:

- Patients with inadequate blood glucose control on metformin, who are obese, should be considered for rosiglitazone combination therapy, as an alternative to sulfonylurea.
- Patients with inadequate blood glucose control on sulfonylurea, who are unable to take metformin, should be considered for rosiglitazone combination therapy.
- Patients treated with an established metformin and sulfonylurea combination should not subsequently be offered a switch from one of these components to glitazone therapy.

The population of interest in the study is defined as those people with type 2 diabetes whose blood glucose levels are poorly controlled with oral monotherapy with either metformin or sulfonylureas.

The study focuses on the possible lifetime clinical and economic outcomes. The key economic results are reported in terms of the cost per QALY.

The perspective of the analysis is on direct medical costs, with a specific focus on the UK NHS costs.

Clinical benefits are discounted at 1.5%, costs are discounted at 6%.

The GSK metabolic submodel

General description

The metabolic model represents the natural interrelationships of clinical metabolic variables and the normal progression of these variables over time in type 2 diabetes patients. *Figure 1* illustrates the relationships between key variables of the metabolic model.

Main output of metabolic model: HbA_{1c} progression

The main output of the metabolic model is the level of glycaemia (HbA_{1c}) during all years after

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FIGURE I Interrelationships of metabolic variables

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FIGURE 2 Description of the progression of glucose levels in the metabolic model

diagnosis if a specific treatment pathway is followed. *Figure 2* and the description that follows explain the methods used to derive the HbA_{1c} values for each treatment pathway.

Secondary outputs of metabolic model: weight, blood pressure and lipids progression

In addition to the glycaemic progression in type 2 diabetes, the metabolic model provides estimates for other interrelated clinical variables.

Using information for males and females with diabetes, separate regression models were developed from the NHANES3 data set²⁵ to predict each one of the following clinical variables:

- BMI
- SBP
- DBP
- total cholesterol
- LDL cholesterol
- triglycerides
- HDL cholesterol.

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Calibration of metabolic model: treatment effects

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To estimate the effect that each treatment or combination of treatments has on insulin sensitivity and β -cell capacity, GSK has used data from the following sources:

- UKPDS results
- published papers detailing trials
- unpublished SKB trial reports.

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TABLE 21 Additive and multiplicative effects of the different regimens on the metabolic variables

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Parameter values and assumptions

The underlying population profile assumes 1000 patients, newly diagnosed with type 2 diabetes, based on the numbers expected from a population structure equivalent to the European standard population.

The insulin sensitivity and β -cell function at diagnosis are assumed to be the same for males and females.

The user can choose the percentage of people in each ethnic group. In the central estimate, the population consists of 100% white people. This proportion is not representative of the British population. However, the company has been conservative in this assumption, as it has assumed only white people, who have a lower risk of progressing to microvascular states than do black people.

The assumptions made for the rest of the key parameter values are presented in *Table 22*.

The prevalence of smoking for each age group of diabetic males and females is taken from Eastman's model^{39,45} and is presented in *Table 23*.

From Eastman's paper, it was found that the prevalence values are for US people with diabetes and are taken from Cowie and Harris⁴⁷ and Fujimoto.⁴⁸

The initial conditions (baselines) for BMI, insulin sensitivity and effective β -cell function are defined as follows.

• The BMI is assumed to be 25 for lean patients and 31 for overweight patients.

TABLE 22	Key parameter	values in the	GSK model
----------	---------------	---------------	-----------

Parameter values	Males	Females
BMI age offset	5	10
BMI peak offset	-0.7	-1
BMI excess for new diabetics	I	I
Initial carbohydrate intake: diabetics	1.43	1.43
Insulin: BMI conversion factor	0.08	0.08
BMI: sensitivity conversion factor	0.035	0.035
Carbohydrate: BMI conversion factor	0.3	0.3
Carbohydrate: plasma glucose conversion factor	0.5	0.5
Baseline output calibration adjustment for BMI: diabetics	5	I

TABLE 23 Prevalence of smoking

Age group (years)	Diabetic males (%)	Diabetic females (%)
<25-44	40.70	25.30
45-64	27.00	22.00
65–90+	13.20	11.8

- The insulin sensitivity level [homeostasis model assessment (HOMA)% S] is set to 40 for lean patients and 30 for overweight patients.
- The β-cell function level (HOMA % B) is set to 40 for lean patients and 50 for overweight patients.

Therapy scenarios

The scenarios examined in the model are shown in *Table 24*.

The model only allows patients to receive one of the four following types of therapy during any half-yearly period:

	1.	1	•
٠	diet	and	exercise

- monotherapy oral hypoglycaemic agents
- combination oral hypoglycaemic agents
- insulin-based therapy.

The switching threshold between treatments is fixed for each scenario, rather than being dependent on the treatment that patients receive. The threshold has been set to: FPG = 9.3 mmol/l(equivalent to HbA_{1c} = 7.5%, using the developed regression model).

There is some consensus that the threshold used by clinicians to change treatments is dependent on the efficacy of the specific treatment. Thus, the threshold used in practice is not the same for all the types of treatment (drugs).

Sensitivity analysis on the switching threshold of HbA_{1c} was performed (see later section in this report). However, a different threshold could not be applied for each treatment, as the model allows

Scenario	Weight	Step I	Step 2	Step 3	Step 4	Step !
I	Normal	D&E	S	S + M	I	
2	Obese	D&E	М	M + S	I	
3	Normal	D&E	S	S + M	S + R (4 mg)	1
4	Obese	D&E	М	M + S	M + R(4 mg)	1
5	Obese	D&E	М	M + S	S + R (4 mg)	1
6	Normal	D&E	S	S + R (4 mg)	I	
7	Obese	D&E	М	M + R(4 mg)	I	
8	Obese	D&E	М	I		
9	Normal	D&E	S	I		
0	Obese	D&E	М	M + S	M + R (8 mg)	1
1	Obese	D&E	М	M + R (8 mg)	l	
2	Obese	D&E	S	S + R (4 mg)	I	
3	Obese	D&E	S	I S		

TABLE 24 Treatment pathways

only one switching threshold, which remains constant for all the drugs in the treatment pathway.

Drug and treatment costs

It is assumed that patients receive fixed dosage and that no drug titration takes place. Moreover, there is a fixed cost for dispensing each prescription.

The daily drug costs used in the model are presented in *Table 25*.

These costs were derived from BNF 42.⁴⁹ The drug costs used in the submission model are identical to the ones reported in BNF 42.

The total annual treatment costs are calculated, which, according to the regimen followed, include:

- **urine testing costs** (only for insulin-treated patients): the annual cost is £64. The model assumes 4 contacts per annum of specialist advice
- **glucose testing costs** (consumables): the annual cost is £34.31 for all of the regimens, except for insulin treatment, for which it is £102.94
- **prescribing costs** (drugs and dispensing): the annual cost for drugs is calculated using the prices provided in *Table 25*. The annual cost for dispensing is £3.90 (4×0.975) for all of the regimens, except for insulin treatment, for which it is £5.85 (6×0.975).

No source was provided in the GSK submission concerning the treatment costs.

The total annual cost for each treatment is calculated using the above data for the drug and treatment costs (*Table 26*).

Long-term progression of glycaemia (HbA_{1c})

Information from the sponsor's submission was submitted in confidence to NICE. This information was made

TABLE 26 Annual treatment cost for each drug

Treatment	Cost	
D&E	£34.31	
S	£88.30	
M	£78.91	
R (4 mg)	£366.10	
R (8 mg)	£732.19	
1	£611.17	

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FIGURE 3 Relationship between insulin sensitivity/ β cells and HbA_{1c} as seen in the metabolic model

available to the NICE Appraisals Committee but has been removed from this version of the report.

Belfast diet study

The Belfast Diet Study⁵⁰ is a 10-year prospective natural history study of 432 newly diagnosed diabetic patients aged 40–69 years, which was undertaken to assess the effect of intensive dietary management. The results of the study demonstrate that patients continuing on diet alone for the first 10 years after diagnosis have a small but progressive rise in FPG, which is associated with an equally progressive fall in β -cell function, but not with a change in either obesity or insulin sensitivity. GSK's model is in accordance with these results.

Furthermore, the Belfast study indicated that failure of diet therapy within the first 10 years is related to higher rates of glucose rise and β -cell decline. This failure occurs earlier in patients with higher initial glucose concentration, lower initial β -cell function, lower age and, for subjects maintained on diet therapy alone for at least 6 months, greater obesity.

TABLE 25 Daily drug costs

Drug	Tablet size	No. per day	Daily cost
Glibenclamide	2.5 mg	4	£0.1371
Metformin	850 mg	3	£0.1114
Rosiglitazone	4 mg	I	£0.9500
Rosiglitazone (max. dose 8 mg)	4 mg	2	£1.9000
Insulin	20 IŬ	3	£0.9000
Needles for insulin		3	£0.0800

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Regression model: $HbA_{1c} = f(FPG)$

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TABLE 27 Diagnostics for the regression model: HbA_{1c} = f(FPG)

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HOMA/continuous infusion of glucose with model assessment (CIGMA) models and progression of glycaemia

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Output risk factor variables

As mentioned earlier, population data from NHANES3²⁵ were used to describe the relationship between risk factors and other demographic and metabolic variables. The regression models developed from the NHANES3 data set are used to estimate the levels of the following variables for each age group and for each time slot after diagnosis:

- BMI (this variable is additionally updated in the subsequent periods in line with the central BMI trajectory, which reflects changes in plasma insulin levels, calorific intake and direct treatment effect on weight)
- SBP
- DBP
- total cholesterol
- LDL cholesterol

- triglycerides
- HDL cholesterol.

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However, diagnostics and accuracy measures were not provided from the company concerning the models described above. Therefore, the reviewers were unable to validate those models.

The GSK economic submodel

A deterministic Markov model was used for the long-term type 2 diabetes complications of newly diagnosed patients. The model consists of a network of classical interconnected Markov chain modules, which all reconcile to a central survival module. It projects the natural progression of type 2 diabetes over time, to assess lifetime costs and complications, and it is a cohort study. Various antidiabetic agents can be introduced individually or in combination. Well-attested probabilities were used from epidemiological and clinical sources to estimate the proportions of patients likely to suffer from a range of common complications of diabetes over their remaining lifetime.

Cohort information

The cohort used in the model is described in *Table 28*.

Inputs

The basic parameters used in the model are the following:

Distribution of new patients developing type 2 diabetes by age and gender

- 40.97% males and 59.03% females
- Average age: 56.50 for males and 57.81 for females

These are used throughout the model as the patients alive at the outset.

TABLE 28	Cohort inforr	mation used	in the	published	models
----------	---------------	-------------	--------	-----------	--------

Author	Disease type	Cohort age range (years)	Source of cohort information	No. of patients in cohort
GSK model ²⁵	Туре 2	57 (mean)	UKPDS	1000

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Relative mortality multipliers for nephropathy, neuropathy, retinopathy, CVD/CHD and stroke states

The default values were determined interactively to reproduce prevalence and mortality figures in published literature. The microvascular multipliers were derived from a varianceminimisation heuristic to reconcile the overall mortality at each age in the mortality module with the implied mortality calculated for each complication module. The macrovascular multipliers were generated by a process of iterative recalibration.

Distribution of new patients in each state by condition, age and gender

These are used to calculate the progression of patient cohorts in each state in the model. The proportions are the ones used in Eastman's model.^{39,45} The baseline prevalence values included in the GSK model are identical to those used by Eastman.

Proportion of patients in each ethnic group

The model assumes 100% white patients. This is not representative of the British population.

Relative risk multipliers per ethnic group in developing each condition

These are taken from Eastman's model^{39,45} and are used to adjust the transition probabilities in each disease state. The risk multipliers are taken from Tull and Roseman,⁵¹ Ghodes⁵² and Stern and Mitchell,⁵³ and are used also by Eastman in the development of a model of complications for type 2 diabetes.

Glycaemic levels

- Standard HbA_{1c} (DCCT) = 10.00
- Average HbA_{1c} at diagnosis = 7.00

These are used to calculate the Eastman power function and the WESDR linear threshold used in the estimation of the transition probabilities. The HbA_{1c} level for standard care used in the GSK model is identical to the one provided from DCCT.

Glycaemic control parameters

These are the parameters for the Eastman power functions and WESDR linear functions. They are used to modify the rates of progression to the several morbid states according to changes in glycaemic control. The glycaemic control parameters that are included in the GSK model are identical to those used by Eastman.

Choice of glycaemic gradient

The user can choose the gradient model (Eastman power function or WESDR linear threshold) to be used for the progression of microvascular complications. The WESDR linear threshold is used in the central estimate in the submitted model.

Annual discounting rates

- Cost rate = 6.00%
- Outcome rate = 1.50%

These factors are calculated for each of the periods throughout the model. The method used is examined and considered appropriate.

Discounting offset from date of diagnosis

This factor specifies the number of years after which the discounting should start and is used when comparisons between scenarios only diverge after a fixed period. In the model, the factor is assumed to be zero, which means that the discounting starts from the first period.

Cohort size

The total size of the annual diagnosis cohort is assumed to be 1000.

Choice of CVD models

The user can choose the CVD risk model (Framingham⁵⁴ or UKPDS⁵⁵) to be used in the estimation of CVD event rates and mortality. The UKPDS-based risk equations are employed for the central estimate in the submitted model.

The following are taken into account by the model only when the Framingham risk equations are used to estimate cardiovascular events and death rates:

Relative reduction in CVD mortality rates from reduced HbA_{1c} for conventional and intensive treatment (Table 29)

These are used to calculate the combined mortality rates. The source for these is UKPDS 33.

Reduction in SBP and relative reduction in CVD mortality for conventional treatment and tight control (Table 30)

These are used to calculate the combined mortality rates. The source for these is UKPDS 38.³

Reduction in triglycerides and relative reduction in CVD mortality for conventional treatment and tight control (Table 31)

The source for these is the Veterans Affairs High-Density Lipoprotein Cholesterol Intervention Trial (VA-HIT) study.

Glycaemia					
Regimen	FPG limit	HbA _{lc} limit	Mortality reduction		
Conventional	15.0	9.25			
Tight control	6.0	5.90	-7.10%		
Target	9.3		-4.50%		

TABLE 29 CVD mortality reduction from reduced glucose

TABLE 30 CVD mortality reduction from reduced SBP

Systolic blood pressure					
Regimen	Mean	Mortality reduction			
Conventional Tight control Target	54 44 54	-28.70% 0.00%			

TABLE 31 CVD mortality reduction from reduced triglycerides

Triglycerides		
Regimen	Mean	Mortality reduction
Conventional	1.82	
Tight control	1.30	-22.00%
Target	1.87	0.00%

Variables transferred from the Metabolic model HbA_{1c}

The following estimates concerning HbA_{1c} levels are transferred from the metabolic model:

- mean HbA_{1c} levels at 6-month intervals
- regimen applicable to treatment in each 6-month period (diet and exercise, oral monotherapy, oral combination therapy, insulin-based therapy)
- average HbA_{1c} in 5-year periods
- average proportions of patients on each of the therapies in 5-year periods; these are used to calculate the cost of diabetes
- mean HbA_{1c} for microvascular (nephropathy, neuropathy and retinopathy) model periods; these are used in calculations in the respective microvascular submodels

• average annual treatment cost corresponding to the specific treatment used in each 6-month period.

These figures provide estimates for mean HbA_{1c}, which are used in the economic model, especially in the microvascular complication modules.

Body mass index, blood pressure and lipids

Mean values for the BMI, blood pressure and lipids are provided from the metabolic model. These are then used to drive the cardiovascular disease risk relationships in this model.

Mean values of the following for each subcohort by age (per 6 months) and gender are transferred from the metabolic model:

- BMI
- SBP
- DBP
- triglycerides
- HDL cholesterol
- total cholesterol
- LDL cholesterol.

The SBP and the ratio of total cholesterol over HDL cholesterol are then averaged per 5-year age group and they are used to drive the cardiovascular disease risk relationships for patients with type 2 diabetes.

Model structure and scope

The economic component of the GSK model (*Table 32*) is based closely on a published economic model by Bagust and colleagues.⁴³

The Bagust⁴³ model projects the natural progression of type 2 diabetes over time, to assess lifetime costs and co-morbidities. The model incorporates the major complications associated with diabetes: CVD, neuropathy, nephropathy, retinopathy and hypoglycaemia.

Within the model proposed by Bagust and colleagues, complications are presented as submodels, linked to the consequences of CVD and a mortality submodel, which, as a whole, forms the overall structure of the model.

TABLE 32	Types of modelling used by the studies
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Author	Type of simulation	Type of model	Decision analysis	Monte Carlo
GSK model	None	Markov	Yes	No

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It can be concluded here that, in terms of model scope, the model proposed by Bagust is at least as comprehensive as the other models identified by the literature search. By incorporating seven complications, CVD and a complication-specific mortality element within each relevant submodel, the underlying structure of the model provides a broad representation of the complications that a type 2 diabetes patient may experience.

The overall structure of the model included in the GSK submission is identical to the published model by Bagust.43 A deterministic Markov model was used for the long-term type 2 diabetes complications of newly diagnosed patients. The model consists of a network of classical interconnected Markov chain modules, which all reconcile to a central survival module. It projects the natural progression of type 2 diabetes over time, to assess lifetime costs and complications, and it is a cohort study. Various antidiabetic agents can be introduced individually or in combination. Well-attested probabilities were used from epidemiological and clinical sources to estimate the proportions of patients likely to suffer from a range of common complications of diabetes over their remaining lifetime.

Structure of submodels *CVD, CHD and stroke models*

The Bagust model⁴³ uses the Framingham multivariate risk model to estimate cardiovascular morbidity and mortality. The model estimates the prevalence of all CVD, CHD and history of stroke. Mortality rates due to cardiovascular causes were estimated using the Framingham equation. The CVD submodel in the GSK model is identical in structure to the published Bagust model.

UKPDS risk equations

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The UKPDS risk equation for CHD is implemented in the model as published in UKPDS 56.⁵⁵ The UKPDS effects on mortality are introduced as risk adjustments in the estimation of cardiovascular deaths, to reflect the benefit that tight control of hypertension and hyperglycaemia has in extended life. In order to integrate with the rest of the model, corresponding risk factors are estimated for CVD and stroke risks, based on scaling Framingham risk estimates for these conditions by factors dependent on the ratio between UKPDS and Framingham risk estimates for CHD.

It is assumed that the UKPDS risk evaluation can be applied incrementally to periods after diagnosis using current risk factor values. Moreover, CHD and stroke risks estimated by Framingham equations are assumed to increase in UKPDS in a similar manner to CHD risks.

The UKPDS 56⁵⁵ model coefficients are provided in the model. These are presented in *Table 33*.

Common factor values for males and females are calculated using the model coefficients, to simplify later calculations (males: 0.000005, and females: 0.000003). The values of the UKPDS 56 intermediate variable (*q*) are then calculated for males and females for each subcohort at diagnosis and at 5-year age points. The parametric model used is:

$$q = q_0 \beta_1^{age-55} \beta_2^{sex} \beta_3^{ac} \beta_4^{smok} \beta_5^{h-6.72}$$
$$\beta_6^{(sbp-135.7)/10} \beta_7^{\ln(lr)-1.59}$$

Thus, the probability (risk) of a CHD event over t years in a patient who has had diabetes for T years is estimated using the previous q calculations and the following formula:

$$R_T(t) = 1 - \exp\left\{-qd^T\left(\frac{1-d^t}{1-d}\right)\right\}$$

The above methods are used and thus the model provides estimates for the following 2.5/5-year period CHD risks for males and females for each subcohort. These estimates are then combined

TABLE 33	UKPDS	56 model	coefficients
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Parameter	Risk ratios for	Estimate
90	Intercept	0.0112
$\beta_{\rm I}$	l year of age at diagnosis	1.059
β_2	Female gender	0.525
β_3	Afro-Caribbean ethnicity	0.39
β_4	Smoking	1.35
β_5	1% increase in HbA _{1c}	1.183
β_6	10 mmHg increase in SBP	1.088
β_7	Unit increase in logarithm of lipid ratio	3.845
d	Each year increase in duration	1.078

with the Framingham⁵⁴ risk estimates in the model to derive CVD and stroke risks, as well as mortality rates for the following 2.5/5-year period for males and females for each subcohort.

Summarising, the model estimates the following 2.5/5-year period risks for males and females for each subcohort:

- CHD risks: UKPDS
- CVD risks: estimated UKPDS
- stroke risks: estimated UKPDS
- CVD annual mortality rate: estimated UKPDS
- CHD risk ratios: UKPDS/Framingham (these figures are descriptive only).

UKPDS 56

The UKPDS 56 study⁵⁵ provided a parametric model for predicting the risk of new CHD events in patients with type 2 diabetes. Unlike previously published models, the risk equation derived from this study has been specifically designed for people with type 2 diabetes, comes from a very large population and includes glycaemia, SBP and lipid levels as risk factors, in addition to age, gender, ethnic group, smoking status and time since diagnosis.

The UKPDS risk equation for CHD events is correctly used in the model included in the GSK submission. In particular, the parameter estimates are identical to the UKPDS ones, the common factors are properly generated, and the equation and risks are correctly calculated.

Framingham risk equations

The Framingham multivariate risk models are used to calculate the death rates and the number

of new cases of CHD, CVD and stroke in each model period and surviving cohort. Moreover, the numbers of new cases of CHD, CVD and stroke are estimated using also the UKPDS-based equations, to reflect the benefit that tight control of hyperglycaemia (not included in Framingham models) has in extended life. The UKPDS equations are used for the central estimate in the submitted model.

The estimates of events are calculated using multivariate models. Definitions of MI, stroke, CHD and CVD are those used in the formulation of the Framingham risk equations.

The Framingham⁵⁴ model coefficients are provided in the model. These are presented in *Table 34*.

The mean SBP, prevalence of smoking, total/HDL cholesterol ratio, proportion with ECG-LVH, age and gender are used to calculate the intermediate multifactors in the six Framingham risk equations for each age group of the model. These are assumed to be a linear function of the risk factors:

$$\mu = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \ldots + \beta_k x_k$$

The macrovascular events examined are the following:

- CVD event
- CVD death
- stroke
- MI
- CHD event
- CHD death.

Factor	CVD death	Stroke	МІ	CHD event	CHD death	CVD event
Beta 0	-5.0385	26.5116	11.4712	15.5305	11.2889	18.8144
Female	0.2243	0.2019	10.5109	28.4441	0.2332	-1.2460
In(Age)	8.2370	-2.3741	-0.7965	-1.4792	-0.9440	-1.8443
In(Age) ^ 2	-1.2109	0.0000	0.0000	0.0000	0.0000	0.0000
In(Age) [Fem. only]	0.0000	0.0000	-5.4216	-14.4588	0.0000	0.3668
In(Age) ^ 2 [Fem. only]	0.0000	0.0000	0.7101	1.8515	0.0000	0.0000
In(SBP)	-0.8383	-2.4643	-0.6623	-0.9119	-0.5880	-1.4032
Smoking	-0.1618	-0.3914	-0.2675	-0.2767	-0.1367	-0.3899
In(Chol/HDL)	-0.3493	-0.0229	-0.4277	-0.7181	-0.3448	-0.5390
Diabetic	-0.0833	-0.3087	-0.1534	-0.1759	-0.0474	-0.3036
Diabetic and Fem.	-0.2067	-0.2627	-0.1165	-0.1999	-0.2233	-0.1697
LVH	-0.2946	-0.2355	0.0000	-0.5865	-0.1237	-0.3362
LVH [Males only]	0.0000	0.0000	-0.1588	0.0000	0.0000	0.0000
Theta 0	0.8207	-0.4312	3.4064	0.9145	2.9851	0.6536
Theta I	-0.4346	0.0000	-0.8584	-0.2784	-0.9142	-0.2402

The calculated factors (μ) are then used to calculate the death rates and number of new cases of CHD, CVD and stroke in each model period by age, gender and subcohort. In particular, CVD death rates are estimated using the Framingham equation:

$$P(T > t) = 1 - \exp\left[-\exp\left(\frac{\ln(t) - \mu}{\sigma}\right)\right]$$

where T = time until the event of interest, and $\ln(\sigma) = \vartheta + \vartheta_1 \mu$ is considered to be a linear function of μ .

These rates are later used in the model to calculate the mortality rates. Moreover, the proportion of new cases of CVD, CHD and stroke developed in the previous 2.5/5-year period are estimated using the same model.

The user can choose between two CVD models. In particular, either the Framingham equations or the UKPDS-based equations can be used to estimate the macrovascular disease event rates. Thus, the proportion of new cases of CVD, CHD and stroke developed in the previous 5-year period are estimated using also the UKPDS multivariate risk equation. The model included in the GSK submission uses the UKPDS-based equations. It should be noted that Framingham equations are always used for non-diabetic CVD, CHD and stroke risks, as UKPDS estimates are only relevant to patients with type 2 diabetes.

Framingham heart study

The Framingham Heart Study⁵⁴ provided equations to predict risk for the following cardiovascular disease end-points: MI, CHD, death from CHD, stroke, CVD and death from CVD. These equations are based on measurements of several known risk factors and have indicated that a multifactor approach, one that takes into account all the risk factors, is probably the best strategy for the prevention of CHD. The parametric model used in this study is considered to have advantages, as it provides predictions for different lengths of time and its probabilities can be expressed in a straightforward way.

The Framingham risk equations for cardiovascular events are correctly used in the model included in the GSK submission. In particular, the parameter estimates are identical to the Framingham ones, the parametric model is properly applied, and the death rates and number of new cases are correctly calculated. Prevalence rates of CVD, CHD and stroke are calculated for males and females by age. These are later used to estimate the inpatient costs in the model.

Mortality

In the models proposed by Bagust and GSK, mortality is contained within a separate submodel model. The various components of the mortality module are combined and analysed by age, gender, and macrovascular status for the different subcohorts. Thus, the combined mortality rate for diabetics is estimated.

'CVD mortality' is based on Framingham (for nondiabetics) and UKPDS (for diabetics) risk equations, whereas 'other mortality' is derived from Office of Population, Censuses and Surveys (OPCS) mortality statistics in the UK combined with the prevalence of diagnosed diabetes from Diabetes in America. Moreover, the UKPDS effects on mortality are introduced as risk adjustments in the estimation of cardiovascular deaths, to reflect the benefit that tight control of hypertension and hyperglycaemia has in extended life. The estimated mortality components are combined with relative risk multipliers for each macrovascular status, to generate the 'combined mortality rates' for diabetics without CVD, and with CVD, CHD and stroke. Finally the 'revised annual death rates' are rates that are calculated by microvascular complication state module, based on mortality multipliers relative to non-diabetic mortality rates.

The morbidity components for the various diabetic causes are estimated combining data from two sources:

- OPCS mortality by cause statistics 1995 (used to determine rates for males and females for all causes, CVD and non-CVD)
- prevalence of diagnosed diabetes from *Diabetes in America* (2nd edition).⁵⁶

The following annual death rates are calculated:

- annual death rates for diabetics from non-CVD cases adjusted for underrecording by age and gender
- annual death rates from non-CVD cases in diabetes adjusted for independence (competing risks) by age and gender
- annual death rates for the major complication subgroups by age and gender: diabetes and renal, diabetes and neurovascular/PVD and other diabetes causes
- annual CVD death rates by age-group and gender. These are estimated for the nondiabetic population using the Framingham risk

equations⁵⁴ and for the diabetic population using the UKPDS-based risk equations⁵⁵

- annual CVD death rates by age group, gender and subcohort. These death rates are adjusted for tighter glycaemic, hypertensive and triglyceride control only when the Framingham CVD models are used. The glycaemic and hypertensive adjustment reflects the UKPDS⁵⁵ 'tight glycaemic/blood pressure control' regimen. The triglyceride adjustment is based on the results of the VA-HIT trial when triglyceride levels are reduced. All the above modifications are later used in the model to calculate the combined mortality rates
- annual death rates by ethnic group for diabetics with ESRD. The default values are those used by Eastman. A weighted average of death rates per age group is provided according to the percentage of people in each ethnic group. The proportion of ESRD deaths due to CVD is assumed to be 50% (this percentage is identical to the one used by Eastman^{39,45}).

The various components of mortality are combined and analysed by age, gender, macrovascular status and subcohort. Thus, the combined mortality rates for diabetics are estimated. These are mortality rates for diabetics 'without CVD', 'with CVD', 'with CHD' and 'with stroke'. All of these rates are the result of combination of previous mortality estimations and the last three are additionally adjusted by mortality multipliers for CVD, CHD and stroke, respectively. The combined mortality rates are used to drive the progression of patient cohorts. Moreover, they are used to derive the revised annual death rates.

The revised annual death rates are finally calculated by age, gender and microvascular complication state module. The rates appropriate to individual morbid states are based on risk multipliers relative to non-diabetic mortality rates. The revised death rates are used as inputs in calculating transition rates in the neuropathy, nephropathy and retinopathy models.

Progression of patient cohorts by age and survival

The Framingham⁵⁴ definitions of CVD, CHD and stroke are used to estimate the progression of patient cohorts by age and survival.

The combined mortality rates and the proportions of new cases of CVD, CHD and stroke are used to calculate the number of patients alive and dead by CVD, CHD and stroke status, by age, gender and subcohort. Then, the total number of patients alive or dead is estimated as the sum of patients in all the subcohorts.

The average number of patients alive per 5-year age group and by CVD, CHD and stroke is derived from the previous estimates. This is then discounted at the cost rate.

Survival functions and lifetime healthcare costs are provided for each diagnosis subcohort. The survival functions are the percentage of patients alive by age and time from diagnosis (number of patients alive in each age group divided by the number of patients at the outset). These functions are used to calculate the lifetime healthcare costs from diagnosis. The latter are later used in the model to estimate the costs.

Finally, the number of all, CHD and stroke patients alive by time from diagnosis is calculated. This is twice the average number of patients alive in each subcohort for each 2.5/5-year period following diagnosis for male and female patients. It is used later in the model to calculate the costs by time from diagnosis.

Neuropathy

Eastman's^{39,45} neuropathy module is redesigned to incorporate PVD and neuropathy in accordance with the pathogenesis of foot ulceration and amputation defined by Boulton.⁵⁷ The reformulation involved the identification of seven morbid states: no neuropathy, neuropathy, PVD, neuropathy and PVD, diabetic foot ulceration (DFU), first amputation and second amputation. The model uses linear functions with a WESDR threshold in the normal range of HbA_{1c} to show the influence of glycaemic levels on transition rates. (Alternatively, the user can use Eastman's power functions.) The duration of diabetes is divided into two phases: initial and regular. The model produces a set of 5-year transition probabilities across all seven states of neuropathy by diagnosis group for each of the phases.

The neuropathy model assumes seven states of disease:

- no neuropathy
- neuropathy
- PVD
- neuropathy and PVD
- DFU
- first amputation
- second amputation.

These states are mutually exclusive and in any year, the patient can either remain in the same state, progress to another state or die.

Figure 4 illustrates the neuropathy module.

A person will first develop either neuropathy or PVD. This can then lead to DFU and then first, second and further amputations. Published annual rates of transition to neuropathic states are used to derive 5-year state transition probabilities.

The elemental transition rates to neuropathy state and to second amputation are identical to those used in Eastman's model.^{39,45}

Two alternative sets of input transition rates can be used:

- transition rates used in Eastman's method
- transition rates by fitting the same data to a linear function with thresholds.

The transition probabilities used in the model increase as the time since diagnosis of diabetes increases. The rates are calculated separately for each of the 30 age/gender diagnosis groups.

The transition probabilities are the net transition rates for a cohort of cases. While an individual patient might move from one state to any other within a year, at the level of net transitions a cohort can only progress to at most the next most severe state within a year. It is not possible to move from neuropathy only to PVD.

The duration of diabetes is divided into two phases:

• **initial phase**: this phase begins at diagnosis and ends at first 5-year age marker. Since diagnoses

are made randomly in time, this phase lasts for 2.5 years on average

• **regular phase**: this phase follows the initial phase and lasts until death or the age of 100.

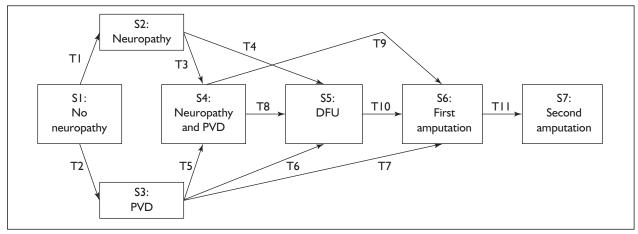
The model produces a set of 5-year transition probabilities across all seven states of neuropathy by diagnosis group. In addition, a modified set of transition rates is derived for the initial period from diagnosis of diabetes to the next model marker age.

Standard risks of developing significant neuropathy and PVD are provided. These are then adjusted by using one of two alternative methods, so that the influence of glycaemic levels on transitions is demonstrated.

The Eastman power function method can be used. In this method, the rates are derived by Eastman by fitting the WESDR study results to an exponential curve and applying it to the risk of developing neuropathy. The current HbA_{1c} is standardised by dividing it by the standard HbA_{1c} of 10.0 from DCCT data and raising it to a power parameter to fit the observations. Moreover, there is an option to adjust transition rates to neuropathy and PVD for ethnic mix variations.

Alternatively, the WESDR linear threshold can be used. In this method, the relationship between HbA_{1c} and the annual risk of developing neuropathy is assumed to be linear, but with a threshold value (HbA_{1c} of 6.0), below which the risk is zero. The basic annual transition rates are assumed to apply when HbA_{1c} = 10.0 and to increase linearly in proportion to the excess of HbA_{1c} above the threshold value (6.0).

In this model, the second method of the WESDR linear threshold is used for the central estimate.





Probabilities for 11 transitions are estimated using a combination of the published and HbA_{1c} adjusted rates included in the model. The methods applied to calculate the compound transition probabilities are correct.

Regular and initial probabilities are then calculated for each transition from one neuropathy state to the next. Residual proportions of patients in each disease state are estimated after transitions to more severe states and deaths.

A compound residue function is used to derive 5year state transition probabilities for every possible pair of start and end states. These are derived by assigning a probability to each possible path from the start state to the end state and summing these probabilities across all pairs.

Finally, separate transition rates are derived for the initial phase between diagnosis and the first 5-year age marker using the early phase annual transition probabilities. Since a patient can be diagnosed with type 2 diabetes in any of the 5 preceding years, the initial phase lasts for 2.5 years on average. For each combination of start state and end state the probability of this pair is calculated over 1, 2, 3, 4 and 5 years. The average of these five probabilities is the 'initial' transition matrix.

The 5-year neuropathy transition rates are used to calculate the numbers of patients in each state by age for each of the 30 age/gender diagnosis groups.

The newly diagnosed patients are distributed across the seven neuropathic states. The model assumes the same distribution for all 30 diagnosis groups. In particular, of the newly diagnosed diabetics:

- 96.5% have no neuropathy
- 3.5% have neuropathy.

The prevalence values used in the model are identical to those used by Eastman.^{39,45}

The number of patients in each disease state at each age is determined by applying the 5-year state transition rates (analysed in the previous section) to the number of patients in the starting state in the previous period and adding up all the possible combinations. The initial rates are applied to the first transition (from outset to the first 5-year age mark). The regular rates apply thereafter until death or the age of 100. In addition, the number of patients in each state is adjusted so that the total number of patients alive (dead) matches the number predicted by the mortality/CVD module. This is done by adjusting pro rata the number of patients in each state against the discrepancy in the overall number of patients alive.

The number of dead by state is calculated by applying the probability of death for that state to the number of patients in the state. These figures are again reconciled to match the values predicted in the mortality/CVD module.

The accumulated alive and dead by state at death derive directly from the two previous calculations. The total number of patients in each state at 5year age points is estimated by summing up all of the initiation groups. The model provides the number of patients alive by neuropathy state and age group for males, females and total. These figures are also discounted at the cost rate.

As far as the prevalence of neuropathy is concerned, the model provides estimations for:

- prevalence by age
- prevalence by time from diagnosis
- prevalent caseload by time from diagnosis
- cumulative incidence by time from diagnosis.

Finally, the model calculates the number of patients in each state (no neuropathy, neuropathy, PVD, neuropathy and PVD, DFU, first amputation, second amputation) by subcohort and time from diagnosis.

Nephropathy

The nephropathy module follows Eastman's formulation with four nephropathic states: no nephropathy, microalbuminuria, gross proteinuria (GPR) and ESRD. The transition probabilities are adjusted to reproduce the main WESDR findings on cumulative incidence of the morbid states. The model employs linear functions with a WESDR threshold in the normal range of HbA_{1c} to show the influence of glycaemic levels on transition rates. (Alternatively, the user can use Eastman's power functions.) The duration of diabetes is divided into four phases: initial, early, middle and late. The model produces a set of 5-year transition probabilities across all four states of nephropathy by diagnosis group for each one of the phases.

The nephropathy model assumes four states of disease:

- no nephropathy
- microalbuminuria

- GPR
- ESRD.

These states are mutually exclusive and in any year, the patient can remain in the same state, progress to the next more severe state or die.

Figure 5 illustrates the nephropathy module.

Published annual rates of transition to nephropathic states are used to derive 5-year state transition probabilities. Two alternative sets of input transition rates can be used:

- transition rates used in Eastman's method^{39,45}
- transition rates by fitting the same data to a linear function with thresholds.

The transition probabilities used in the model increase as the time since diagnosis of diabetes increases. The rates are calculated separately for each of the 30 age/gender diagnosis groups.

The transition probabilities are the net transition rates for a cohort of cases. Although an individual patient might move from one state to any other within a year, at the level of net transitions a cohort can only progress to at most the next most severe state within a year. However, within 5 years, progression through all four stages is possible.

The duration of diabetes is divided into four phases:

- **initial phase**: this phase begins at diagnosis and ends at the first 5-year age marker. Since diagnoses are made randomly in time, this phase lasts for 2.5 years on average
- **early phase**: this phase follows the initial phase and lasts for 10 years
- **middle phase**: this phase follows the early phase and lasts for 10 years
- **late phase**: this phase follows the middle phase and lasts until death or the age of 100.

The model produces a set of 5-year transition probabilities across all four states of nephropathy by diagnosis group for each of the early, middle and late periods since diagnosis. In addition, a modified set of early-phase transition rates is derived for the initial period from diagnosis of diabetes to the next model marker age.

WESDR-based standard transition rates are provided in the model. These are the annual transition rates derived by adjusting the values reported by Eastman. The rates have been rescaled to approximate outcomes reported by WESDR as part of the calibration of the model.

The transition rates used in GSK model are identical to the probabilities published in Eastman's model.

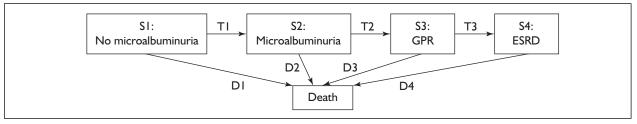
As mentioned earlier, the user can choose between two alternative sets of input transition rates, so that the influence of glycaemic levels on transitions is demonstrated.

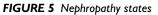
The Eastman power function method can be used. In this method, the rates are derived by Eastman by fitting the WESDR study results to an exponential curve and applying it to the risk of developing nephropathy. The current HbA_{1c} is standardised by dividing it by the standard HbA_{1c} of 10.0 from DCCT data and raising it to a power parameter to fit the observations. Moreover, there is an option to adjust transition rates to nephropathy and PVD for ethnic mix variations.

Alternatively, the WESDR linear threshold can be used. In this method, the relationship between HbA_{1c} and the annual risk of developing nephropathy is assumed to be linear, but with a threshold value (HbA_{1c} of 6.0), below which the risk is zero. The basic annual transition rates are assumed to apply when HbA_{1c} = 10.0 and to increase linearly in proportion to the excess of HbA_{1c} above the threshold value (6.0).

In this model, the second method of the WESDR linear threshold is used for the central estimate.

Regular probabilities are then calculated for each transition from one nephropathy state to the next





for the initial, early, middle and late phases of the model. Residual proportions remaining in state after transitions and deaths occurring in one year are calculated in the model.

From the ESRD state there is no transition but to death. The nephropathy mortality rates are age and gender specific and are based on estimates of the relative risk of death for patients in the first three states compared with non-diabetics. Although originally set to figures used by Eastman, these have been modified as part of the model calibration process to match published outcomes. The mortality rates for patients in the ESRD are extracted from Eastman's paper^{39,45} and were derived from a dialysis register in the USA. The proportions are calculated by gender and each progression phase.

A compound residue function is used to derive 5year state transition probabilities for every possible pair of start and end states. These are derived by assigning a probability to each possible path from the start state to the end state and summing these probabilities across all pairs. Each year is an independent event; therefore, the probability of a particular 5-year path is the product of annual probabilities of remaining in the same state, progression to the next state or dying. This process is repeated for the early, middle and late phases, and the rates are referred to as the 'regular' transition rates.

As mentioned earlier, separate transition rates are derived for the initial phase between diagnosis and the first 5-year age maker, using the early phase annual transition probabilities. Since a patient can be diagnosed with type 2 diabetes in any of the 5 preceding years, the initial phase lasts for 2.5 years on average. For each combination of start state and end state the probability of this pair is calculated over 1, 2, 3, 4 and 5 years. The average of these five probabilities is the 'initial' transition matrix.

The method used in the model to estimate the transition probabilities is correct.

The 5-year nephropathy transition rates are used to calculate the number of patients in each state by age for each of the 30 age/gender diagnosis groups.

The newly diagnosed patients are distributed across the four nephropathic states. The model assumes the same distribution for all 30 diagnosis groups. In particular, of the newly diagnosed diabetics:

- 87.1% have no nephropathy
- 11% have microalbuminuria
- 1.9% have GPR.

The prevalence of microalbuminuria used in the model is identical to the one used by Eastman.^{39,45}

The number of patients in each disease state at each age is determined by applying the 5-year state transition rates (analysed in the previous section) to the number of patients in the starting state in the previous period and adding up all of the possible combinations. The initial rates are applied to the first transition (from outset to the first 5-year age mark). The early rates are applied to the next two transitions (next 10 years), the middle rates are applied to the following two transitions (next 10 years) and the late rates apply thereafter until death or the age of 100. In addition, the number of patients in each state is adjusted so that the total number of patients alive (dead) matches the number predicted by the mortality/CVD module. This is done by adjusting pro rata the number of patients in each state against the discrepancy in the overall number of patients alive.

The number of dead by state is calculated by applying the probability of death for that state to the number of patients in the state. These figures are again reconciled to match the values predicted in the mortality/CVD module.

The accumulated alive and dead by state at death are derived directly from the two previous calculations.

The total number of patients in each state at 5-year age points is then derived, by summing up all of the initiation groups.

Moreover, the number of new cases per annum of GPR by transition phase and gender is estimated from prevalent cases and transition rates. The model provides the number of patients alive by nephropathy state and age group for males, females and total. These figures are also discounted at the cost and then at the outcome rate.

As far as the prevalence of nephropathy is concerned, the model provides estimations for:

- prevalence by age
- prevalence by time from diagnosis
- prevalent caseload by time from diagnosis
- cumulative incidence by time from diagnosis.

Finally, the model calculates the number of patients in each state (no microalbuminuria, microalbuminuria, GPR, ESRD) by subcohort and time from diagnosis.

Retinopathy

Eastman's retinopathy module is redesigned to represent combinations of possible conditions and severities. Eighteen compound transition rates were calculated from five elemental transition probabilities. The structure comprises nine distinct morbid states: no retinopathy, background diabetic retinopathy (BDR), MO, MO with blindness in one eye (B1), proliferative diabetic retinopathy (PDR), PDR with blindness in one eye, MO and PDR, MO and PDR with blindness in one eye and blindness in both eyes (B2). The model uses linear functions with a WESDR threshold in the normal range of HbA_{1c} to show the influence of glycaemic levels on transition rates. (Alternatively, the user can use Eastman's power functions.) The duration of diabetes is divided into four phases: early, middle, late and continuing. The model produces a set of 5-year transition probabilities across all nine states of retinopathy by diagnosis group for each one of the phases.

The retinopathy model assumes nine states of disease:

- no retinopathy
- BDR
- MO
- MO and B1
- PDR

- PDR and B1
- MO and PDR
- MO, PDR and B1
- B2.

These states are mutually exclusive and in any year, the patient can either remain in the same state, progress to another state or die. *Figure 6* illustrates the retinopathy module.

Published annual rates of transition to retinopathic states are used to derive 5-year state transition probabilities.

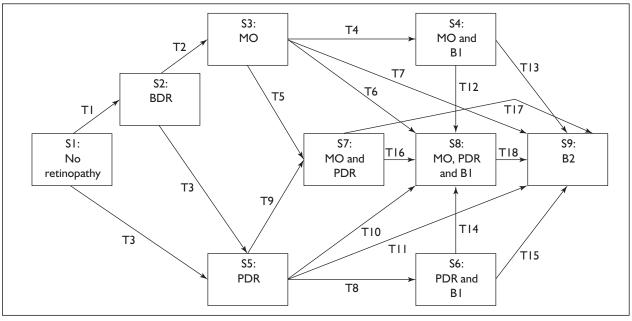
Two alternative sets of input transition rates can be used:

- transition rates used in Eastman's^{39,45} method
- transition rates by fitting the same data to a linear function with thresholds.

The transition probabilities used in the model increase as the time since diagnosis of diabetes increases. The rates are calculated separately for each of the 30 age/gender diagnosis groups.

The transition probabilities are the net transition rates for a cohort of cases. While an individual patient might move from one state to any other within a year, at the level of net transitions a cohort can only progress to at most the next most severe state within a year.

The duration of diabetes is divided into four phases:



- early phase: this phase begins at diagnosis and ends at the first 5-year age marker. Since diagnoses are made randomly in time, this phase lasts for 2.5 years on average
- middle phase: this phase follows the early phase and lasts for 10 years
- late phase: this phase follows the middle phase and lasts for 10 years
- continuing phase: this phase follows the late phase and lasts until death or the age of 100 years.

It is assumed that transition rates for retinopathy and MO are statistically independent.

The model produces a set of 5-year transition probabilities across all nine states of retinopathy by diagnosis group. In addition, a modified set of transition rates is derived for the initial period from diagnosis of diabetes to the next model marker age. In this module, the early phase coincides with the initial period, so no 5-year transition rates are calculated for the early phase.

In total, 18 distinct transition probabilities are used, derived from five underlying probabilities.

Annual risks of developing MO and PDR blindness are provided with and without photocoagulation treatment. Moreover, the standard risks of developing BDR, MO and PDR are calculated for the early, middle, late and final stages. The original Eastman hazard rates were amended to calibrate outcomes against WESDR results.

The risks of developing BDR, ME, PDR and blindness that are used to calculate the transition probabilities are provided in the model. The elemental hazard rates used for the estimation of the transitions to non-proliferative retinopathy, proliferative retinopathy, MO and blindness are identical to those used by Eastman.^{39,45}

Multipliers then calibrate these rates, in order to achieve an approximation to published WESDR incidence and prevalence rates.

The values of these basic parameters depend on the choice of glycaemia gradient (Eastman or WESDR) and are adjusted for current mean HbA_{1c} values and ethnic mix variations.

The Eastman power function method can be used. In this method, the rates are derived by Eastman by fitting the WESDR study results to an exponential curve and applying it to the risk of developing neuropathy. The current HbA_{1c} is standardised by dividing it by the standard HbA_{1c} of 10.0 from DCCT data and raising it to a power parameter to fit the observations. Moreover, there is an option to adjust transition rates to neuropathy and PVD for ethnic mix variations.

Alternatively, the WESDR linear threshold can be used. In this method, the relationship between HbA_{1c} and the annual risk of developing neuropathy is assumed to be linear, but with a threshold value (HbA_{1c} of 5.35 for 'normal to BDR' and 7.0 for 'BDR to MO' and 'BDR to PDR'), below which the risk is zero. The basic annual transition rates are assumed to apply when HbA_{1c} = 10.0 and to increase linearly in proportion to the excess of HbA_{1c} above the threshold value (6.0).

In this model, the second method of the WESDR linear threshold is used for the central estimate.

Transition probabilities are then calculated for the early, middle, late and final phases of the model, using the basic model parameters mentioned above, the WESDR linear threshold method and assuming independent probabilities.

The methods applied to calculate the compound transition probabilities are correct.

Residual proportions of patients in each disease state are estimated after transitions to more severe states and deaths. A compound residue function is used to derive 5-year state transition probabilities for every possible pair of start and end states. These are derived by assigning a probability to each possible path from the start state to the end state and summing these probabilities across all pairs. The transition probabilities are estimated for each of the middle, late and final phases of the disease.

Finally, separate transition rates are derived for the initial phase between diagnosis and the first 5-year age marker using the early phase annual transition probabilities. Since a patient can be diagnosed with type 2 diabetes in any of the 5 preceding years, the initial phase lasts for 2.5 years on average. For each combination of start state and end state the probability of this pair is calculated over 1, 2, 3, 4, and 5 years. The average of these five probabilities is the 'initial' transition matrix.

The 5-year neuropathy transition rates are used to calculate the numbers of patients in each state by age for each of the 30 age/gender diagnosis groups.

The newly diagnosed patients are distributed across the nine retinopathic states. The model

assumes the same distribution for all 30 diagnosis groups. In particular, of the newly diagnosed diabetics:

- 80% have no background diabetic retinopathy
- 20% have background diabetic retinopathy.

The prevalence values used in the model are identical to those used by Eastman.^{39,45}

The number of patients in each disease state at each age is determined by applying the 5-year state transition rates (analysed in the previous section) to the number of patients in the starting state in the previous period and adding up all the possible combinations. The initial rates are applied to the first transition (from outset to the first 5-year age mark). The regular rates apply thereafter until death or the age of 100. In addition, the number of patients in each state is adjusted, so that the total number of patients alive (dead) matches the number predicted by the mortality/CVD module. This is done by adjusting pro rata the number of patients in each state against the discrepancy in the overall number of patients alive.

The number of dead by state is calculated by applying the probability of death for that state to the number of patients in the state. These figures are again reconciled to match the values predicted in the mortality/CVD module.

The accumulated alive and dead by state at death derive directly from the two previous calculations.

The total number of patients in each state at 5year age points is estimated, by summing up all of the initiation groups.

The model provides the number of patients alive by retinopathy state and age group for males, females and total. These figures are also discounted at the cost rate.

As far as the prevalence of retinopathy is concerned, the model provides estimations for the states: background retinopathy, MO, MO and PDR, blind in one eye and blind in both eyes. In particular, the following are calculated:

- prevalent caseload: estimated number of patients in some states by age band and transition phase
- prevalence by age (in 5-year age points)
- prevalence by time from diagnosis
- prevalent caseload by time from diagnosis
- cumulative incidence by time from diagnosis.

Finally, the model calculates the number of patients in each state (no retinopathy, BDR, MO, PDR, MO and B1, PDR and B1, MO and PDR, MO/PDR and B1, B2) by subcohort and time from diagnosis.

Hypoglycaemia

The model estimates the number of hypoglycaemic events per annum in the cohort by age and gender. However, these estimates are not linked to the costing module, as the latter includes the cost of hypoglycaemia implicitly within the excess cost of diabetes.

It is assumed that DCCT relationships between mean HbA_{1c} and incidence of hypoglycaemia apply (as used by Eastman).

The number of severe hypoglycaemic events (requiring medical attention) expected per year by age group and gender is calculated by multiplying period event rates by the average number of patients. The period event rates are weighted average event rates applicable to the 2.5/5-year periods used in the model and are based on the annual event rates. The latter are calculated as the average of the treatment-specific event rates, weighted by the proportions of patients receiving each mode of treatment. The treatment-specific event rates are calculated using DCCT relationships as used by Eastman.

Weight gain

The weight gain is included in the model indirectly. In particular, BMI is one of the variables used to calculate blood pressure. The latter is then used in the model to estimate the CVD death rates. It was not feasible to interfere in the model and single out the effect of change in the BMI only.

Cataract and cancer

The model estimates the prevalence of cancer and cataract in the diabetic cohort, based on analysis of data collected from South Glamorgan residents.²⁵ Cataract prevalence is estimated by fitting a simple linear regression model to the actual rates of cataract. Cancer prevalence is calculated by fitting an exponential model for males and a simple linear regression model for females to the actual rates of cancer.

The model provides estimates of cataract and cancer in the diabetic cohort by age and gender. Relationships are based on analysis of South Glamorgan data, which are assumed to be representative.

The number of cataract operations per year by age and gender is calculated among patients with type 2 diabetes. In particular, the number of patients alive in each age group is multiplied by the excess of the average HbA_{1c} above a threshold of 5% and an assumed gradient of 0.2% per 1% increase in HbA_{1c}. These estimates are descriptive only and are not linked to other parts of the model.

Simple linear regression models were fitted to the actual rates of cataract prevalence from the South Glamorgan data, to obtain model equations and estimate prevalence by age and gender. The coefficient of determination (R^2) in the above two models was very high (75% and 95%, respectively), indicating that the models have sufficiently fitted the South Glamorgan data.

An exponential model for males and a simple linear regression model for females were fitted to the actual rates of cancer prevalence from the South Glamorgan data, to obtain model equations and estimate prevalence by age. The coefficient of determination (R^2) in the above two models was very high (86% and 73%, respectively), indicating that the models have sufficiently fitted the South Glamorgan data.

Multimorbidity

The model estimates the number of patients suffering from more than one of the five major complication states of diabetes (CHD, stroke, blindness in one or both eyes, chronic DFU and/or LEA and ESRD). Frequencies of multiple complications are derived from analysis of the Cardiff database of South Glamorgan residents. These frequencies and univariate prevalence values of the above complication states are used to calculate the prevalence of every possible combination of the conditions.

The model estimates the proportions of male and female patients with type 2 diabetes who suffer from various combinations of the five most severe complication states:

- CHD
- stroke
- blindness in one or both eyes
- chronic DFU and/or LEA
- ESRD.

The frequencies of multiple complications are assumed to be those derived from analysis of the Cardiff database of South Glamorgan residents.

The univariate prevalence by age of the above complication states is used to calculate the crude prevalence of every combination of the conditions. Each possible combination is represented by a set of 0s/1s. These are combined with the five separate univariate prevalences, to yield the probability of patients having the specified combination of complications, assuming statistical independence. Finally, the resulting figure is multiplied by the relative probability factor derived from analysis of the Cardiff database, to represent the relative frequencies of different combinations. This prevalence is then adjusted, by dividing the prevalence of every combination of the condition by the total per age group, to standardise the sum of probabilities in each age range to unity.

The probabilities for the number of morbidities present by age and gender are then calculated, by adding together the appropriate adjusted prevalences for the 0, 1, 2, 3, 4 and 5 serious morbid conditions.

Finally, the total number of patients with 0, 1, 2, 3, 4 and 5 morbidities present is estimated by multiplying the number of patients alive in each age group by the probabilities for the number of morbidities present.

The method used to calculate the prevalence and number of morbidities is appropriate.

Cost aspects

Costs are in 2000 UK pounds.

The main costing areas included within the models are:

- inpatient costs
- outpatient costs
- primary care costs
- drug costs.

Inpatient costs

Hospital admissions are assigned by primary diagnosis to one of nine broad categories (diabetes, nephropathy, retinopathy, neuropathy and skin infections, heart disease, stroke, cancer, cataract and other diseases). A polynomial or exponential model is used to calculate the annual admission rate and average length of stay by age. Hospital costs are calculated from a fixed cost per admission plus a variable cost per bed-day. The cost parameters are derived from national NHS speciality costs for the dominant specialities for each costing category.

Age/gender models for admission rates and length of stay are derived from original analysis of the Cardiff database of South Glamorgan residents.²⁵

Inpatient costs consist of a fixed treatment cost per admission plus a variable cost dependent on the length of hospital stay. In general, day-case admissions are assumed to incur the fixed cost and a nominal half-day length of stay. Admission rates vary by age, gender and diabetic status, whereas the unit costs are common to all categories of patient.

The number of male and female admissions per annum per 1000 is calculated by age and for each disease state, using polynomial and exponential model coefficients for 4 years. Thus, the total number of admissions can be then calculated as the sum of the individual admissions.

The male and female average length of stay is calculated by age and for each disease state, using polynomial model coefficients. Thus, the total bed-days and the overall length of stay are then estimated. The former is the length of stay multiplied by the number of admissions in each case and the latter is the total bed-days over the total admissions in each case.

Inpatient cost parameters are provided by disease state for an admission (fixed treatment cost) and for a bed-day (variable cost). These are presented in *Table 35*.

Inpatient unit costs are derived from CIPFA English Trust Financial Returns 1995/96.²⁵

Total inpatient costs over all years of life are calculated by age, gender, disease state and diabetic status. The cost data in Table 35, the number of patients alive and the number of patients in individual disease states are used for the calculations. Moreover, the number of admissions for microvascular states is estimated by multiples of non-diabetic rates, which reflect the overall diabetic admission rates. The annual inpatient cost per patient alive is also provided by age, gender and diabetic status. The mentioned costs are then discounted at cost rate.

Finally, the discounted inpatient costs by time from diagnosis are estimated by gender, diabetic status and time from diagnosis.

Outpatient treatment costs

Attendances and outpatient costs are estimated for patients identified as suffering from one of eight disease conditions (CHD, stroke, neuropathy, nephropathy, retinopathy, cataract, cancer and other diseases). A polynomial or exponential model is used to calculate the number of attendances for each of nine UK outpatient specialities. The cost parameters are derived from national NHS speciality-specific clinic costs.²⁵

The UK outpatient specialities are represented in nine broad groups, which are assumed to be homogeneous:

- general medicine; geriatrics; haematology
- ophthalmology
- obstetrics and gynaecology
- cardiology; cardiothoracic surgery
- general surgery; trauma and orthopaedic; dermatology; rehabilitation
- radiotherapy; anaesthetics
- mental handicap and mental illness
- neurology; neurosurgery
- ENT; urology; rheumatology.

The model outpatient activity is calculated through eight disease drivers:

- CHD
- stroke
- neuropathy
- nephropathy
- retinopathy
- cataract

TABLE 35 Inpatient costs by morbid state per admission and length of stay

Disease/complication	Cost per admission	ion Cost per bed-day	
Diabetes	£420	£136	
Nephropathy	£696	£129	
Retinopathy	£680	£180	
Neuropathy/skin infection	£180	£204	
Heart disease	£904	£156	
Stroke	£518	£147	
Cancer	£267	£204	
Cataract	£260	£200	
Other diseases	£300	£140	

- cancer
- other.

Overall Cardiff outpatient data²⁵ were decomposed into incremental additive models to resource use within each speciality group, with a residue associated only with patient numbers. Since driver models are incremental, they are not constrained to non-negative values. Therefore, results should only be interpreted in aggregate form.

Age/gender models for attendance rates are derived from original analysis of a Cardiff database of South Glamorgan residents.

Polynomial and other model coefficients of attendances for 4 years are used to estimate the number of male and female attendances per annum per 1000 by age, speciality group and condition.

Outpatient unit costs are derived from CIPFA English Trust Financial Returns 1995/96.²⁵

Costs per outpatient attendance and per speciality are provided in the model. These costs are used to calculate the average outpatient costs for males and females. Relative activity weights are assigned to individual specialities in each group based on activity volumes in the Cardiff database. The average costs per speciality are presented in *Table 36*.

The model provides estimations for the total cost of outpatient care by age group, gender, diabetic status and major complication/disease. These are based on attendance rates, unit costs and patient numbers. Attendance rate weights for microvascular states are used to exclude 'no complication' states from the calculation of attendance numbers. Moreover, the excess cost due to diabetes is calculated, as well as the overall attendance rates for diabetic and non-diabetic patients per 1000 per year. These costs are then discounted at the cost rate.

Finally, the model estimates the undiscounted costs of outpatient care by gender, diabetic status, major complication/disease and time from diagnosis.

Primary care services costs

The proportion of GP consultations attributable to type 2 diabetes is assumed to be the same as the proportion of prescriptions dispensed to type 2 diabetic patients. The number of GP consultations is estimated using age-related curves (exponential model) from the General Practice Morbidity Database Project (GPMDP). The average cost per GP consultation is assumed to be £12 [Personal Social Services Research Unit (PSSRU)].²⁵

Coefficients from polynomial models are used to calculate the number of GP consultations per annum per person by age, gender and diabetic status. The latter are then used to estimate the total over all years of life of GP consultations by age, gender and diabetic status. Finally, multiplying the total consultations by the average cost of a GP consultation gives the total cost of GP consultations by age and gender. This is then discounted at the cost rate.

Moreover, the total cost of GP consultation by gender and time from diagnosis is calculated, by multiplying the consultation rates by the number of patients alive by time from diagnosis. This is then discounted at the cost rate.

Unit costs are derived from PSSRU 'Health and social services costs'. 25

Community health costs

The proportion of community health contacts attributable to type 2 diabetes is assumed to be the same as the proportion of prescriptions dispensed to type 2 diabetic patients. The number of

Speciality grouping	Male diabetics	Female diabetics
General medicine; geriatrics; haematology	£84.59	£85.42
Ophthalmology	£49.00	£49.00
Obstetrics and gynaecology	-	£72.13
Cardiology; cardiothoracic surgery	£75.35	£75.32
General surgery; trauma and orthopaedic; dermatology; rehabilitation	£66.44	£65.95
Radiotherapy; anaesthetics	£90.52	£92.65
Mental handicap and mental illness	£104.87	£106.31
Neurology; neurosurgery	£117.78	£117.40
ENT; urology; rheumatology	£68.05	£68.61

TABLE 36 Outpatient costs per speciality

community health contacts is estimated using agerelated curves (exponential model) from GPMDP. The contacts are calculated per health professional (chiropodist, dietitian, practice nurse and other) and costed with data from PSSRU.²⁵

The proportion of community health contacts attributable to type 2 diabetes is assumed to be the same as the proportion of prescriptions dispensed to type 2 diabetic patients.

Coefficients from polynomial models and the proportion of diabetic prescriptions for type 2 diabetes are used to calculate the community health contacts per annum per person by age, gender and diabetic status. In particular, number of contacts is calculated for chiropodists, dietitians, practice nurses and others.

Unit costs are derived from PSSRU 'Health and social services costs'. 25

The average contacts per referral and the average cost per contact used in the model are presented in *Table 37*.

The community health contacts per annum by age, gender, diabetic status and type of contact are calculated by multiplying the number of contacts by the number of patients alive. These are then multiplied by the respective costs to give the community health costs per annum by age, gender, diabetic status and by type of contact. The contacts and the costs are also discounted at the cost rate.

The community health costs per annum by gender, diabetic and time from diagnosis are also provided. They are calculated by multiplying the number of contacts by the number of patients alive and contact rates and costs.

Drug costs

A logistic model from a Tayside prescriptions paper⁵⁸ is used to calculate the proportion of diabetic prescriptions attributable to type 2

Professional	Average contacts per referral	Average cost per contact
Chiropodist	6	£11.50
Dietitian	3	£9.00
Practice nurse	2	£6.00
Other	4	£10.00

diabetes. These are then used in polynomial agerelated models, to estimate the number of prescriptions. The average cost of prescriptions (excluding drugs for diabetes) is derived from the Tayside paper and is assumed to be £8.90 for diabetics.

Average prescriptions per annum by age and gender are calculated, using polynomial model coefficients multiplied by the proportion of diabetic prescriptions attributable to type 2 diabetes. The proportion of diabetic prescriptions is calculated using logistic model coefficients. The model is based on data from the Tayside prescriptions paper. Total prescriptions per annum by age and gender are then provided, by multiplying the average prescriptions by the number of patients alive. These are then discounted at cost rate.

The average cost of prescriptions (excluding drugs for diabetes) is £8.8961 for diabetics and £8.7311 for non-diabetics. The source used is the Tayside diabetes paper (type 2 only).

The total cost of prescriptions is then calculated by multiplying the total prescriptions by their average cost. The cost is also adjusted for inflation changes. The total cost of prescriptions is also discounted at cost rate.

The total cost of prescriptions by time from diagnosis is calculated by multiplying the prescribing rates by the number of patients alive by time from diagnosis and the average costs.

The cost of diabetic therapy by age, gender and type of therapy is estimated by multiplying the average annual medication cost by the number of patients treated with this medication. This cost is then discounted. Moreover, the same cost is provided by time from diagnosis.

Costs

The inpatient, outpatient, primary and drug costs estimated in the previous sections are summarised and the following costs are additionally calculated:

- total costs by age and gender (undiscounted and discounted)
- total costs by time of diagnosis and by gender (undiscounted and discounted)
- 1-year total costs per patient alive
- lifetime healthcare costs from diagnosis by age and gender
- aggregate total costs for each type of cost by gender (undiscounted and discounted)

- total cost per survival year
- total cost per person.

Clinical outcomes

The model provides the following outputs:

- average age at diagnosis and death by sex
- average survival by gender
- average life expectancy by age at diagnosis and gender
- patients alive by time from diagnosis and gender (number and proportion)
- years of life (discounted at outcomes rate and undiscounted) by gender in the following states:
 nephropathic

 - neuropathic
 - retinopathicCHD
 - chD – stroke
- proportion of patients in state of each complication by gender.

The GSK model²⁵ estimates mean life expectancy, total lifetime costs, costs per LYG and incremental cost per QALY.

Utility scores

QALYs visual analogue scale

Utility values recently available from the CODE-2 study are used and a multivariate model of EuroQol visual analogue scale (VAS) scores is generated.²⁵ The latter is used to calculate the mean QALY scores in the model.

A multivariate model of EuroQol VAS scores derived from CODE-2 data together with cohort details are used to calculate the mean QALY score for patients in each diagnosis subcohort by age and gender, and for each model period. Moreover, mean scores applying across each 2.5/5-year model period from diagnosis for each subcohort by age are estimated by averaging figures from VAS scores calculated earlier.

The CODE-2 EuroQol responses are representative of a general population of people with Type 2 diabetes and offer the patient perspective.

The model used to calculate the QALY scores consists of an exponential equation. The percentage of patients in each morbid state is multiplied by the model coefficients and they are all summed together with the constant of the model. The model is additionally adjusted for age characteristics. The exponential of this sum is then calculated (plus some adjustments) to provide the QALY score of the model.

TABLE 38 EuroQol VAS model coefficients

EuroQol VAS	Model coefficients
Constant	4.40500
ESRD	-0.45700
Insulin	-0.19700
Stroke	-0.19200
Amputation/DFU	-0.1310
Retinopathy	-0.12060
PVD only	-0.10100
Proteinuria	-0.09630
Female	-0.09080
Tablets	-0.06060
CHD	-0.03250
Age	-0.00221

The coefficients used in the model are presented in *Table 38*.

QALYs EuroQol 5 Dimensions (EQ-5D)

Utility values recently available from the CODE-2 study are used and a multivariate model of EQ-5D scores is generated.²⁵ The latter is used to calculate the mean QALY scores in the model.

A multivariate model of EQ-5D scores derived from CODE-2 data together with cohort details are used to calculate the mean QALY score for patients in each diagnosis subcohort by age and gender and for each model period. Moreover, mean scores applying across each 2.5/5-year model period from diagnosis for each subcohort by age are estimated by averaging figures from VAS scores calculated earlier.

The CODE-2 EuroQol responses are representative of a general population of people with Type 2 diabetes and offer the societal perspective.

The model used to calculate the QALY scores consists of an exponential equation. The percentage of patients in each morbid state is multiplied by the model coefficients and they are all summed together with the constant of the model. The model is additionally adjusted for age characteristics. The exponential of this sum is then calculated (plus some adjustments) to provide the QALY score of the model.

The coefficients used in the model are presented in *Table 39*.

The following QALY calculations were performed in the model:

TABLE 39 EuroQol EQ-5D model coefficients

EuroQol EQ5-D	Model coefficients
Constant	-0.002133
Nephropathy	-0.130000
Insulin	-0.074230
Stroke	-0.165000
Amputation	-0.207000
DFU	-0.182000
Retinopathy	-0.027160
Blind	-0.153000
PVD only	-0.088640
Tablets	-0.021130
CHD	-0.013150
Age	-0.001758

- QALYs from time of diagnosis to death (discounted at outcomes rate and undiscounted) by sex: two different utility measures were used, derived from the following sources:
 - CODE-2
 - EuroQol data
- life-years by time from diagnosis and gender: these are split for each complication state.
- QALYs VAS by time from diagnosis: these are aggregated QALYs using the VAS (discounted at outcomes rate and undiscounted) for gender and time from diagnosis.
- QALYs EQ-5D by time from diagnosis: these are aggregated QALYs using the EQ-5D scale (discounted at outcomes rate and undiscounted) for gender and time from diagnosis.

Validation of the GSK model

Assumptions made within the GSK model

In the absence of available type 2 data, type 1 diabetes data have been used in their place.

Again, the submission states that a weakness exists in that, for the AMI submodel, calculations were based on a predominantly white population and, therefore, the population used may not be representative of the diabetic population.

Conclusions on the critical appraisal of the GSK model 25

The model proposed by GSK contains a detailed representation of diabetic health states. The model presented by GSK is almost identical in structure to the model proposed by Bagust and colleagues.⁴³ The only significant difference is in the HbA_{1c} module. As explained earlier, the GSK

metabolic model estimates the glucose levels for each period after diagnosis, based on an impaired insulin sensitivity and a declining β -cell function. The levels of those two variables depend on the effect of the treatment applied. Patients move to the next therapy in the treatment pathway when their glucose level increases to a certain level. In contrast, the Bagust HbA_{1c} module is a distinct Markovian submodel, which incorporates annual transition rates to indicate the change from one treatment to another.

There are two main criticisms of the GSK submission. 25

- There was no transparency in the presentation of the model. A large amount of data was included in the model, for which no information or explanation was provided in the form of a report.
- The company performed no sensitivity analysis. Given the complexity of the model and the large number of parameters, this is a critical failing in establishing the credibility of the favourable economic outcomes. Specifically, the stability of the results and identification of key drivers would be of greatest value.

Key economic results for rosiglitazone

The economic model described by GSK²⁵ is used to investigate the use of rosiglitazone as an adjunctive therapy when monotherapy has failed to achieve an adequate level of control of average blood glucose levels. The two indications as detailed in the licensing documentation are examined:

- where metformin is contraindicated and where the use of sulfonylurea has failed to achieve adequate control (defined in the model analysis as $HbA_{1c} > 7.5\%$), the choices evaluated are to add rosiglitazone to sulfonylurea or to switch to insulin therapy
- in obese patients who have failed to achieve adequate glycaemic control with metformin the alternatives evaluated are to add rosiglitazone or sulfonylurea to metformin.

GSK has compared several scenarios of treatment pathways and the cost-effectiveness results are presented below. Note that all of the costs and QALYs included in the tables are cumulative per 1000 patients.

Lean patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression to insulin

$$\begin{array}{ccc} S & \longrightarrow & 1 \\ vs & & & lean \\ S & \longrightarrow & S+R & \longrightarrow & 1 \end{array}$$

This comparison reveals that rosiglitazone dominates in both cost and efficacy outcomes, as the cost of the scenario that adds rosiglitazone is lower than the cost of the scenario that progresses directly to insulin after sulfonylurea. Thus, in lean patients on sulfonylurea who are contraindicated to or intolerant to metformin, adding rosiglitazone after monotherapy failure is a costeffective strategy (*Table 40*).

The model was used to calculate the above figures and exactly the same results were found. In addition, the respective incremental costeffectiveness ratios (ICERs) were generated using LYG instead of QALYs (*Table 41*):

$$\begin{array}{ccc} S & \longrightarrow & 1 \\ v_S & & & lean \\ S & \longrightarrow & S+R & \longrightarrow & 1 \end{array}$$

The conclusion remains the same: adding rosiglitazone after monotherapy failure with sulfonylurea is preferable to progressing directly to insulin. Although the gain in life-years is small, the cost of rosiglitazone combination therapy is lower than the cost of insulin and this makes the second scenario dominant.

Finally, it was verified that the mechanism to calculate the QALYs and the LYG in the model works properly. In particular, all of the utility values were set equal to zero and exactly the same life-years results as those presented in the model were found.

Obese patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression to insulin



This comparison reveals that rosiglitazone dominates in both cost and efficacy outcomes, as the cost of the scenario that adds rosiglitazone is lower than the cost of the scenario that progresses directly to insulin after sulfonylurea. Thus, in obese patients on a sulfonylurea who are contraindicated to or intolerant to metformin, adding rosiglitazone after monotherapy failure is a cost-effective strategy (*Table 42*).

The model was used to calculate the above figures and exactly the same results were found.

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	-470,619	204.6	NA	-229,686	171.3	NA
22.5	12.5	-465,676	207.3	NA	-228,116	173.3	NA
27.5	17.5	-465,278	207.5	NA	-228,022	173.5	NA
32.5	22.5	-466,076	207.4	NA	-228,163	173.4	NA
	20.0	-465,677	207.5	NA	-228,093	173.5	NA

TABLE 40 Cost per QALY of adding R to S (lean)

TABLE 41 Cost per LYG of adding R to S (lean)

Time from diagnosis	Years from start of combination	Undiscounted			Discounted			
	therapy	Costs	LYG	ICER	Costs 6%	LYG 1.5%	ICER	
17.5	7.5	-470,619	1.7	NA	-229,686	1.4	NA	
22.5	12.5	-465,676	3.8	NA	-228,116	2.9	NA	
27.5	17.5	-465,278	4.1	NA	-228,022	3.1	NA	
32.5	22.5	-466,076	3.8	NA	-228,163	2.9	NA	
	20.0	-465,677	3.9	NA	-228,093	3.0	NA	

In addition, the respective ICERs were generated using LYG instead of QALYs (*Table 43*):

$$\begin{array}{cccc} S & & & & \\ v_{s} & & \\ S & & & \\ \end{array} \xrightarrow{} & S + R & & & 1 \end{array} \qquad obese$$

The conclusion remains the same: adding rosiglitazone after monotherapy failure with sulfonylurea is preferable to progressing directly to insulin. Although the gain in life-years is small, the cost of rosiglitazone combination therapy is lower than the cost of insulin and this makes the second scenario dominant.

Finally, it was verified that the mechanism to calculate the QALYs and the LYG in the model

works properly. In particular, all of the utility values were set equal to zero and exactly the same life-years results as those presented in the model were found.

Obese patients uncontrolled on maximum dose of metformin and not prescribed sulfonylurea: comparing addition of rosiglitazone with progression to insulin



This comparison reveals that rosiglitazone dominates in both cost and efficacy outcomes, as the cost of the scenario that adds rosiglitazone is

TABLE 42 Cost per QALY of adding R to S (obes

Time from diagnosis	Years from start of combination	Undiscounted			Discounted			
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER	
17.5	7.5	-439,172	204.3	NA	-196,501	171.1	NA	
22.5	12.5	-433,541	207.3	NA	-194,713	173.3	NA	
27.5	17.5	-431,161	208.2	NA	-194,148	173.9	NA	
32.5	22.5	-429,914	208.6	NA	-193,927	174.2	NA	
	20.0	-430,538	208.4	NA	-194,038	174.1	NA	

TABLE 43 Cost per LYG of adding R to S (obese)

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	LYG	ICER	Costs 6%	LYG 1.5%	ICER
17.5	7.5	-439,172	1.7	NA	-196,501	1.4	NA
22.5	12.5	-433,541	4.2	NA	-194,713	3.2	NA
27.5	17.5	-431,161	5.4	NA	-194,148	4.0	NA
32.5	22.5	-429,914	6.0	NA	-193,927	4.4	NA
	20.0	-430,538	5.7	NA	-194,038	4.2	NA

TABLE 44 Cost per QALY of adding R to M

Time from diagnosis	Years from start of combination	Ur	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER	
17.5	7.5	-658,971	292.0	NA	-317,472	241.5	NA	
22.5	12.5	-573,369	377.1	NA	-290,294	305.0	NA	
27.5	17.5	-519,917	398.1	NA	-277,613	319.5	NA	
32.5	22.5	-488,060	406. I	NA	-271,967	324.6	NA	
	20.0	-503,988	402. I	NA	-274,790	322.1	NA	

lower than the cost of the scenario that progresses directly to insulin after metformin. Thus, in obese patients on metformin whose physicians do not want to prescribe sulfonylureas, adding rosiglitazone after monotherapy failure is a cost-effective strategy (*Table 44*).

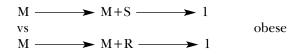
The model was used to calculate the above figures and exactly the same results were found.

In addition, the respective ICERs were generated using LYG instead of QALYs (*Table 45*):



The conclusion remains the same: adding rosiglitazone after monotherapy failure with metformin is preferable to progressing directly to insulin. Although the gain in life-years is very small, the cost of rosiglitazone combination therapy is lower than the cost of insulin and this makes the second scenario dominant. Finally, it was verified that the mechanism to calculate the QALYs and the LYG in the model works properly. In particular, all of the utility values were set equal to zero and exactly the same life-years results as those presented in the model were found.

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then progressing to insulin



This comparison reveals that rosiglitazone improves efficacy levels when added to metformin after monotherapy failure, compared with adding sulfonylureas. Thus, in obese patients on metformin who are able to take sulfonylureas, adding rosiglitazone after monotherapy failure is a cost-effective option, with an ICER of £5137/QALY at 20 years (*Table 46*).

Time from diagnosis	Years from start of combination	Un	Undiscounted			Discounted		
	therapy	Costs	LYG	ICER	Costs 6%	LYG 1.5%	ICER	
17.5	7.5	-658,971	3.4	NA	-317,472	2.8	NA	
22.5	12.5	-573,369	19.0	NA	-290,294	14.3	NA	
27.5	17.5	-519,917	40.8	NA	-277,613	29.4	NA	
32.5	22.5	-488,060	54.4	NA	-271,967	38.2	NA	
	20.0	-503,988	47.6	NA	-274,790	33.8	NA	

To achieve a realistic output representative of the UK for rosiglitazone licence and use, GSK has apportioned the final costs and effects at 85% for 4 mg and 15% for 8 mg.

TABLE 46	Cost per	QALY o	of M +	S versus	М	$^{+}$	R
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Time from diagnosis	Years from start of combination therapy	Undiscounted			Discounted			
		Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER	
17.5	7.5	1,449,421	89.1	16,274	713,459	71.5	9972	
22.5	12.5	1,532,021	170.8	8968	739,684	132.5	5582	
27.5	17.5	1,575,845	188.5	8361	750,081	144.7	5183	
32.5	22.5	1,597,716	193.7	8248	753,956	148.1	5091	
	20.0	1,586,781	191.1	8303	752,019	146.4	5137	

The model was used to calculate the above figures and exactly the same results were found.

In addition, the respective ICERs were generated using LYG instead of QALYs (*Table 47*):



Table 47 demonstrates that the scenario including rosiglitazone is not cost-effective. Adding rosiglitazone after monotherapy failure with metformin is not a cost-effective option, compared with adding sulfonylurea. The cost of rosiglitazone combination therapy is very large compared with the cost of sulfonylurea combination therapy. Moreover, the gain in life-years of using rosiglitazone therapy is very small. Thus, very high ICERs, which reach £430,095, are observed.

Finally, it was verified that the mechanism to calculate the QALYs and the LYG in the model works properly. In particular, all of the utility values were set equal to zero and exactly the same life-years results as those presented in the model were found. Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then switching to rosiglitazone

$$M \longrightarrow M+S \longrightarrow M+R \longrightarrow 1$$

obese
$$M \longrightarrow M+R \longrightarrow 1$$

This comparison reveals that rosiglitazone improves efficacy levels when added to metformin after monotherapy failure, compared with adding rosiglitazone after combination therapy of metformin plus sulfonylurea. Thus, in obese patients on metformin who are able to take sulfonylureas, adding rosiglitazone after monotherapy failure is a cost-effective option compared with introducing sulfonylurea and then switching to rosiglitazone. The ICER for this comparison is £6479/QALY at 20 years (*Table 48*).

The model was used to calculate the costeffectiveness for each of these scenarios and the same results were obtained.

In addition, the respective ICERs were generated, using LYG instead of QALYs (*Table 49*):

TABLE 47	Cost per LYG of $M + S$ versus $M + R$	

Time from diagnosis	Years from start of combination	Ur	ndiscounte	d	Discounted			
	therapy	Costs	LYG	ICER	Costs 6%	LYG 1.5%	ICER	
17.5	7.5	1,449,421	3.4	430,095	713,459	2.7	262,253	
22.5	12.5	1,532,021	16.9	90,773	739,684	12.8	57,85 I	
27.5	17.5	1,575,845	34.1	46,175	750,081	24.7	30,352	
32.5	22.5	1,597,716	43.3	36,894	753,956	30.6	24,627	
	20.0	1,586,781	38.7	40,984	752,019	27.7	27,184	

To achieve a realistic output representative of the UK for rosiglitazone licence and use, GSK has apportioned the final costs and effects at 85% for 4 mg and 15% for 8 mg.

Time from diagnosis	Years from start of combination	Undiscounted			Discounted			
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER	
17.5	7.5	1,482,499	76.4	19,417	727,504	61.4	11,857	
22.5	12.5	1,506,582	138.0	10,917	735,150	107.3	6850	
27.5	17.5	1,538,948	146.9	10,480	742,830	113.4	6549	
32.5	22.5	1,556,219	151.4	10,276	745,889	116.4	6411	
	20.0	I,547,584	149.1	10,376	744,359	114.9	6479	

$$M \longrightarrow M+S \longrightarrow M+R \longrightarrow 1$$

obese
$$M \longrightarrow M+R \longrightarrow 1$$

Table 49 indicates that the scenario including rosiglitazone is not cost-effective. Adding rosiglitazone after monotherapy failure with metformin is not a cost-effective option, compared with adding sulfonylurea and then switching to rosiglitazone. The cost of rosiglitazone combination therapy is very large compared with the cost of sulfonylurea combination therapy. Moreover, the gain in life-years of using rosiglitazone therapy is very small. Thus, very high ICERs, which reach £259,591, are observed.

Finally, it was verified that the mechanism to calculate the QALYs and the LYG in the model works properly. In particular, all of the utility values were set equal to zero and exactly the same life-years results as those presented in the model were found.

Sensitivity analysis on the HbA_{1c} threshold

A sensitivity analysis was performed on the HbA_{1c} threshold of switching therapies. In particular, the threshold was changed from HbA_{1c} = 7.5% to HbA_{1c} = 8.5%. The results of the analysis

TABLE 49 Cost per LYG of switching from M + S to M + R

concerning the cost-effectiveness of rosiglitazone are presented below.

Lean patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression to insulin



The conclusion remains the same: rosiglitazone dominates in both cost and QALYs in this scenario (*Table 50*).

Obese patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression to insulin



Adding rosiglitazone after sulfonylurea monotherapy failure dominates the scenario of progressing immediately to insulin. The change in threshold has not affected the final conclusion (*Table 51*).

Time from diagnosis	Years from start of combination	Ur	ndiscounte	counted		Discounted	
	therapy	Costs	LYG	ICER	Costs 6%	LYG 1.5%	ICER
17.5	7.5	1,482,499	3.5	428,221	727,504	2.8	259,591
22.5	12.5	1,506,582	14.5	104,013	735,150	11.0	66,771
27.5	17.5	1,538,948	27.0	56,994	742,830	19.7	37,770
32.5	22.5	1,556,219	34.5	45,067	745,889	24.5	30,437
	20.0	1,547,584	30.8	50,301	744,359	22.1	33,702

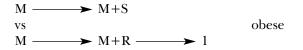
TABLE 50	Cost per Q	ALY of adding R to	o S (lean) using	8.5% threshold
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Time from diagnosis	Years from start of combination	U	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER	
27.5	7.5	-28,328	722.6	NA	-104,938	70.9	NA	
32.5	12.5	-263,244	425.5	NA	-103,497	72.5	NA	
37.5	17.5	-333,143	233.0	NA	-103,054	72.9	NA	
42.5	22.5	-340,507	136.2	NA	-102,941	73.1	NA	
	20.0	-336,825	184.6	NA	-102,998	73.0	NA	

Obese patients uncontrolled on maximum dose of metformin and not prescribed sulfonylurea: comparing addition of rosiglitazone with progression to insulin



Adding rosiglitazone after metformin monotherapy failure dominates the scenario of progressing immediately to insulin. The change in threshold has not affected the final conclusion (*Table 52*). Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then progressing to insulin



Adding rosiglitazone instead of sulfonylurea after metformin monotherapy failure remains a costeffective strategy even with a change in the switching threshold (*Table 53* versus *Table 46*). Moreover, the cost-effectiveness is now improved,

TABLE 51	Cost per QALY of adding R to S (obese) using 8.5% threshold
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Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
27.5	7.5	-282,840	95.5	NA	-82,116	5845.4	NA
32.5	12.5	-282,715	96.3	NA	-82,094	5070. I	NA
37.5	17.5	-281,342	96.6	NA	-81,912	4286.1	NA
42.5	22.5	-280,508	96.8	NA	-90,829	3614.1	NA
	20.0	-280,925	96.7	NA	-86,371	3950.1	NA

TABLE 52 Cost per QALY of adding R to M using 8.5% threshold

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
27.5	7.5	-228,952	43.	NA	-60,024	102.5	NA
32.5	12.5	-275,321	180.8	NA	-68,244	126.7	NA
37.5	17.5	-225,257	192.8	NA	-61,612	133.8	NA
42.5	22.5	-206,675	196.3	NA	-59,774	146.3	NA
	20.0	-215,966	194.6	NA	-60,693	140.0	

To achieve a realistic output representative of the UK for rosiglitazone licence and use, GSK has apportioned the final costs and effects at 85% for 4 mg and 15% for 8 mg.

TABLE 53 Cost per QALY of M + S versus M + R using 8.5% threshold

Time from diagnosis	Years from start of combination	U	Undiscounted		Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
27.5	7.5	498,865	46.9	10,636	129,107	32.8	3937
32.5	12.5	449,172	83.I	5407	120,297	56.0	2147
37.5	17.5	496,202	94.4	5258	126,527	62.8	2016
42.5	22.5	513,143	97.6	5258	128,205	75.1	1708
	20.0	504,672	96.0	5258	127,366	68.9	1848

as the incremental ratios are lower than those observed before the change in threshold.

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then switching to rosiglitazone

$$\begin{array}{ccc} M & \longrightarrow & M+S & \longrightarrow & M+R & \longrightarrow & 1 \\ vs & & & obese \\ M & \longrightarrow & M+R & \longrightarrow & 1 \end{array}$$

This comparison indicates that the marginal ICERs have increased after the increase in the glycaemia threshold (*Table 54* versus *Table 48*). However, adding rosiglitazone to metformin after monotherapy failure is still cost-effective compared with adding first sulfonylurea and then switching to insulin after the change in threshold; the discounted ICER is less than £30,000 (£24,962).

Sensitivity analysis on the treatment effects

Sensitivity analysis was performed on the treatment effects to investigate the impact of variations in the calibrators on the costeffectiveness results presented in the GSK submission. The main focus was on the scenarios that include rosiglitazone combination therapy. All of the changes made refer to the treatment effect of rosiglitazone. In particular, the effects of rosiglitazone treatment on insulin sensitivity and β -cell function were varied.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Exactly the same scenarios were examined as presented in the previous sections.

Lean patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression to insulin Change in insulin sensitivity



Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The results are presented Table 55.

TABLE 54 Cost per QALY of switching from M + S to M + R using 8.5% threshold

Time from diagnosis				Discounted			
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
27.5	7.5	539,161	7.8	69,092	139,213	5.6	24,962
32.5	12.5	539,200	14.5	37,189	139,220	15.9	8766
37.5	17.5	550,187	17.1	32,083	140,676	11.5	12,282
42.5	22.5	554,951	18.0	30,808	141,146	22.4	6291
	20.0	552,569	17.6	31,430	140,911	16.9	8316

TABLE 55 Cost per QALY of adding R to S (lean) following change in insulin sensitivity

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	-320,347	199.5	NA	-165,612	167.2	NA
22.5	12.5	-335,575	196.4	NA	-170,447	164.9	NA
27.5	17.5	-344,893	193.7	NA	-172,657	163.0	NA
32.5	22.5	-351,078	192.0	NA	-173,754	162.0	NA
	20.0	-347,986	192.9	NA	-173,206	162.5	NA

The difference in costs between the two scenarios has decreased compared with the results before the change in the effect (*Table 55* versus *Table 40*). However, the scenario of adding rosiglitazone still dominates in both cost and efficacy outcomes, as its cost is lower than the cost of progressing immediately to insulin after failure in monotherapy with sulfonylurea and the number of QALYs is greater.

Change in β -cell function

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The results are presented below.

$$S \longrightarrow 1$$

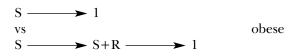
vs lean
$$S \longrightarrow S+R \longrightarrow 1$$

The difference in costs between the two scenarios has decreased compared with the results before the change in the effect (*Table 56* versus *Table 40*). However, the scenario of adding rosiglitazone still dominates in both cost and efficacy outcomes, as its cost is lower than the cost of progressing immediately to insulin after failure in monotherapy with sulfonylurea and the number of QALYs is greater.

Obese patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression to insulin Change in insulin sensitivity

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The results are presented below.



The difference in costs between the two scenarios has decreased compared with the results before the change in the effect (*Table 57* versus *Table 42*). However, the scenario of adding rosiglitazone still dominates in both cost and efficacy outcomes, as its cost is lower than the cost of progressing immediately to insulin after failure in monotherapy with sulfonylurea and the number of QALYs is greater.

Change in β -cell function

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

TABLE 56 Cost per QALY of adding R to S (lean) following change in β -cell function

Time from diagnosis	Years from start of combination therapy	Undiscounted			Discounted		
		Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	-317,731	200.5	NA	-164,526	168.0	N/A
22.5	12.5	-328,518	198.5	NA	-167,951	166.5	NA
27.5	17.5	-336,325	196.1	NA	-169,803	164.9	NA
32.5	22.5	-342,059	194.5	NA	-170,820	163.9	NA
	20.0	-339,192	195.3	NA	-170,312	164.4	NA

TABLE 57 Cost per QALY of adding R to S (obese) following change in insulin sensitivity

Time from diagnosis	Years from start of combination therapy	Undiscounted			Discounted		
		Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	-337,220	202.0	NA	-153,133	169.3	NA
22.5	12.5	-341,065	202.1	NA	-154,353	169.3	NA
27.5	17.5	-343,938	201.2	NA	-155,035	168.7	NA
32.5	22.5	-346,117	200.6	NA	-155,422	168.3	NA
	20.0	-345,028	200.9	NA	-155,229	168.5	NA

The results are presented below.

$$\begin{array}{ccc} S & \longrightarrow & 1 \\ vs & & \\ S & \longrightarrow & S+R & \longrightarrow & 1 \end{array} \quad obese$$

The difference in costs between the two scenarios has decreased compared with the results before the change in the effect (*Table 58* versus *Table 42*). However, the scenario of adding rosiglitazone still dominates in both cost and efficacy outcomes, as its cost is lower than the cost of progressing immediately to insulin after failure in monotherapy with sulfonylurea and the number of QALYs is greater.

Obese patients uncontrolled on maximum dose of metformin and not prescribed sulfonylurea: comparing addition of rosiglitazone with progression to insulin

Change in insulin sensitivity

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The results are presented below.



The results indicate that the difference in costs between the two scenarios has considerably decreased compared with the results before the sensitivity analysis (*Table 59* versus *Table 44*). However, the scenario of adding rosiglitazone still dominates in both cost and efficacy outcomes, as its cost is lower than the cost of progressing immediately to insulin after failure in monotherapy with metformin and the number of QALYs is greater.

Change in β -cell function

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The results are presented below.



Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	-336,266	202.4	NA	-152,733	169.6	NA
22.5	12.5	-338,532	202.9	NA	-153,452	170.0	NA
27.5	17.5	-340,905	202.2	NA	-154,015	169.5	NA
32.5	22.5	-342,938	201.7	NA	-154,376	169.1	NA
	20.0	-341,922	202.0	NA	-154,196	169.3	NA

TABLE 58 Cost per QALY of adding R to S (obese) following change in β -cell function

TABLE 59 Cost per QALY of adding R to M following change in insulin sensitivity

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	-305,465	195.1	NA	-166,509	163.6	NA
22.5	12.5	-327,856	190.3	NA	-173,618	160.0	NA
27.5	17.5	-335,726	187.9	NA	-175,486	158.3	NA
32.5	22.5	-338,35 I	187.1	NA	-175,951	157.8	NA
	20.0	-337,039	187.5	NA	-175,718	158.1	NA

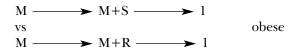
The results indicate that the difference in costs between the two scenarios has decreased considerably compared with the results before the sensitivity analysis (*Table 60* versus *Table 44*). However, the scenario of adding rosiglitazone still dominates in both cost and efficacy outcomes, as its cost is lower than the cost of progressing immediately to insulin after failure in monotherapy with metformin and the number of QALYs is greater.

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then progressing to insulin Change in insulin sensitivity

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The results are presented below.

62



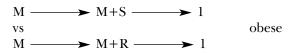
The scenario of adding rosiglitazone instead of sulfonylureas after monotherapy failure with metformin is no longer cost-effective after the change in the treatment effect (*Table 61* versus *Table 46*). In particular, it is dominated by the scenario of adding sulfonylureas, as the QALYs in this scenario are greater than in the one with rosiglitazone.

Hence, it may be concluded that the investigation of this comparison of scenarios proves to be very sensitive to changes in the treatment effects.

Change in β -cell function

The results are presented below.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.



The scenario of adding rosiglitazone instead of sulfonylureas after monotherapy failure with

TABLE 60	Cost per	QALY	of adding	R to M	following	change in	β-cell function
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Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER NA NA NA
17.5	7.5	-357,983	196.7	NA	-188,923	164.9	NA
22.5	12.5	-374,933	193.4	NA	-194,304	162.4	NA
27.5	17.5	-380,128	191.8	NA	-195,537	161.3	NA
32.5	22.5	-381,276	191.4	NA	-195,740	161.0	NA
	20.0	-380,702	191.6	NA	-195,638	161.2	NA

To achieve a realistic output representative of the UK for rosiglitazone licence and use, GSK has apportioned the final costs and effects at 85% for 4 mg and 15% for 8 mg.

TABLE 61 Cost per QALY of M + S versus M + R following change in insulin sensitivity

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	1,802,927	-7.8	NA	864,422	-6.4	NA
22.5	12.5	1,777,534	-16.0	NA	856,360	-12.5	NA
27.5	17.5	1,760,036	-21.7	NA	852,208	-16.4	NA
32.5	22.5	1,747,425	-25.3	NA	849,973	-18.8	NA
	20.0	1,753,730	-23.5	NA	851,090	-17.6	NA

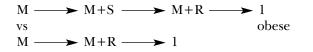
metformin is no longer cost-effective after the change in the treatment effect (*Table 62* versus *Table 46*). In particular, it is dominated by the scenario of adding sulfonylurea, as the QALYs in this scenario are greater than in the one with rosiglitazone.

Hence, it may be concluded that the investigation of this comparison of scenarios proves to be very sensitive to changes in the treatment effects.

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then switching to rosiglitazone Change in insulin sensitivity

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The results are presented below.



The results indicate that the scenario of adding rosiglitazone directly after monotherapy failure with metformin is no longer cost-effective after the sensitivity analysis is performed (*Table 63* versus *Table 48*). In particular, it is dominated by the scenario of adding first sulfonylurea and then switching to rosiglitazone after failure of combination therapy with sulfonylurea, as the QALYs in this scenario are greater than in the one with direct adding of rosiglitazone.

Hence, it may be concluded that the investigation of this comparison of scenarios proves to be very sensitive to changes in the treatment effects.

Change in β -cell function

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The results are presented below.

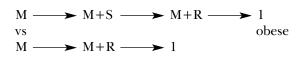


TABLE 62 Cost per QALY of M + S versus M + R following change in β -cell function

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER NA NA NA
17.5	7.5	1,750,410	-6.2	NA	842,008	-5.I	NA
22.5	12.5	1,730,457	-12.9	NA	835,674	-10.1	NA
27.5	17.5	1,715,635	-17.8	NA	832,157	-13.4	NA
32.5	22.5	1,704,500	-21.0	NA	830,183	-15.5	NA
	20.0	1,710,067	-19.4	NA	831,170	-14.5	NA

To achieve a realistic output representative of the UK for rosiglitazone licence and use, GSK has apportioned the final costs and effects at 85% for 4 mg and 15% for 8 mg.

TABLE 63 Cost per QALY of switching from M + S to M + R following change in insulin sensitivity

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER NA NA NA
17.5	7.5	1,842,385	-7.1	NA	881,133	-11.8	NA
22.5	12.5	1,819,052	-13.8	NA	873,726	-10.8	NA
27.5	17.5	1,809,773	-16.7	NA	871,524	-12.8	NA
32.5	22.5	1,807,713	-17.5	NA	871,159	-13.3	NA
	20.0	1,808,743	-17.1	NA	871,342	-13.1	NA

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER NA NA NA
17.5	7.5	1,789,855	-5.5	NA	858,714	-4.5	NA
22.5	12.5	1,768,742	-11.6	NA	852,011	_9.1	NA
27.5	17.5	1,757,395	-14.6	NA	849,319	-11.2	NA
32.5	22.5	1,755,386	-15.3	NA	848,963	-11.6	NA
	20.0	1,756,390	-15.0	NA	849,141	-11.4	NA

TABLE 64 Cost per QALY of switching from M + S to M + R following change in β -cell function

To achieve a realistic output representative of the UK for rosiglitazone licence and use, GSK has apportioned the final costs and effects at 85% for 4 mg and 15% for 8 mg.

The results indicate that the scenario of adding rosiglitazone directly after monotherapy failure with metformin is no longer cost-effective after the sensitivity analysis is performed (*Table 64* versus *Table 48*). In particular, it is dominated by the scenario of adding first sulfonylurea and then switching to rosiglitazone after failure of combination therapy with sulfonylurea, as the QALYs in this scenario are greater than in the one with direct adding of rosiglitazone.

Hence, it may be concluded that the investigation of this comparison of scenarios proves to be very sensitive to changes in the treatment effects.

Sensitivity analysis on the inpatient costs

Sensitivity analysis was performed on the inpatient costs, as these comprise the greatest part of the total health costs. In particular, the inpatient costs by disease state shown in *Table 65* were examined.

The cost parameters for cancer and other diseases were left unchanged.

Exactly the same scenarios were examined as presented in the previous sections.

TABLE 65 Sensitivity analysis on the inpatient costs

Disease/complication	Inpatient cost per year
Diabetes	£218
Nephropathy	£1,478
Retinopathy	£480
Neuropathy/skin infection	£1,264
Heart disease	£1,320
Stroke	£1,791
Cataract	£1,073
Source: HRG reference costs 2	2001.59

Lean patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression with insulin



The difference in costs is almost identical to that before the sensitivity analysis was performed (*Table 66* versus *Table 40*). Although the absolute costs of the two scenarios have now decreased,

TABLE 66 Cost per QALY of adding R to S (lean) following change in inpatient costs

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER NA NA NA
17.5	7.5	-470,852	204.6	NA	-229,778	171.3	NA
22.5	12.5	-466,446	207.3	NA	-228,380	173.3	NA
27.5	17.5	-466,105	207.5	NA	-228,299	173.5	NA
32.5	22.5	-466,829	207.4	NA	-228,427	173.4	NA
	20.0	-466,467	207.5	NA	-228,363	173.5	NA

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	-439,434	204.3	NA	-196,611	171.1	NA
22.5	12.5	-434,385	207.3	NA	-195,007	173.3	NA
27.5	17.5	-432,240	208.2	NA	-194,499	173.9	NA
32.5	22.5	-431,108	208.6	NA	-194,298	174.2	NA
	20.0	-431,674	208.4	NA	-194,399	174.1	NA

TABLE 67 Cost per QALY of adding R to S (obese) following change in inpatient costs

TABLE 68 Cost per QALY of adding R to M following change in inpatient costs

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER NA NA
17.5	7.5	-659,332	292.0	NA	-317,622	241.5	NA
22.5	12.5	-576,334	377.1	NA	-291,271	305.0	NA
27.5	17.5	-527,573	398.1	NA	-279,703	319.5	NA
32.5	22.5	-498,882	406.I	NA	-274,617	324.6	NA
	20.0	-513,228	402.1	NA	-277,160	322.1	NA

To achieve a realistic output representative of the UK for rosiglitazone licence and use, GSK has apportioned the final costs and effects at 85% for 4 mg and 15% for 8 mg.

their difference remains almost the same. Thus, the scenario of adding rosiglitazone still dominates in both cost and efficacy outcomes, as its cost is lower than the cost of progressing immediately to insulin after failure in monotherapy with sulfonylurea and the number of QALYs is greater.

Obese patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression to insulin

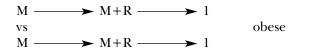


The difference in costs is almost identical to that before the sensitivity analysis was performed (*Table 67* versus *Table 42*). Although the absolute costs of the two scenarios have now decreased, their difference remains almost the same. Thus, the scenario of adding rosiglitazone still dominates in both cost and efficacy outcomes, as its cost is lower than the cost of progressing immediately to insulin after failure in monotherapy with sulfonylurea and the number of QALYs is greater. Obese patients uncontrolled on maximum dose of metformin and not prescribed sulfonylurea: comparing addition of rosiglitazone with progression to insulin



The difference in costs is almost identical to that before the sensitivity analysis was performed (*Table 68* versus *Table 44*). Although the absolute costs of the two scenarios have now decreased, their difference remains almost the same. Thus, the scenario of adding rosiglitazone still dominates in both cost and efficacy outcomes, as its cost is lower than the cost of progressing immediately to insulin after failure in monotherapy with metformin and the number of QALYs is greater.

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone to addition of sulfonylurea and then progressing to insulin



65

The difference in costs is almost identical to that before the sensitivity analysis was performed (*Table 69* versus *Table 46*). Although the absolute costs of the two scenarios have now decreased, their difference remains almost the same. The scenario of adding rosiglitazone instead of sulfonylureas after monotherapy failure with metformin remains cost-effective after the change in the inpatient costs.

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then switching to rosiglitazone

M → M+S →	\rightarrow M+R \longrightarrow	1
VS		obese
$M \longrightarrow M+R$	▶ 1	

The difference in costs is almost identical to that before the sensitivity analysis was performed (*Table 70* versus *Table 48*). Although the absolute costs of the two scenarios have now decreased, their difference remains almost the same. Thus, the scenario of adding rosiglitazone directly after monotherapy failure with metformin remains costeffective after the sensitivity analysis is performed compared with the scenario of adding first sulfonylurea and then switching to rosiglitazone after failure of combination therapy with sulfonylurea.

Sensitivity analysis on the insulin cost

Sensitivity analysis was performed on the annual cost of insulin treatment. In particular, the insulin cost was reduced by 50%, to investigate possible changes in the cost-effectiveness results. Thus, the insulin cost was set equal to $\pounds 305.59$.

Exactly the same scenarios were examined as presented in the previous sections.

Lean patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression to insulin



Table 71 shows that after the change in the cost of insulin treatment the rosiglitazone scenario is no longer dominant (*Table 71* versus *Table 40*). The results indicate that the addition of rosiglitazone after monotherapy failure with sulfonylurea has a

TABLE 69 Cost per QALY of M + S versus M + R following change in inpatient costs

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	1,449,038	89.1	16,270	713,304	71.6	9968
22.5	12.5	1,529,229	170.8	8951	738,764	132.5	5575
27.5	17.5	1,569,218	188.5	8326	748,252	144.7	5170
32.5	22.5	1,588,863	193.7	8202	751,734	148.1	5076
	20.0	1,579,041	191.1	8263	749,993	146.4	5123

To achieve a realistic output representative of the UK for rosiglitazone licence and use, GSK has apportioned the final costs and effects at 85% for 4 mg and 15% for 8 mg.

TABLE 70 Cost per QALY of switching from M + S to M + R following change in inpatient costs

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	1,482,083	76.4	19,411	727,334	61.4	11,854
22.5	12.5	1,504,106	138.0	10,899	734,326	107.3	6842
27.5	17.5	1,533,688	146.9	10,444	741,345	113.4	6535
32.5	22.5	1,549,267	151.4	10,234	744,107	116.4	6395
	20.0	1,541,478	149.1	10,337	742,726	114.9	6464

higher cost than progressing immediately to insulin after monotherapy failure. However, it is a cost-effective option, as the ICERs are very low (average of 996).

Obese patients uncontrolled on maximum dose of sulfonylurea, unable to take metformin: comparing addition of rosiglitazone with progression to insulin



Table 72 shows that after the change in the cost of insulin treatment the rosiglitazone scenario is no

longer dominant (*Table 72* versus *Table 42*). The results indicate that the addition of rosiglitazone after monotherapy failure with sulfonylurea has a higher cost than progressing immediately to insulin after monotherapy failure. However, it is a cost-effective option, as the ICERs are very low (average of 857).

Obese patients uncontrolled on maximum dose of metformin and not prescribed sulfonylurea: comparing addition of rosiglitazone with progression to insulin



Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	351,591	204.6	1718	171,159	171.3	999
22.5	12.5	355,837	207.3	1717	172,508	173.3	995
27.5	17.5	356,149	207.5	1716	172,581	173.5	995
32.5	22.5	355,449	207.4	1714	172,458	173.4	994
	20.0	355,799	207.5	1715	172,520	173.5	995

TABLE 72 Cost per QALY of adding R to S (obese) after changing insulin cost

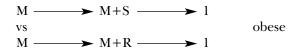
Time from diagnosis	Years from start of combination	U	Undiscounted		Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	s I.5% ICER
17.5	7.5	328,942	204.3	1610	147,003	171.1	859
22.5	12.5	333,742	207.3	1610	148,527	173.3	857
27.5	17.5	335,728	208.2	1613	148,999	173.9	857
32.5	22.5	336,770	208.6	1614	149,183	174.2	856
	20.0	336,249	208.4	1613	149,091	174.1	856

TABLE 73 Cost per QALY of adding R to M after changing insulin cost

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	861,734	292.0	2951	414,269	241.5	1715
22.5	12.5	1,043,727	377.1	2767	472,050	305.0	1548
27.5	17.5	1,096,565	398.1	2755	484,586	319.5	1517
32.5	22.5	1,123,832	406.I	2767	489,419	324.6	1508
	20.0	1,110,198	402.1	2761	487,002	322.1	1512

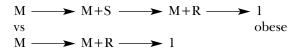
The sensitivity analysis demonstrated that rosiglitazone combination therapy with metformin does not dominate the alternative scenario of progressing directly to insulin after metformin failure (*Table 73* versus *Table 44*). The results in *Table 73* indicate that the addition of rosiglitazone after monotherapy failure with metformin has a higher cost than progressing immediately to insulin after monotherapy failure. However, it is a cost-effective option, as the ICERs are very low (average of 1560).

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then progressing to insulin



Rosiglitazone combination therapy with metformin is less cost-effective after the sensitivity analysis compared with the alternative scenario of sulfonylurea combination therapy after metformin failure (*Table 74* versus *Table 46*). The results in *Table 74* indicate that the addition of rosiglitazone has a higher cost than the addition of sulfonylurea after monotherapy failure with metformin. However, it is still a cost-effective option, as the ICERs are lower than £30,000.

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then switching to rosiglitazone



Rosiglitazone combination therapy with metformin is less cost-effective after the sensitivity analysis compared with the alternative scenario of sulfonylurea combination therapy after metformin failure and then substitution of sulfonylurea with rosiglitazone (*Table 75* versus *Table 48*). The results in *Table 48* indicate that the addition of rosiglitazone after monotherapy failure with metformin has a higher cost than the addition of sulfonylurea and then switching to rosiglitazone. However, it is still a cost-effective option, as the ICERs are lower than £30,000.

TABLE 74 Cost per QALY of M + S versus M + R after changing insulin cost

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	1,634,376	89.1	18,351	792,041	71.5	11,070
22.5	12.5	1,814,055	170.8	10,619	849,087	132.5	6407
27.5	17.5	I,858,804	188.5	9862	859,705	144.7	5940
32.5	22.5	1,877,575	193.7	9692	863,032	148.1	5828
	20.0	1,868,189	191.1	9776	861,368	146.4	5884

To achieve a realistic output representative of the UK for rosiglitazone licence and use, GSK has apportioned the final costs and effects at 85% for 4 mg and 15% for 8 mg.

TABLE 75 Cost per QALY switching from M + S to M + R after changing insulin cost

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	1,561,105	76.4	20,446	760,901	61.4	12,401
22.5	12.5	1,582,470	138.0	11,466	767,685	107.3	7153
27.5	17.5	1,617,349	146.9	11,014	775,960	113.4	6841
32.5	22.5	1,632,073	151.4	10,777	778,570	116.4	6692
	20.0	1,624,711	149.1	10,893	777,265	114.9	6765

Sensitivity analysis on the dosage of rosiglitazone

According to GSK, 85% of rosiglitazone use is seen in the 4 mg dose. The rest of the use (15%) is in the 8 mg dose, which is combined only with metformin. In all the previous scenarios of combination therapy with metformin and rosiglitazone, the GSK assumption was followed and the costs and outputs were calculated using 85% of the cost–effect of rosiglitazone on the 4 mg dosage and 15% of the cost–effect of using rosiglitazone on the 8 mg dosage.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Thus, in this section, several scenarios are compared, assuming that 100% of rosiglitazone is seen in the 8 mg dose. The scenarios compared here are those that include metformin.

Obese patients uncontrolled on maximum dose of metformin and not prescribed sulfonylurea: comparing addition of rosiglitazone with progression to insulin

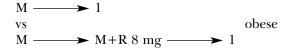
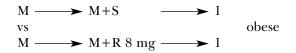


TABLE 76 Cost per QALY of adding R to M (obese) using 8 mg dose

The sensitivity analysis demonstrated that the scenario of adding 8 mg of rosiglitazone after failure of monotherapy with metformin does not dominate the scenario of progressing directly to insulin therapy (*Table 76* versus *Table 44*). The results indicate that the cost of the second scenario is now higher than the cost of the first scenario. However, the strategy of adding rosiglitazone after metformin failure is a cost-effective option compared with progressing to insulin, as the ICERs are very low in all cases (average of 1399).

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then progressing to insulin



The results in *Table* 77 indicate that adding 8 mg of rosiglitazone after monotherapy failure with metformin remains a cost-effective option compared with adding sulfonylurea. Nevertheless, the marginal ICERs are higher after the sensitivity analysis (*Table* 77 versus *Table* 46).

Obese patients uncontrolled on maximum dose of metformin: comparing addition of rosiglitazone with addition of sulfonylurea and then switching to rosiglitazone

Time from diagnosis	Years from start of combination	U	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER	
17.5	7.5	818,962	304.6	2689	392,194	251.8	1558	
22.5	12.5	1,255,955	461.5	2722	530,934	368.7	1440	
27.5	17.5	1,411,696	542.3	2603	567,883	424.7	1337	
32.5	22.5	1,514,910	567.8	2668	586,164	441.0	1329	
	20.0	1,463,303	555.0	2636	577,024	432.9	1333	

TABLE 77 Cost per QALY of M + S versus M + R using 8 mg dose

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	2,927,354	101.7	28,798	1,423,125	81.8	17,389
22.5	12.5	3,361,345	255.2	13,174	1,560,912	196.3	7953
27.5	17.5	3,507,458	332.7	10,543	1,595,577	249.9	6384
32.5	22.5	3,600,686	355.4	10,132	1,612,087	264.5	6095
	20.0	3,554,072	344.0	10,330	1,603,832	257.2	6236

$$M \rightarrow M+S \rightarrow M+R \ 8 \ mg \rightarrow 1$$

vs obese
$$M \rightarrow M+R \ 8 \ mg \rightarrow 1$$

The sensitivity analysis clearly demonstrates that adding 8 mg of rosiglitazone after metformin monotherapy failure is not a cost-effective option compared with adding sulfonylurea and then switching to rosiglitazone. The marginal ICERs have changed after the variation in the rosiglitazone dosage (*Table 78* versus *Table 48*).

Conclusions on the health economics of treatment with rosiglitazone in Type 2 diabetes

 There is an inadequate exploration of uncertainty within the model. No sensitivity analysis was performed by the company on the variables that drive the model: treatment effects on insulin sensitivity and β-cell function, costs and glucose threshold. Furthermore, it was difficult to perform sensitivity analysis on the treatment effects, as the exact relationship between insulin sensitivity and β -cell function was unknown to the reviewers.

- Several submodels are presented in the GSK model, which are built based on large databases. However, no information about the methods or diagnostics of the models was provided by the company.
- The baseline economic results for rosiglitazone indicate that rosiglitazone is a cost-effective treatment for type 2 diabetes. Nevertheless, the sensitivity analysis indicated that some of the scenarios are very sensitive to changes in key effectiveness variables.

The key economic results and their variations after the sensitivity analyses are summarised in *Table 79*.

TABLE 78 Cost per QALY of switching from M + S to M + R using 8 mg dose

Time from diagnosis	Years from start of combination	Undiscounted			Discounted		
	therapy	Costs	QALYs	ICER	Costs 6%	QALYs 1.5%	ICER
17.5	7.5	2,833,917	14.0	202,857	1,383,611	11.4	121,263
22.5	12.5	2,879,952	29.8	96,740	1,398,227	23.2	60,320
27.5	17.5	2,926,344	38.6	75,773	1,409,233	29.3	48,080
32.5	22.5	2,939,483	43.2	67,996	1,411,545	32.3	43,742
	20.0	2,932,914	40.9	71,666	1,410,389	30.8	45,807

TABLE 79 Summary of key economic results and sensitivity analyses (discounted cumulative values per 1000 patients)

7.5 years after start of combination therapy	Comparison of scenarios	S→I vs S→S+R→I lean	S→I vs S→S+R→I overweight	M→I vs M→M+R→I	M→M+S→I vs M→M+R→I	M→M+S→M+R→I vs M→M+R→I
Central estimate	Cost/QALY	R dominates	R dominates	R dominates	9972	11,857
	Cost/LYG	R dominates	R dominates	R dominates	262,253	259,591
Sensitivity analysis Cost/QALY	HbA _{1c} threshold	R dominates	R dominates	R dominates	3937	24,962
	Insulin sensitivity	R dominates	R dominates	R dominates	Comparator dominates	Comparator dominates
	eta-cell function	R dominates	R dominates	R dominates	Comparator dominates	Comparator dominates
	Inpatient costs	R dominates	R dominates	R dominates	9968	11,854
	Insulin cost	999	859	1715	11,070	12,401
	Rosiglitazone dosage	-	-	1558	17,389	121,263

Comparison of the Takeda model and the GSK model

TABLE 80 Comparing the cost per LYG of the Takeda and GSK models

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

TABLE 81 Efficacy of comparators in the two models

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

TABLE 82 Cost per LYG of pioglitazone using data inputs of the GSK model

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Conclusions

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

The results of the sensitivity analyses suggest that the cost per QALY of rosiglitazone is most sensitive to two variables: dosage and a change in the treatment effect, that is, the effect of rosiglitazone on β -cell function and insulin sensitivity (*Table 79*).

The cost per QALY ratio of rosiglitazone is most sensitive to a change in its treatment effect. In the central estimate provided by GSK, rosiglitazone is cost-effective in all scenarios, whether it is compared with metformin or sulfonylurea. However, if the effect of rosiglitazone on β -cell function and insulin sensitivity is reduced so that the resulting HbA_{1c} reduction is that of the higher confidence interval, the cost-effectiveness changes dramatically. In the two scenarios where rosiglitazone is compared with metformin and sulfonylurea combination therapy, the costeffectiveness of rosiglitazone switches from being cost-effective to being dominated by the comparator strategy.

This suggests that small changes in the effect of rosiglitazone on β -cell function and insulin sensitivity induce large changes in the cost per QALY ratios. Therefore, there is a high level of uncertainty associated with the treatment effect of rosiglitazone. Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report. As seen in Table 79, the cost per QALY ratio of rosiglitazone increases considerably if it is assumed that 100% of patients receive 8 mg.

GSK did not perform any univariate or multivariate sensitivity analyses on the model. Furthermore, owing to the complexity of the model it has not been possible within the rapid review timescales to conduct a full multivariate analysis. The univariate sensitivity analyses that have been performed indicate that there is a wide degree of uncertainty; for example, ranging from a cost-effectiveness under £20,000 per QALY to being dominated, in the key scenarios under consideration.

The cost per QALY ratios presented in the GSK model suggest that rosiglitazone therapy, combined with either metformin or sulfonylurea, is an economically attractive option. The current NICE guideline suggests that glitazone combination therapy should only be tried after metformin and sulfonylurea combination therapy has failed, whereas the cost per QALY ratios presented in the GSK model suggest that it is potentially economically viable to use rosiglitazone combination therapy directly after failure of monotherapy with either metformin or sulfonylurea. However, since the baseline estimate of cost-effectiveness is not robust to changes in the treatment effect and is reliant on the many assumptions included within the metabolic and long-term economic models, caution should be applied in interpreting this baseline favourable result.

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.

Chapter 5 Impact on the NHS

There is a high level of uncertainty in the potential budgetary impact of the glitazones on the NHS; any estimates inevitably are based heavily on a series of assumptions, many of which cannot be justified easily.

It is thought that an estimated 800,000 people within England and Wales have type 2 diabetes.⁶⁰ This figure may be an underestimate as the King's Fund Report⁶¹ of 1996 estimates that roughly two million people over the age of 16 in the UK suffer from NIDDM. Crudely weighting this prevalence of type 2 diabetes in the UK to England and Wales alone⁶² suggests that approximately 1.7 million people, diagnosed and undiagnosed, may suffer from the disease.

The UK Drug Information Pharmacists Group (UKDIPG)^{63,64} makes the assumption that 50% of all diabetic patients currently on oral monotherapy are controlled inadequately. These estimates have a major impact on the number of patients eligible for treatment with the glitazones. The upper estimate of people potentially eligible for treatment with a glitazone in England and Wales is 212,000.

Pioglitazone

A valid estimation of the costs to the NHS must include both the 30 mg pioglitazone dose and the lower 15 mg dose. In the absence of any information to support an estimate of the breakdown of prescribing between doses, an assumption of a 50:50 split is just as valid as any other estimate. It should be noted in this regard that the economics of treatment with the 15 mg dose have not been addressed within the economic submission.

The daily cost of pioglitazone is £1.32 for the 30 mg dose and £0.95 for the 15 mg dose.⁸ The average annual cost per patient, therefore, is the average of the annual treatment at 15 mg and at 30 mg. This gives a figure of £414.28 for the estimated annual cost of pioglitazone. This is then multiplied by the number of eligible patients taking it in combination with sulfonylurea and metformin, respectively. The resulting estimates of

gross drug costs for pioglitazone in England and Wales, assuming all eligible patients are prescribed, are:

- P+S: £414.28 × 34,700 people = £14,375,516
- P+M: $\pounds 414.28 \times 40,000$ people = $\pounds 16,571,200$
- Gross cost of pioglitazone in England and Wales = £30,946,716

The gross cost of pioglitazone should, however, be considered against potential savings in other available treatment options. When used in combination with sulfonylureas, the only alternative add-on therapy considered in the Takeda submission²⁶ is acarbose, although this is not widely used in the UK. The only other intervention considered is to switch to intensive insulin therapy. When used in combination with metformin, the alternative add-on therapies considered in the Takeda submission are acarbose and sulfonylureas. As discussed above, acarbose is not widely used in the UK. The maximum potential savings from the reduced usage of insulin and sulfonylureas are detailed below, assuming that all patients using pioglitazone would have used one of these two alternative therapies (a potentially generous assumption) and that the average insulin daily dose saved is 60 units.

• Saving from insulin (average 60 U) £611 × 34,700 people = £21,201,700

•	Saving from sulfonylureas	
	$\pm 132 \times 40,000$ people	= £5,280,000
-	T 1 1 1 1 1 1 1 1 1 1	COC 401 700

• Total maximum potential saving = £26,481,700

A baseline estimate of the net annual cost of pioglitazone to the NHS would therefore be in the region of £4.5 million. Depending on the different assumptions used in the analysis, this figure could vary between an estimated saving of £12 million and a cost of £30 million.

The key assumptions are:

- the proportion of eligible patients who receive treatment
- the average saving in the use of insulin by patients who receive pioglitazone

• the average saving in the use of sulfonylurea by patients who receive pioglitazone.

Rosiglitazone

A valid estimation of the costs to the NHS must include both the 4 mg and 8 mg rosiglitazone doses. In the absence of any information to support an estimate of the breakdown of prescribing between doses, an assumption of a 50:50 split is just as valid as any other estimate.

The daily cost of rosiglitazone is £0.95 for the 4 mg dose and £1.95 for the 8 mg dose.⁸ The average annual cost per patient, therefore, is the average of the annual treatment at 4 mg and at 8 mg. This gives a figure of £529.25 for the estimated annual cost of rosiglitazone. This is then multiplied by the number of eligible patients taking it in combination with sulfonylurea and metformin, respectively. The resulting estimates of gross drug costs for rosiglitazone in England and Wales, assuming all eligible patients are prescribed, are:

- R+S: $\pounds 529.25 \times 34,700$ people
- R+M: = $\pounds 18,364,975$ $\pounds 529.25 \times 40,000$ people = $\pounds 21,170,000$
- Gross cost of rosiglitazone in England and Wales = £39,534,975

The gross cost of rosiglitazone should, however, be considered against potential savings in other

available treatment options. When used in combination with sulfonylureas, the only alternative is to switch to intensive insulin therapy.²⁵ When used in combination with metformin, the alternative add-on therapies considered in the GSK submission are sulfonylureas. The maximum potential savings from the reduced usage of insulin and sulfonylureas are detailed below, assuming that all patients using rosiglitazone would have used one of these two alternative therapies (a potentially generous assumption) and that the average insulin daily dose saved is 60 units.

- Saving from insulin (average 60 U) £611 × 34,700 people = £21,201,700
- Saving from sulfonylureas $\pounds 132 \times 40,000$ people = $\pounds 5,280,000$
- Total maximum potential savings = £26,481,700

A baseline estimate of the net annual cost of rosiglitazone to the NHS would therefore be in the region of £13 million, with a similar degree of uncertainty to that for pioglitazone.

The key assumptions are:

- the proportion of eligible patients who receive treatment
- the average saving in the use of insulin by patients who receive rosiglitazone
- the average saving in the use of sulfonylurea by patients who receive rosiglitazone.

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Chapter 6 Recommendations for research

T he report demonstrates that a number of relevant clinical studies of pioglitazone and rosiglitazone are not within the public domain. It is recommended that research already undertaken be published, preferably in peer-reviewed journals.⁴

The authors would recommend that further research into the effects of the glitazones be considered. Specifically, prospective RCTs would be helpful in comparing the effects of rosiglitazone in combination with metformin or sulfonylurea against pioglitazone in combination with metformin or sulfonylurea.

The use of a glitazone in triple oral combination therapy or in combination with insulin is classified in the licences under 'special warning and special precautions for use'. This precaution is based on the fact that at the time the licences were issued there was no clinical experience of triple therapy. There is, however, emerging evidence of use of the glitazones within such combination therapies and prospective RCTs would be useful. These studies could examine short-term transition strategies and longer term management.

The current evidence base focuses largely on physiological outcome measures. The impact of the glitazones in delaying transfer to insulin and the impact on longer term outcomes could also be investigated.



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Contributions of authors

Carolyn Czoski-Murray (Researcher in Health and Related Research) and Johanna Cowan (Research Assistant in Health and Related Research) carried out the review of clinical effectiveness. Emma Warren (Operational Research Analyst in Health and Related Research), Jim Chilcott (Senior Operational Research Analyst in Health and Related Research) and Maria Psyllaki (Operational Research Analyst in Health and Related Research) carried out the review of cost-effectiveness. Catherine Beverley (Systematic Reviews Information Officer in Health and Related Research) carried out the electronic searches.

Relationship of reviewer(s) with sponsor

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Appendix I

Electronic bibliographic databases searched

- 1. CINAHL
- 2. Cochrane Controlled Trials Register (CCTR)
- 3. Cochrane Database of Systematic Reviews (CDSR)
- 4. Database of Abstracts of Reviews of Effectiveness (DARE)
- 5. EMBASE

- 6. HTA Database
- 7. MEDLINE
- 8. NHS EED
- 9. OHE HEED
- 10. PreMEDLINE
- 11. Science Citation Index (SCI)
- 12. Social Sciences Citation Index (SSCI)

Appendix 2

Other sources consulted

- 1. Agency for Healthcare Research and Quality (AHRQ)
- 2. Aggressive Research Intelligence Facility (ARIF)
- 3. Association of British Clinical Diabetologists
- 4. Association of Diabetes Specialist Nurses
- 5. Aventis
- 6. Bandolier
- 7. British Dietetic Association
- 8. Canadian Co-ordinating Centre for Health Technology Assessment (CCOHTA)
- 9. CenterWatch Trials Register
- 10. Centre for Health Economics, University of York
- 11. Copernic
- 12. Current Controlled Trials (CCT)
- 13. Current Research in Britain (CRiB)
- 14. Department of Health
- 15. Diabetes Foundation
- 16. Diabetes UK
- 17. eBNF
- 18. Electronic Medicines Compendium
- 19. eGuidelines
- 20. European Agency for the Evaluation of Medicinal Products (EMEA)
- 21. Food and Drug Administration (FDA)
- 22. Google
- 23. Health Evidence Bulletins, Wales
- 24. Heart Disease and Diabetes Research Trust

- 25. International Network of Agencies for Health Technology Assessment (INAHTA) Clearinghouse
- 26. Index to Theses
- 27. Medlineplus Drug Information
- 28. MeReC
- 29. Medical Research Council (MRC) Funded Projects Database
- 30. National Assembly for Wales
- 31. National Guideline Clearinghouse (NGC)
- 32. National Research Register (NRR)
- 33. National Coordinating Centre for Health Technology Assessment (NCCHTA)
- 34. Organising Medical Networked Information (OMNI)
- 35. Primary Care Diabetes UK
- 36. Research Findings Register (ReFeR)
- 37. Royal College of Physicians
- 38. ScHARR Library Catalogue
- 39. Scottish InterCollegiate Guideline Network (SIGN)
- 40. Trent Working Group on Acute Purchasing
- 41. Turning Research into Practice (TRIP) Database
- 42. Wessex Development and Evaluation Committee (DEC) Reports
- 43. West Midlands Development and Evaluation Services (DES) Reports
- 44. World Health Organization (WHO)

Appendix 3

Search strategies used in the major electronic bibliographic databases

CDSR and CCTR

2002 Issue 1 The Cochrane Library, Update Software (CD-ROM version) Search undertaken April 2002

#1. GLITAZONE*
#2. THIAZOLIDINEDION*
#3. PPAR GAMMA AGONIST*
#4. PIOGLITAZONE*
#5. ACTOS
#6. ROSIGLITAZONE*
#7. AVANDIA
#8. TROGLITAZONE*
#9. #1 OR #2 OR #3 OR #4 OR #5 OR #6 OR #7 OR #8

CINAHL

1982–2002 Ovid Biomed Search undertaken April 2002

- 1 glitazone\$.af
- 2 thiazolidinedione.af
- 3 peroxisome proliferator activated receptor gamma agonist\$.tw
- 4 ppar gamma agonist\$.af
- 5 actos.af
- 6 1101025-46-8.rn
- 7 ad-4833.af
- 8 u-72107.af
- 9 pioglitazone\$.af
- 10 rosiglitazone\$.af
- 11 avandia.af
- 12 122320-73-4.rn
- 13 brl-49653.af
- 14 troglitazone\$.af
- 15 97322-87-7.rn
- 16 thiazole\$.ti
- 17 or/1-16
- 18 exp diabetes mellitus/
- 19 diabet\$.tw
- 20 or/18-19
- $21 \hspace{.1in} 17 \hspace{.1in} and \hspace{.1in} 20$

CRD databases (NHS DARE, EED, HTA)

CRD website – complete databases Search undertaken April 2002

glitazone or thiazolidinedione or ppar gamma agonist or pioglitazone or actos or rosiglitazone or avandia or troglitazone/All fields

Embase

1980–2002 SilverPlatter WebSPIRS Search undertaken April 2002

- #1 glitazone*
- #2 thiazolidineione*
- #3 peroxisome proliferator activated receptor gamma agonist*
- #4 ppar gamma agonist*
- #5 pioglitazone*
- #6 actos
- #7 111025-46-8
- #8 ad-4833
- #9 u-72107
- #10 rosiglitazone*
- #11 avandia
- #12 122320-73-4
- #13 brl-49653
- #14 troglitazone*
- #15 97322-87-7
- #16 #1 or #2 or #3 or #4 or #5 or #6 or #7 or #8 or #9 or #10 or #11 or #12 or #13 or #14 or #15
- #17 explode 'diabetes-mellitus' / all subheadings
- #18 diabet*
- #19 #17 or #18
- #20 #16 and #19

OHE HEED

CD ROM version Search undertaken April 2002

Search terms

• glitazone* or pioglitazone* or rosiglitazone* or avandia or actos or ppar gamma agonist*

Fields searched

• All data

MEDLINE

1966–2002 Ovid Biomed Search undertaken April 2002

- 1 glitazone\$.af
- 2 thiazolidinedione.af
- 3 peroxisome proliferator activated receptor gamma agonist\$.tw
- 4 ppar gamma agonist\$.af
- 5 actos.af
- 6 1101025-46-8.rn
- 7 ad-4833.af
- 8 u-72107.af
- 9 pioglitazone\$.af
- 10 rosiglitazone\$.af
- 11 avandia.af

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- 12 122320-73-4.rn
- 13 brl-49653.af
- 14 troglitazone\$.af
- 15 97322-87-7.rn
- 16 thiazole\$.ti
- 17 or/1-16
- 18 exp diabetes mellitus/
- 19 diabet\$.tw
- $20 \quad \text{or}/18-19$
- 21 17 and 20

SCI and SSCI

1981–2002 Web of Science Search undertaken April 2002

Title=(glitazone* or thiazolidinedione or ppar gamma agonist* or actos or pioglitazone* or rosiglitazone* or avandia or troglitazone*) and diabet* not (rat or rats or mice or mouse); DocType=All document types; Languages=All languages; Databases=SCI-EXPANDED, SSCI; Timespan=All Years

Appendix 4

Economic evaluations and quality of life methodological search filters used in MEDLINE (Ovid) 1966–June 2002

Economic evaluations

- 1 economics/
- 2 exp "costs and cost analysis"/
- 3 economic value of life/
- 4 exp economics, hospital/
- 5 exp economics, medical/
- 6 economics, nursing/
- 7 economics, pharmaceutical/
- 8 exp models, economic/
- 9 exp "fees and charges"/
- 10 exp budgets/
- 11 ec.fs
- 12 (cost or costs or costed or costly or costing\$).tw
- 13 (economic\$ or pharmacoeconomic\$ or price\$ or pricing).tw
- 14 or/1-13

Quality of life

- 1 exp quality of life/
- 2 quality of life.tw

- 3 life quality.tw
- 4 hql.tw
- 5 (sf 36 or sf36 or sf thirtysix or sf thirty six or short form 36 or short form thirty six or short form thirtysix or shortform 36).tw

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- 6 qol.tw
- 7 (euroqol or eq5d or eq 5d).tw
- 8 qaly\$.tw
- 9 quality adjusted life year\$.tw
- 10 hye\$.tw
- 11 health\$ year\$ equivalent\$.tw
- 12 health utilit\$.tw
- 13 hui.tw
- 14 quality of wellbeing\$.tw
- 15 quality of well being.tw
- 16 qwb.tw
- 17 (qald\$ or qale\$ or qtime\$).tw
- 18 or/1-17

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Appendix 5 Correspondence with GSK

Information from the sponsor's submission was submitted in confidence to NICE. This information was made available to the NICE Appraisals Committee but has been removed from this version of the report.



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Feedback

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The Correspondence Page on the HTA website (http://www.ncchta.org) is a convenient way to publish your comments. If you prefer, you can send your comments to the address below, telling us whether you would like us to transfer them to the website.

We look forward to hearing from you.

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